# III. Approximation of Distributions of Sums of Weakly Dependent Random Variables by the Normal Distribution

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# Contents

§1.	Weak Dependence Conditions and Covariance Inequalities	114
§2.	Estimation of the Rate of Convergence in the Central Limit Theorem for Weakly Dependent Random Variables	116
	2.1. Introduction and Notation	116 122 128 133 141
§3.	Estimation of the Rate of Convergence in the Central Limit Theorem for Weakly Dependent Random Fields .	146
Refe	erences	150

This article is a review of known results on approximating distributions of sums of weakly dependent random variables (r.v.'s) and random fields defined on the integer lattice  $\mathbb{Z}^d$  by the normal distribution. It also contains some new results of the author. Attention is given mainly to methods of proof due to Stein, Tikhomirov and Heinrich.

# §1. Weak Dependence Conditions and Covariance Inequalities

Let  $X_1, X_2, \ldots$  be a sequence of real r.v.'s,  $\mathcal{F}_a^b$  the  $\sigma$ -algebra generated by  $X_i, a \leq i \leq b, R$  the real line, and  $N = \{1, 2, \ldots\}$ .

Later on we shall assume that the sequence  $X_1, X_2, \ldots$  satisfies one of the following weak dependence conditions:

- 1. m-dependence:  $\mathcal{F}_1^r$  and  $\mathcal{F}_{r'}^{\infty}$  are independent for all integers r and r' such that  $1 \le r < r' < \infty$ , r' r > m;
- 2.  $\psi$ -mixing:

$$\psi(\tau) = \sup_{\substack{t \in N \\ P\{A\} > 0, \ P\{B\} > 0}} \sup_{\substack{P \in \mathcal{F}_{t}^{\omega}, \ B \in \mathcal{F}_{t}^{\omega} + \tau \\ P\{A\} > 0, \ P\{B\} > 0}} |\mathbf{P}\{AB\} - \mathbf{P}\{A\} \mathbf{P}\{B\}| / \mathbf{P}\{A\} \mathbf{P}\{B\} \downarrow 0, \ \tau \to \infty;$$

3. uniformly strong mixing (u.s.m.):

$$\varphi(\tau) = \sup_{t \in N} \sup_{\substack{A \in \mathcal{F}_1^t, \ B \in \mathcal{F}_{t+\tau}^{\infty} \\ \mathbf{P}(A) > 0}} |\mathbf{P}\{AB\} - \mathbf{P}\{A\}\mathbf{P}\{B\}|/\mathbf{P}\{A\} \downarrow 0, \ \tau \to \infty;$$

4. absolute regularity (a.r.):

$$\beta(\tau) = \sup_{t \in N} \mathbf{E} \left( \sup_{B \in \mathcal{F}_{t+\tau}^{\infty}} |\mathbf{P}\{B/\mathcal{F}_{1}^{t}\} - \mathbf{P}\{B\}| \right) \downarrow 0, \ \tau \to \infty;$$

5. strong mixing (s.m.):

$$\alpha(\tau) = \sup_{t \in N} \sup_{A \in \mathcal{F}_1^t, \ B \in \mathcal{F}_{t+\tau}^\infty} |\mathbf{P}\{AB\} - \mathbf{P}\{A\}\mathbf{P}\{B\}| \downarrow 0, \ \tau \to \infty.$$

Strong mixing was introduced by Rosenblatt (1956), absolute regularity by Kolmogorov (Volkonskii and Rozanov (1959)), uniformly strong mixing by Ibragimov (1962),  $\psi$ -mixing by Blum, Hanson and Koopmans (1963) and m-dependence by Hoeffding and Robbins (1948) (and goes back to Bernstein (1926)).

All of these conditions are requirements of weak dependence between the start and end of a sequence of r.v.'s.

