

It is enough to show that on this subsequence the second and third terms in decomposition (5) tends to 0. For the second term

$$\begin{aligned} \left| \int_A (X_\tau - X_{\tau \wedge n_k}) d\mathbf{P} \right| &= \left| \int_{A \cap \{\tau > n_k\}} (X_\tau - X_{\tau \wedge n_k}) d\mathbf{P} \right| \\ &\leq \int_{A \cap \{\tau > n_k\}} (|X_\tau| + |X_{n_k}|) d\mathbf{P} \\ &\leq \int_{\{\tau > n_k\}} |X_\tau| d\mathbf{P} + \int_{\{\tau > n_k\}} |X_{n_k}| d\mathbf{P}. \end{aligned}$$

Similarly, for the third term

$$\begin{aligned} \left| \int_A (X_\sigma - X_{\sigma \wedge n_k}) d\mathbf{P} \right| &= \left| \int_{A \cap \{\sigma > n_k\}} (X_\sigma - X_{n_k}) d\mathbf{P} \right| \\ &\leq \int_{\{\sigma > n_k\}} |X_\sigma| d\mathbf{P} + \int_{\{\tau > n_k\}} |X_{n_k}| d\mathbf{P}. \end{aligned}$$

Using (2) both upper bounds tend to 0. \square

Corollary 2. Assume that (X_n) is (super-, sub-) martingale, τ is a stopping time, $\mathbf{E}(|X_\tau|) < \infty$ and (3) holds. Then

- (i) $\mathbf{E}(X_\tau | \mathcal{F}_1) \leq X_1$ and $\mathbf{E}(X_\tau) \leq \mathbf{E}(X_1)$ for supermartingales;
- (ii) $\mathbf{E}(X_\tau | \mathcal{F}_1) \geq X_1$ and $\mathbf{E}(X_\tau) \geq \mathbf{E}(X_1)$ for submartingales;
- (iii) $\mathbf{E}(X_\tau | \mathcal{F}_1) = X_1$ and $\mathbf{E}(X_\tau) = \mathbf{E}(X_1)$ for martingales.

Some conditions are needed for the optional stopping to hold.

Example 2 (Simple symmetric random walk). Let ξ, ξ_1, ξ_2, \dots are iid random variables with $\mathbf{P}(\xi = \pm 1) = 1/2$. Let $S_0 = 1$ and $S_n = S_{n-1} + \xi_n$. Then (S_n) is martingale. Let $\tau = \min\{n : S_n = 0\}$. Then τ is a stopping time and the martingale $(S_{\tau \wedge n})_n$ tends to 0 a.s. The optional stopping does not hold as $S_\tau \equiv 0$ a.s., while $S_0 = 1$. Clearly, condition (3) does not hold.

Theorem 5 (Wald identity). Let X, X_1, X_2, \dots be iid random variables with $\mathbf{E}X = \mu \in \mathbb{R}$, and let τ be a stopping time with $\mathbf{E}(\tau) < \infty$. Put $S_n = X_1 + \dots + X_n$, $n \in \mathbb{N}$. Then $\mathbf{E}(S_\tau) = \mu \mathbf{E}(\tau)$.

$$\begin{aligned} \mathbf{E}(S_n) &= \mathbf{E} \sum_{i=1}^n X_i \\ &= \sum_{i=1}^n \mathbf{E} X_i = n\mu. \end{aligned}$$

$$\{\tau \geq k-1\} = \{\tau \geq k\} \in \mathcal{F}_{k-1}$$

Proof. First assume $X \geq 0$. We have

$$\begin{aligned} \mathbf{E}(X) &= \int_0^{\infty} \mathbf{P}(X > x) dx \\ &= \sum_{k=1}^{\infty} \mathbf{P}(X \geq k) \end{aligned}$$

$$\begin{aligned} \mathbf{E}(S_{\tau}) &= \mathbf{E}\left(\sum_{k=1}^{\tau} X_k\right) \\ &= \sum_{k=1}^{\infty} \mathbf{E}(\mathbf{I}(\tau \geq k) X_k) \\ &\stackrel{?}{=} \sum_{k=1}^{\infty} \mathbf{E}(\mathbf{I}(\tau \geq k)) \mathbf{E}(X_k) \\ &= \mu \sum_{k=1}^{\infty} \mathbf{P}(\tau \geq k) \\ &= \mu \mathbf{E}(\tau). \end{aligned}$$

$$\begin{aligned} \sum_{i=1}^{\tau} X_i &\in \mathcal{F}_{k-1} = \sigma(X_1, \dots, X_{k-1}) \\ \mathbf{E}\left[\mathbf{I}(\tau \geq k) \cdot X_k\right] &= \\ &= \mathbf{E}(\mathbf{I}(\tau \geq k)) \cdot \mathbf{E}(X_k) \\ &\uparrow \\ &\text{independence} \quad \checkmark \end{aligned}$$

To see the general case consider the decomposition $S_{\tau} = S_{\tau}^{(+)} - S_{\tau}^{(-)}$ where

$$S_{\tau}^{(+)} = \sum_{k=1}^{\infty} X_k^{+} \mathbf{I}(\tau \geq k)$$

$$a^{+} = \max(0, a)$$

and

$$S_{\tau}^{(-)} = \sum_{k=1}^{\infty} X_k^{-} \mathbf{I}(\tau \geq k).$$

$$a^{-} = \max(0, -a)$$

$$a = a^{+} - a^{-}$$

As a simple application of the optional stopping problem we consider the gambler's ruin problem. There is an elementary but longer way to derive these formulas.

Example 3 (Gambler's ruin). Let X, X_1, X_2, \dots be iid random variables such that $\mathbf{P}(X = 1) = p = 1 - \mathbf{P}(X = -1)$, $0 < p < 1$, and put $S_n = X_1 + \dots + X_n$, $n \in \mathbb{N}$. Fix $a, b \in \mathbb{N}$ and let