Information about these and other measures of weak dependency (and the corresponding mixing coefficients) as well as appropriate references may be found, for example, in Ibragimov and Linnik (1965), Ibragimov and Rozanov (1970), Statulevičius (1974), Philipp and Stout (1975), Bradley, Bryc and Janson (1985), Bradley (1986) and Bulinskii (1987), (1989). The mixing coefficients defined above are interrelated by the following inequalities (Ibragimov and Linnik (1965), Hipp (1979a)):

$$2\alpha(\tau) \leq \varphi^{1/2}(\tau),$$
  

$$\alpha(\tau) \leq \beta(\tau) \leq \varphi(\tau) \leq \psi(\tau),$$
  

$$\alpha(\tau) \leq 1/4, \ \varphi(\tau) \leq 1.$$

Important refinements of these inequalities, a comparison of the mixing coefficients and other measures of dependency are found in Kolmogorov and Rozanov (1960), Ibragimov and Linnik (1965), Ibragimov and Rozanov (1970), Peligrad (1982), Bradley (1983), (1986), Bradley and Bryc (1985), Bradley, Bryc and Janson (1985), Bulinskii (1985b), (1989) and Liptser and Shiryaev (1986). We denote by  $\text{cov}(\xi,\eta) = \mathbf{E}\xi\eta - \mathbf{E}\xi\mathbf{E}\eta$  the covariance of real r.v.'s  $\xi$  and  $\eta$ . We shall subsequently need bounds for the covariance in terms of the mixing coefficients  $\alpha$ ,  $\beta$ ,  $\varphi$  and  $\psi$ .

**Lemma 1.** Suppose that  $\xi$  is  $\mathcal{F}_1^t$ -measurable and  $\eta$  is  $\mathcal{F}_{t+\tau}^{\infty}$ -measurable;  $t, \tau \in \mathbb{N}$ .

1. If  $\mathbf{E}|\xi| < \infty$  and  $\mathbf{E}|\eta| < \infty$ , then (see Philipp (1969a)):

$$|\operatorname{cov}(\xi,\eta)| \le \mathbf{E}|\xi|\mathbf{E}|\eta|\psi(\tau); \tag{1}$$

2. If  $\mathbf{E}|\xi| < \infty$  and  $\mathbf{P}\{|\eta| > C\} = 0$ , then (see Billingsley (1977))

$$|\operatorname{cov}(\xi,\eta)| \le 2C\mathbf{E}|\xi|\varphi(\tau);$$
 (2)

3. If  $\mathbf{E}|\xi|^q < \infty$  and  $\mathbf{E}|\eta|^r < \infty$ , q, r > 1,  $q^{-1} + r^{-1} = 1$ , then (see Ibragimov (1962))

$$|\operatorname{cov}(\xi,\eta)| \le 2\mathbf{E}^{1/q} |\xi|^q \mathbf{E}^{1/r} |\eta|^r \varphi^{1/q}(\tau); \tag{3}$$

4. If  $\mathbf{P}\{|\xi| > C_1\} = \mathbf{P}\{|\eta| > C_2\} = 0$ , then (see Volkonskii and Rozanov (1959), (1961), Ibragimov (1962))

$$|\operatorname{cov}(\xi,\eta)| \le 4C_1 C_2 \alpha(\tau); \tag{4}$$

5. If  $\mathbf{E}|\xi|^r < \infty$ , r > 1, and  $\mathbf{P}\{|\eta| > C\} = 0$ , then (see Davydov (1968), Hipp (1979a))

$$|\operatorname{cov}(\xi,\eta)| < 4C\mathbf{E}^{1/r}|\xi|^r(\alpha(\tau))^{1-r^{-1}};$$
 (5)

6. If  $\mathbf{E}|\xi|^q < \infty$  and  $\mathbf{E}|\eta|^r < \infty$ , q, r > 1,  $q^{-1} + r^{-1} < 1$ , then (see Davydov (1968), Hipp (1979a))

$$|\operatorname{cov}(\xi,\eta)| \le 6\mathbf{E}^{1/q} |\xi|^q \mathbf{E}^{1/r} |\eta|^r (\alpha(\tau))^{1-q^{-1}-r^{-1}}.$$
 (6)

Inequalities (1–6) may be extended easily by induction to finitely many r.v.'s (see, for example, Volkonskii and Rozanov (1959), (1961), Roussas and Ionnides (1987)). We shall need a result that follows from (4) and (5).