$$\tau = \tau_{a,b}(p) = \inf\{n : S_n \geq b \text{ or } S_n \leq -a\},$$

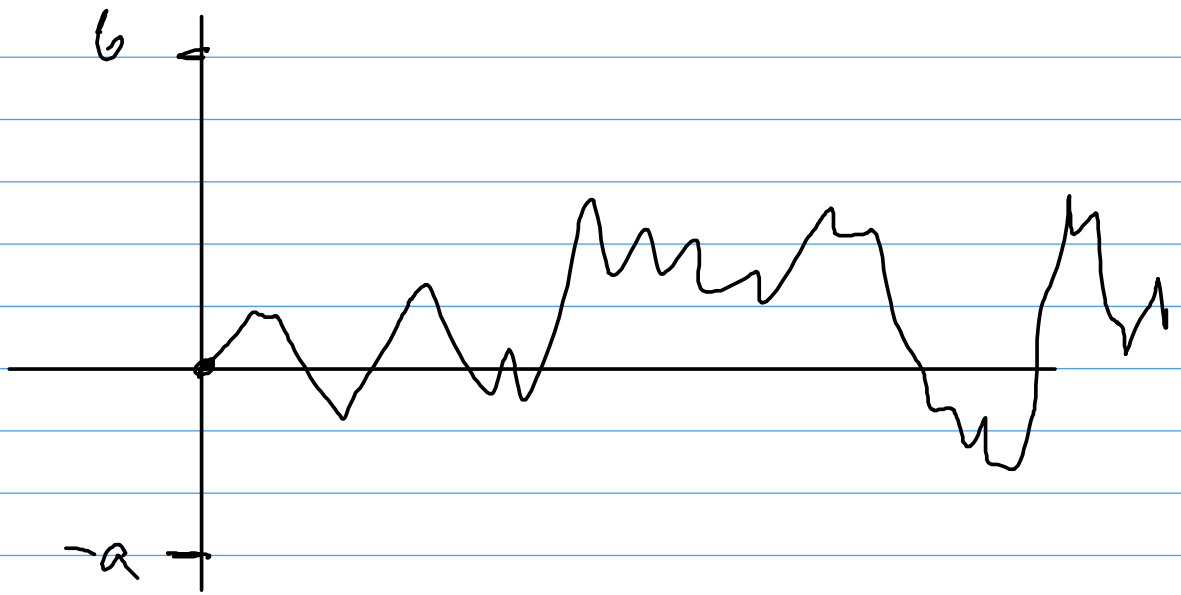
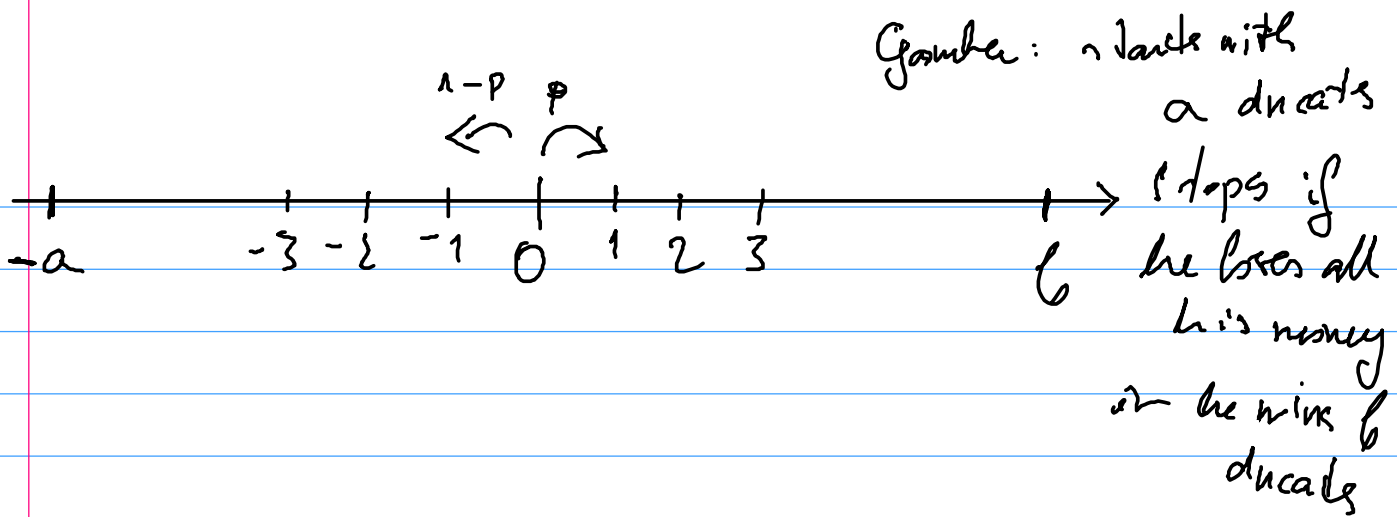
with the convention $\inf \emptyset = \infty$. Let (\mathcal{F}_n) be the natural filtration, i.e. $\mathcal{F}_n = \sigma(X_1, \dots, X_n)$, $n \in \mathbb{N}$. HW

It is easy to show that $\mathbf{P}(\tau < \infty) = 1$, and τ is a stopping time. Furthermore, $|S_{\tau}| \leq \max(a, b)$, in particular $\mathbf{E}|S_{\tau}| < \infty$ and

$$\liminf_{n \rightarrow \infty} \int_{\{\tau > n\}} |S_n| d\mathbf{P} \leq \liminf_{n \rightarrow \infty} \max(a, b) \mathbf{P}(\tau > n) = 0.$$

8

We can apply optional st. pp's



$$\boxed{p = \frac{1}{2}}$$

If $\boxed{p = \frac{1}{2}}$ then $E(X) = 0 \Rightarrow (S_n)$ is a martingale
 $E(S_T) = E(S_0) = 0$

$$= -a \cdot P(S_T = -a) + b \cdot \underbrace{[1 - P(S_T = -a)]}_{P(S_T = b)}$$

$$-a P(S_T = -a) - b P(S_T = -a) + b = 0$$

$$\boxed{\mathbb{P}(S_T = -a) = \frac{b}{a+b}}$$

First assume that $p = 1/2$. Then $\mathbf{E}X = 0$ and (S_n) is a martingale. Therefore, by the optional stopping theorem

$$0 = \mathbf{E}S_0 = \mathbf{E}S_\tau = -a\mathbf{P}(S_\tau = -a) + b\mathbf{P}(S_\tau = b) \\ = -a(1 - \mathbf{P}(S_\tau = b)) + b\mathbf{P}(S_\tau = b).$$

Thus

$$\mathbf{P}(S_\tau = b) = \frac{a}{a+b} \quad \text{and} \quad \mathbf{P}(S_\tau = -a) = \frac{b}{a+b}.$$

Using that $(S_n^2 - n)$ is a martingale, we can determine $\mathbf{E}\tau$. Since

opt. stopping $0 = \mathbf{E}(S_0^2 - 0) = \mathbf{E}(S_\tau^2 - \tau)$

we obtain

$$\mathbf{E}\tau = \mathbf{E}(S_\tau^2) = a^2\mathbf{P}(S_\tau = -a) + b^2\mathbf{P}(S_\tau = b) = a^2\frac{b}{a+b} + b^2\frac{a}{a+b} = ab.$$

The case $p \neq 1/2$ is different. Introduce

with respect to the natural filtration $Z_n = s^{S_n} = \prod_{k=1}^n s^{X_k}$

with $s = (1-p)/p = 1/r$. Then (Z_n) is a martingale and

$$Z_\tau = s^b\mathbf{I}(S_\tau = b) + s^{-a}\mathbf{I}(S_\tau = -a) \leq s^b + s^{-a},$$

thus $\mathbf{E}Z_\tau < \infty$ and

$$\liminf_{n \rightarrow \infty} \int_{\{\tau > n\}} |Z_n| d\mathbf{P} \leq (s^b + s^{-a}) \liminf_{n \rightarrow \infty} \mathbf{P}\{\tau > n\} = 0.$$

Again, by the optional sampling theorem

$$\mathbf{E}(Z_0) = \mathbf{E}(Z_\tau) \\ = s^{-a}\mathbf{P}(S_\tau = -a) + s^b(1 - \mathbf{P}(S_\tau = -a)) \\ = s^{-a}\mathbf{P}(S_\tau = -a) + s^b\mathbf{P}(S_\tau = b) \\ = \mathbf{E}(s^{S_\tau}) = \mathbf{E}(Z_\tau) \\ = \mathbf{E}(Z_1) = \mathbf{E}(s^X) = 1.$$

Rearranging we obtain

$$\mathbf{P}(S_\tau = -a) = \frac{1 - s^b}{s^{-a} - s^b} \frac{r^b}{r^b} = \frac{r^b - 1}{r^{a+b} - 1} = \frac{1 - r^b}{1 - r^{a+b}}.$$

$$\mathcal{F}_n = \sigma(X_1, \dots, X_n)$$

$$\mathbf{E}[Z_n | \mathcal{F}_{n-1}] = s^{S_{n-1}}.$$

$$s^{S_{n-1} + X_n} \cdot \mathbf{E}[s^{X_n} | \mathcal{F}_{n-1}]$$

$$= Z_{n-1} \cdot \mathbf{E}(s^X)$$

$$\text{HW } \mathbf{E}(s^X) = 1$$

$$E(X) = P(X=1) - P(X=-1)$$

To obtain $E\tau$, using the Wald identity $E(X) = P - (1-p) = 2p - 1$

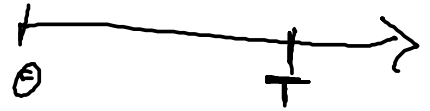
$$ES_\tau = (2p - 1)E\tau,$$

from which

$$E\tau = \frac{1}{2p - 1} ES_\tau = \frac{1}{2p - 1} [-aP(S_\tau = -a) + bP(S_\tau = b)].$$

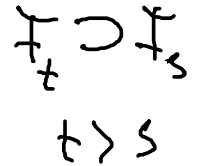
Exercise 4. Show that $\tau < \infty$ a.s.

2 Continuous time martingales



2.1 Definitions and simple properties

Let $(\Omega, \mathcal{F}, \mathbf{P})$ be a probability space and $(\mathcal{F}_t)_{t \geq 0}$ a *filtration*, i.e. an increasing sequence of σ -algebras. The time horizon is either finite or infinite, $t \in [0, T]$ or $t \in [0, \infty)$.