**Lemma 2.** Suppose that  $\xi_j$  is  $\mathcal{F}_{s_j}^{t_j}$ -measurable, j = 1, 2, ..., k, where  $1 \le s_1 \le t_1 < s_2 \le t_2 < ... < s_k \le t_k < \infty$ , and that  $\tau = \min_{1 \le j < k} (s_{j+1} - t_j)$ .

1. If  $P\{|\xi_j| > C_j\} = 0$ , j = 1, 2, ..., k, then for  $k \ge 2$ 

$$\left| \mathbf{E} \left( \prod_{j=1}^{k} \xi_j \right) - \prod_{j=1}^{k} \mathbf{E} \xi_j \right| \le 4(k-1) \prod_{j=1}^{k} C_j \alpha(\tau); \tag{7}$$

2. If  $\mathbf{E}|\xi_1|^r < \infty$ , r > 1, and  $\mathbf{P}\{|\xi_j| > C_j\} = 0$ , j = 2, 3, ..., k, then for  $k \ge 2$ 

$$\left| \mathbf{E} \left( \prod_{j=1}^{k} \xi_j \right) - \prod_{j=1}^{k} \mathbf{E} \xi_j \right| \le 4(k-1) \mathbf{E}^{1/r} |\xi_1|^r \prod_{j=2}^{k} C_j(\alpha(\tau))^{1-r^{-1}}.$$
 (8)

Sometimes it is advantageous to have an estimate of the covariance of r.v.'s in terms of the mixing coefficients when the  $\sigma$ -algebras  $\mathcal{F}_1^t$  (the "past") and  $\mathcal{F}_{t+\tau}^{\infty}$  (the "future") are replaced by  $\sigma$ -algebras of a more complex structure (see, for example, Lemma 3).

Let  $\sigma(\mathcal{G} \cup \mathcal{H})$  be the  $\sigma$ -algebra generated by the  $\sigma$ -algebras  $\mathcal{G}$  and  $\mathcal{H}$ .

**Lemma 3** (Takahata (1981)). Let  $\xi$  be  $\mathcal{F}_{t+\tau}^{t+\tau+n-1}$ -measurable and  $\eta$  be  $\sigma\{\mathcal{F}_1^t \cup \mathcal{F}_{t+2\tau+n-1}^{\infty}\}$ -measurable,  $t, \tau, n \in \mathbb{N}$ .

1. If  $\mathbf{E}|\xi|^q < \infty$  and  $\mathbf{E}|\eta|^r < \infty$ , q, r > 0,  $q^{-1} + r^{-1} = 1$ , then

$$|\cos(\xi, \eta)| \le 6\mathbf{E}^{1/q} |\xi|^q \mathbf{E}^{1/r} |\eta|^r \psi^{1/q}(\tau);$$
 (9)

2. If  $\mathbf{E}|\xi|^q < \infty$  and  $\mathbf{E}|\eta|^r < \infty$ , q, r > 0,  $q^{-1} + r^{-1} < 1$ , then

$$|\operatorname{cov}(\xi,\eta)| \le 18\mathbf{E}^{1/q}|\xi|^q \mathbf{E}^{1/r}|\eta|^r (\beta(\tau))^{1-q^{-1}-r^{-1}}.$$
 (10)

If  $\xi$  and  $\eta$  in inequalities (1)–(10) are complex r.v.'s, then the right-hand sides of these inequalities must be multiplied by 4 (see, for example, Ibragimov and Linnik (1965), Roussas and Ionnides (1987)).

§2. Estimation of the Rate of Convergence in the Central Limit Theorem for Weakly Dependent Random Variables

#### 2.1. Introduction and Notation.

We shall concentrate on estimating the rate of convergence in the onedimensional central limit theorem (CLT).