In what follows we always assume that the filtration satisfies the *usual properties*:

(i) \mathcal{F}_0 contains the \mathbf{P} -null sets;

$\cap_{s > t} \mathcal{F}_s = \mathcal{F}_t$ (increasing)

(ii) $(\mathcal{F}_t)_t$ is right-continuous, i.e. $\cap_{s > t} \mathcal{F}_s =: \mathcal{F}_{t+} = \mathcal{F}_t$

Let (X_t) and (Y_t) be stochastic processes. The process Y is a *modification* of X if $X_t = Y_t$ a.s. for any fix t , i.e. $\mathbf{P}(X_t = Y_t) = 1$ for each $t \geq 0$. The processes X and Y are *indistinguishable* if their sample paths are the same almost surely, i.e.

intersection of σ -alg. is σ -alg.

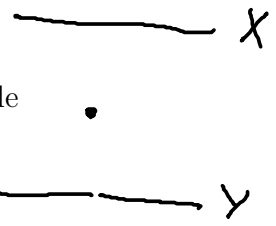
$$\mathbf{P}(X_t = Y_t, t \geq 0) = 1.$$

They have the *same finite dimensional distributions* if for all $0 \leq t_1 < t_2 < \dots < t_n < \infty$ and $A \in \mathcal{B}(\mathbb{R}^n)$

$$\mathbf{P}((X_{t_1}, \dots, X_{t_n}) \in A) = \mathbf{P}((Y_{t_1}, \dots, Y_{t_n}) \in A).$$

Example 4. Let U be uniform(0, 1), and $X_t \equiv 0, t \in [0, 1]$, and $Y_t = \mathbf{I}(U = t)$. Then Y is a modification of X , but they are not indistinguishable, since

$$\mathbf{P}(X_t = Y_t, t \geq 0) = 0.$$



The process $(X_t)_t$ is *adapted to the filtration* $(\mathcal{F}_t)_t$, if X_t is \mathcal{F}_t -measurable for each $t \geq 0$. The process $(X_t, \mathcal{F}_t)_t$ is a *martingale* if

- (i) $(X_t)_t$ is adapted to $(\mathcal{F}_t)_t$;
- (ii) $\mathbf{E}|X_t| < \infty$ for all $t \geq 0$;
- (iii) $\mathbf{E}[X_t | \mathcal{F}_s] = X_s$ a.s. for all $t \geq s$.

It is sub- or supermartingale if (i) and (ii) holds, and (iii) holds with \geq or \leq instead of $=$.

A random variable $\tau : \Omega \rightarrow [0, \infty)$ is a *stopping time* if $\{\tau \leq t\} \in \mathcal{F}_t$. The σ -algebra of the events prior to τ , or pre- τ - σ -algebra is

$$\mathcal{F}_\tau = \{A \in \mathcal{F} : A \cap \{\tau \leq t\} \in \mathcal{F}_t \text{ for all } t \geq 0\}.$$

Exercise 5. Show that \mathcal{F}_τ is indeed a σ -algebra.

The next result is obvious, but very useful.

Proposition 2. Let (X_t, \mathcal{F}_t) be a (sub-, super-) martingale. Then for any sequence $0 \leq t_0 < t_1 < \dots < t_N < \infty$ the process $(X_{t_n}, \mathcal{F}_{t_n})_{n=0}^N$ is a discrete time martingale.

Lemma 3. Let σ, τ be stopping times.

- (i) τ is \mathcal{F}_τ -measurable.
- (ii) If $\tau \equiv t$ then $\mathcal{F}_\tau = \mathcal{F}_t$.
- (iii) $\sigma \wedge \tau = \min(\sigma, \tau)$ and $\sigma \vee \tau = \max(\sigma, \tau)$ are stopping times.
- (iv) If $\sigma \leq \tau$, then $\mathcal{F}_\sigma \subset \mathcal{F}_\tau$.
- (v) If $(X_t)_t$ is right-continuous and adapted then X_τ is \mathcal{F}_τ -measurable.

Exercise 6. Prove the lemma.

Remark 1. In continuous time the technical details are trickier.

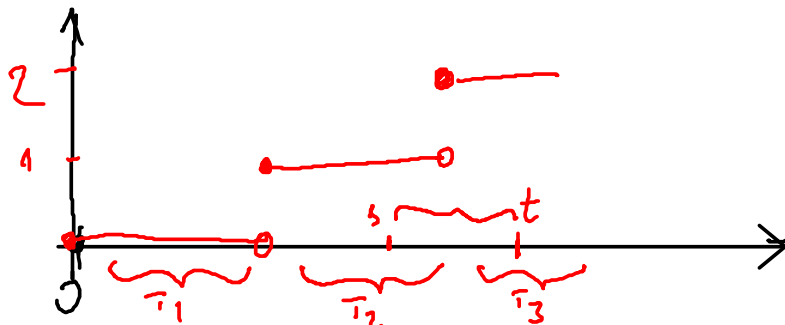
The process $(X_t)_t$ is adapted to $(\mathcal{F}_t)_t$, if X_t is \mathcal{F}_t -measurable, and it is *progressively measurable* with respect to $(\mathcal{F}_t)_t$, if for all $t \geq 0$ and $A \in \mathcal{B}(\mathbb{R}^d)$

$$[0, t] \times \Omega \quad \{(s, \omega) : s \leq t, X_s(\omega) \in A\} \in \mathcal{B}([0, t]) \otimes \mathcal{F}_t,$$

where \mathcal{B} stands for the Borel sets, and \otimes is the product σ -algebra. In what follows we always need progressive measurability, adaptedness is not enough.

The next statement says that the situation is not too bad.

Proposition 3. If $(X_t)_t$ is right continuous and adapted, then it is progressively measurable.



T_1, T_2, T_3, \dots iid
Exp(1)

Example 5 (Poisson process). A Poisson process with intensity $\lambda > 0$ is an adapted integer valued RCLL (right continuous with left limits) process $N = (N_t, \mathcal{F}_t)_{t \geq 0}$ such that

- (i) N has independent increments, that is $N_t - N_s$ is independent of \mathcal{F}_s for any $s < t$,
- (ii) $N_0 = 0$ a.s.,
- (iii) $N_t - N_s \sim \text{Poisson}(\lambda(t - s))$.

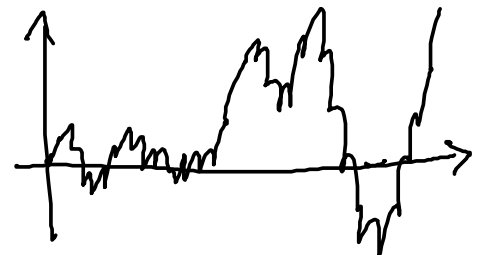
Exercise 7. Show that $(N_t - \lambda t)$ is martingale. ✓

Proposition 4. Let (X_t) be a martingale, and φ a convex function such that $\mathbf{E}|\varphi(X_t)| < \infty$ for all $t \geq 0$. Then $(\varphi(X_t))$ is submartingale.

Furthermore if (X_t) is a submartingale and φ nondecreasing and convex that $\mathbf{E}|\varphi(X_t)| < \infty$ for all $t \geq 0$, then $(\varphi(X_t))$ is a submartingale.

Example 6 (Wiener process). The Wiener process or standard Brownian motion is an adapted process $W = (W_t, \mathcal{F}_t)_{t \geq 0}$ such that

- (i) W has independent increments, that is $W_t - W_s$ is independent of \mathcal{F}_s for any $s < t$,
- (ii) $W_0 = 0$ a.s.,
- (iii) $W_t - W_s \sim N(0, t - s)$, ← Gaussian
- (iv) W_t has continuous sample path.



✓ **Exercise 8.** Show that (W_t) and $(W_t^2 - t)$ are martingales.

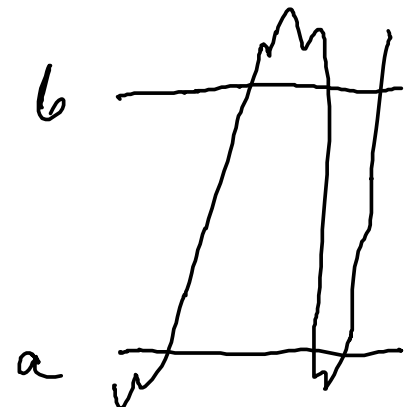
2.2 Martingale convergence theorem

Consider an adapted stochastic process $(X_t)_{t \geq 0}$. Fix $a < b$, and a finite set $F \subset [0, \infty)$. Let U_F denote the number of *upcrossings* of the interval $[a, b]$ by the restricted process $(X_t)_{t \in F}$. Formally, let $\tau_0 = 0$, and

$$\begin{aligned} \tau_{2k-1} &= \min\{t \in F : t \geq \tau_{2k-2}, X_t < a\}, \\ \tau_{2k} &= \min\{t \in F : t \geq \tau_{2k-1}, X_t > b\}. \end{aligned}$$