Conditions for the applicability of the CLT to weakly dependent r.v.'s have been studied by many authors.

Devoted specifically to m-dependent r.v.'s are the papers by Hoeffding and Robbins (1948), Diananda (1955), Kallianpur (1955), Orey (1958), Zaremba (1958), and Berk (1973); to functionals defined on Markov chains, the papers by Bernstein (1926), Sirazhdinov (1955), Dobrushin (1956a,b), Nagaev (1957), (1962), Statulevičius (1961), (1969a,b), (1970a), Gudynas (1977), and Lifshits (1984). Ibragimov (1962) is basic as far as strictly stationary sequences of weakly dependent r.v.'s are concerned. The results in that paper were subsequently refined and generalized by Rosén (1967), Serfling (1968), Philipp (1969a), Berk (1973), Ibragimov (1975), Bradley (1981), (1988), Peligrad (1982), (1985), (1986a,b), (1986), (1990), Herrndorff (1983a,b), (1985), Utev

(1984), (1990a,b), Samur (1984), Grin' (1990a,b), etc. Bergström (1970), (1971), (1972) studied this problem by means of his comparison method. The latest results as well as a bibliography on this topic may be found, for example, in the above papers by Peligrad, Bradley, Utev and Grin'. The advances by Utev in (1990a,b) were made by taking advantage of Peligrad's bounds for large deviation probabilities (1985) which were also used in a modified form by Grin' (1990a) (see also (1990b)).

As of today, there exists much research devoted to studying the rate of convergence in the CLT for weakly dependent r.v.'s.

This question was first investigated by Petrov for m-dependent r.v.'s in (1960). He obtained a bound in the uniform metric of order  $O(n^{-(s-2)/(6s-4)})$  under the assumption that the s-th absolute moments of the variables,  $2 < s \le 3$ , are uniformly bounded and the variance of their sum increases linearly; here and elsewhere n is the number of terms. The problem was later investigated for m-dependent r.v.'s by Ibragimov (1967), Egorov (1970), Stein (1972), Erickson (1973)–(1975), Tikhomirov (1980), Shergin (1976), (1979), (1983), Maejima (1978), Yudin (1981), (1989a), Zuparov (1981), Heinrich (1982), (1984), (1985a,b,d), Sunklodas (1982), (1984), (1989), Rhee (1985), (1986a), Zuev (1986) and others. The first optimal estimate, that is, of order  $O(n^{-(s-2)/2})$ , in the uniform metric for m-dependent r.v.'s in the CLT was obtained by Ibragimov (1967). He considered a sequence of r.v.'s of the form  $f(\varepsilon_i, \varepsilon_{i+1}, \ldots, \varepsilon_{i+m-1}), i = 1, 2, \ldots$ , where  $\varepsilon_1, \varepsilon_2, \ldots$  are independent and identically distributed r.v.'s with  $\mathbf{E}|\varepsilon_1|^s < \infty, 2 < s \le 3$ .

The rate of convergence in the one-dimensional CLT has been estimated under other weak dependence conditions by Statulevičius (1962), (1974), (1977a,b), Iosifescu (1968), Philipp (1969b), Stein (1972), Dubrovin (1971), Bulinskii (1977), Yoshihara (1978a), Negishi (1977), Tikhomirov (1976), (1980), Dubrovin and Moskvin (1979), Schneider (1981), Zuparov (1981), Yudin (1984), (1987), (1989b), (1990), (1991), Sunklodas (1977a), (1982), (1984), (1989), Lappo (1986) and others.

These papers made use of diverse methods of proof.