The number of upcrossings on F is

$$U_F(a, b) = U_F = \max\{k : \tau_{2k} < \infty\}.$$



$(N_t) \sim$ Poisson process $\lambda > 0$

- ind. inc.
- $N_t - N_s \sim \text{Poisson}(\lambda(t-s))$
- $N_0 = 0$ a.s.

$\mathcal{F}_t = \sigma(N_s : s \leq t)$ natural filtration

$N_t - \lambda t$ is wtg.

$s < t$:

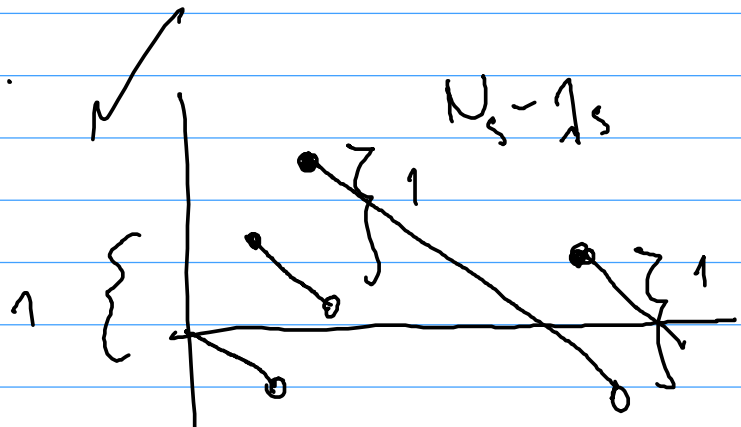
$$E[N_t - \lambda t \mid \mathcal{F}_s] = E[N_s + N_t - N_s - \lambda t \mid \mathcal{F}_s]$$

$$= N_s + E[N_t - N_s - \lambda t \mid \mathcal{F}_s] =$$

indep.

$$= N_s + E(N_t - N_s - \lambda t) = N_s + \lambda(t-s) - \lambda t$$

$$= N_s - \lambda s.$$



(W_t) :

W_t is martingale.

$t > s$

$$E(W_t | \mathcal{F}_s) = E\left[W_s + \overbrace{W_t - W_s}^{\substack{\text{mean } 0 \\ \text{indep. of } \mathcal{F}_s}} \mid \mathcal{F}_s\right]$$

$$= W_s + E(W_t - W_s) = W_s \quad \checkmark$$

$\sim N(0, t-s)$

$\Rightarrow (W_t)$ martingale.

$\Rightarrow (W_t^2)$ submartingale.

$(W_t^2 - t)$ is martingale.

$$E[W_t^2 - t | \mathcal{F}_s] = E\left[W_s^2 + \widehat{W_t^2 - W_s^2} - t \mid \mathcal{F}_s\right]$$

$$= W_s^2 - t + E\left[\underbrace{W_t^2 - W_s^2}_{(W_t - W_s)(W_t + W_s)} \mid \mathcal{F}_s\right]$$

$$= W_s^2 - t + E \left[\underbrace{W_s}_{\text{meas.}} (V_t - W_s) + W_t \cdot (W_t - W_s) | \mathcal{F}_s \right]$$

$$= W_s^2 - t + \underbrace{W_s \cdot E[W_t - W_s | \mathcal{F}_s]}_0 + E \left[\underbrace{(W_s + \underbrace{V_t - W_s}_t)}_t (W_t - W_s) | \mathcal{F}_s \right]$$

$$= W_s^2 - t + \underbrace{W_s \cdot E(W_t - W_s)}_0 + \underbrace{W_s \cdot E(W_t - W_s)}_0 + \underbrace{E((W_t - W_s)^2)}_t$$

$$= W_s^2 - t + 0 + 0 + t - s = W_s^2 - s$$

$\Rightarrow (W_t^2 - t)$ is a martingale. ✓

We can extend the definition of infinite sets $I \subset [0, \infty)$ as

$$U_I = \sup\{U_F : F \subset I, F \text{ finite}\}.$$

We have the upcrossing inequality.