The method of cumulants (semi-invariants) involving the logarithmic derivatives of the characteristic function (c.f.) of the original sum was developed by Statulevičius in (1961), (1969a,b), (1970a) in which he found exact bounds for nonhomogeneous Markov chains. This method was investigated under other weak dependence conditions in (1962), (1974), (1977a,b). Many of the papers mentioned above (see, for example, Petrov (1960), Iosifescu (1968), Philipp (1969b), Egorov (1970) and Bulinskii (1977)) employed methods of proof based on Bernstein's fruitful idea of splitting the sum  $Z_n$  of r.v.'s into two sums  $Z_n = U_n + V_n$ . Then by virtue of the weak dependency,  $U_n$  behaves as if its terms are independent and  $V_n$  has a relatively small variance. For example, when  $Z_n$  is centered and normalized, this was accomplished as a rule via the following inequality (Petrov (1960)): For any positive  $\varepsilon$  and  $x \in R$ ,

$$|\mathbf{P}\{U_n + V_n < x\} - \Phi(x)| \le \sup_{x} |\mathbf{P}\{U_n < x\} - \Phi(x)| + \frac{\varepsilon}{\sqrt{2\pi}} + \mathbf{P}\{|V_n| \ge \varepsilon\},$$
(11)

where  $\Phi(x)$  is the standard normal distribution function. However the estimates obtained in this way for the one-dimensional case are not optimal and the order is no better than  $O(n^{-1/4})$  in the uniform metric.

Nevertheless, Bernstein's classical method has received further development in recent years in the estimation of the rate of convergence in the CLT for weakly dependent n-dimensional and infinite-dimensional r.v.'s. It is precisely in the infinite-dimensional case that this method has manifested its generality. But since we are confining ourselves to only the one-dimensional (that is, the R-valued) case, the reader interested in Bernstein's method in the n-dimensional and infinite-dimensional cases is referred to the papers by Hipp (1979b) and Zuparov (1983), (1984) (see also Dubrovin (1974), Lapinskas (1976), Gabbasov (1977), Sunklodas (1978)). Multi-dimensional and infinite-dimensional analogues of (11) may be found in Lapinskas (1976), Sunklodas (1978), Hipp (1979b) and Zuparov (1984). The passage to independent infinite-dimensional r.v.'s under absolute regularity can be accomplished by means of appropriate approximating inequalities in Gudynas (1989) and Eberlein (1984) (see also Yoshihara (1976), Hipp (1979b)), Zuparov (1983), (1984)).

In (1972), Stein gave a new way of estimating the rate of convergence in the CLT for weakly dependent r.v.'s. Stein's method involves using a linear differential equation in terms of the difference between the distribution function (d.f.) of the sum of weakly dependent r.v.'s and the normal distribution. By means of this method, Stein was able to derive a uniform bound for the rate of convergence in the CLT of order  $O(n^{-1/2} \ln^2 n)$  under complete regularity (this is weaker than u.s.m.), and the optimal order  $O(n^{-1/2})$  for m-dependency. However, he required the eight-order moments of the terms to be finite and that the original sequence of r.v.'s be strictly stationary.

Developing Stein's idea further, Tikhomirov (1980) constructed a similar differential equation for the c.f. of a sum of weakly dependent one-dimensional r.v.'s. Using this, he succeeded in obtaining the wanted bound for how close this c.f. is to the c.f. of the normal distribution. Finally, under minimal restrictions on the moments (the finiteness of the s-th order absolute moments,  $2 < s \le 3$ ), in particular Tikhomirov derived the following uniform estimates for the rate of convergence in the CLT:

- (a) when the s.m. coefficient decreases exponentially, the order is  $O(n^{-(s-2)/2} (\ln n)^{s-1});$
- (b) for m-dependent r.v.'s, the optimal estimate is of order  $O(n^{-(s-2)/2})$ .

Tikhomirov (1980) also obtained a nonuniform estimate for strongly mixing r.v.'s (Theorem 11).

However the class of r.v.'s considered in (1980) is narrowed down by the requirement of strict stationarity. Schneider (1981) considered a uniformly

strong mixing sequence and he replaced strict stationarity and finiteness of the s-th order absolute moments of the variables,  $2 < s \le 3$ , by uniform boundedness of the third-order absolute moments of the variables and linear growth of the variance of their sum. Then assuming the exponential decay of the u.s.m. coefficient, he obtained a uniform estimate for the rate of convergence in the CLT of order  $O(n^{-1/2} \ln n)$ .