Theorem 6 (Upcrossing inequality). *Let (X_t) be a right-continuous submartingale. For any $a < b$ and $0 \leq S \leq T < \infty$*

$$(b - a)\mathbf{E}U_{[S,T]} \leq \mathbf{E}(X_T - a)^+ - \mathbf{E}(X_S - a)^+.$$

Proof. Consider an enumeration of the countable set $\mathbf{Q} \cap [S, T]$ as

$$\mathbf{Q} \cap [S, T] = \{q_1, q_2, \dots\},$$

and let $F_n = \{q_1, \dots, q_n\} \cup \{S, T\}$. Then $(X_t, \mathcal{F}_t)_{t \in F_n}$ is a discrete time martingale, therefore, by the upcrossing inequality

$$(b - a)\mathbf{E}U_{F_n} \leq \mathbf{E}(X_T - a)^+ - \mathbf{E}(X_S - a)^+.$$

Since F_n is increasing, U_{F_n} is increasing, and by the right-continuity of (X_t)

$$\lim_{n \rightarrow \infty} U_{F_n} = U_{[S,T]} \quad \text{a.s.}$$

In particular, $U_{[S,T]}$ is measurable, and by the monotone convergence theorem the result follows. \square

Theorem 7 (Martingale convergence theorem). *Let (X_t) be a right-continuous submartingale such that*

$$\sup_{t \geq 0} \mathbf{E}(X_t^+) < \infty.$$

Then $\lim_{t \rightarrow \infty} X_t = X$ exists a.s. and $\mathbf{E}|X| < \infty$.

Proof. By the upcrossing inequality and the monotone convergence theorem for any $a < b$

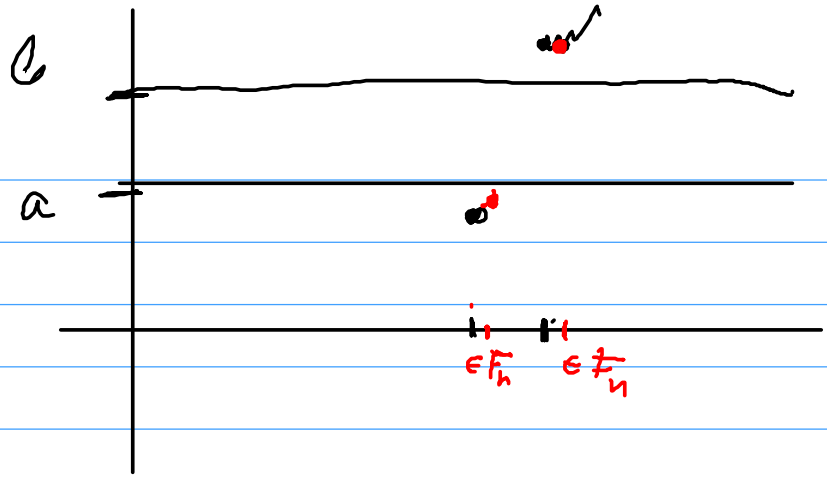
$$\mathbf{E}U_{[0,\infty)}(a, b) \leq \frac{\sup_{t \geq 0} \mathbf{E}X_t^+ + |a|}{b - a}.$$

Therefore, for any $a < b$ the upcrossings $U_{[0,\infty)}(a, b)$ are a.s. finite. Thus almost surely the upcrossings are finite for all $a < b$ rationals, implying the existence of the limit.

The integrability of the limit follows from Fatou's lemma. \square

$$A_{a,b} = \left\{ U_{[0,\infty)}(a,b) = \infty \right\} \quad \mathbb{P}(A_{a,b}) = 0$$

$$\bigcup_{a,b \in \mathbf{Q}} A_{a,b} = A \quad \mathbb{P}(A) = 0.$$



~~2~~

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