Specifically devoted to the m-dependency case are the papers by Shergin (1976), (1979), (1983) and Heinrich (1982), (1984), (1985a,b,d). They employ the c.f. to work out differing ways of estimating the rate of convergence in the one-dimensional CLT to optimal order.

Yudin solved a more general problem in (1981)–(1991). He gives a general method of estimating the rate of convergence of the distributions of sums of weakly dependent r.v.'s to infinitely divisible distributions. His method is also applicable to the normal case. In addition, he studies how to approximate the distributions of sums of weakly dependent r.v.'s by distributions in the class L (1989a,b), (1991). We point out also that Yudin uses Bernstein's method of partitioning a sum and the basic idea behind Tikhomirov's method (1987), (1989)–(1991) when approximating the distributions of sums of weakly dependent r.v.'s by the distributions in class L. These studies are represented in detail in Yudin's book (1990).

Because they yield optimal estimates for the rate of convergence in the onedimensional CLT for weakly dependent r.v.'s under minimal restrictions on the moments of the terms and because they are simple to use and are capable of solving other problems, the methods of Stein, Tikhomirov and Heinrich are currently some of the most extensively used techniques. Therefore we shall give a more detailed presentation below of precisely these three methods. Before doing this, we introduce some notation.

 $\mathcal N$  will denote a real r.v. with standard normal d.f.  $\Phi(x)$  and density  $\varphi(x)=\Phi'(x)$ .

Later on, we shall estimate the difference  $\mathbf{E}h(Z_n) - \mathbf{E}h(\mathcal{N})$ , where  $Z_n$  is a normalized sum of weakly dependent centered r.v.'s and h is either the indicator of an interval or satisfies a smoothness condition.

For any function  $g: R \to R$ , let

$$\mathcal{L}(g; p, lpha) = \sup_{x 
eq y} rac{|g(x) - g(y)|}{|x - y|^{lpha} (1 + |x|^p + |y|^p)}, \ \ ||g||_{\infty} = \sup_{x} |g(x)|,$$

and g' be the derivative of g.

In §2, 2.3. and §3, we shall assume that  $h:R\to R$  satisfies  $||h||_{\infty}<\infty$  (except in Theorems 3 and 16) and one of the conditions  $H_1^{(p)}=\mathcal{L}(h;p,\alpha)<\infty$  or  $H_2^{(p)}=\mathcal{L}(h';p+1,\alpha)<\infty$  with  $p\geq 0$  and  $0<\alpha\leq 1$ .

We now single out the space BL(R) of bounded functions  $h:R\to R$  satisfying a Lipschitz condition, that is, such that

$$||h||_{\infty} < \infty$$
 and  $||h||_{L} = \sup_{x \neq y} \frac{|h(x) - h(y)|}{|x - y|} < \infty.$ 

Write

$$||h||_{\mathcal{B}H_i^{(p)}} = ||h||_{\infty} + H_i^{(p)}, \quad ||h||_{\mathcal{B}L} = ||h||_{\infty} + ||h||_{L}.$$

Let

$$X_1, X_2, \dots \tag{12}$$

be a sequence of real r.v.'s with  $\mathbf{E}X_j = 0$  and  $\mathbf{E}X_j^2 < \infty, j = 1, 2, \dots, n$ . Let

$$S_{n} = \sum_{j=1}^{n} X_{j}, \quad B_{n}^{2} = \mathbf{E}S_{n}^{2}, \quad Z_{n} = S_{n}/B_{n}, \quad F_{n}(x) = \mathbf{P}\{Z_{n} < x\},$$

$$A_{j} = X_{j}/B_{n}, \quad \overline{A}_{j} = A_{j}1_{(|A_{j}| \leq t)}, \quad \overline{\overline{A}}_{j} = A_{j}1_{(|A_{j}| > t)},$$

$$L_{r,n} = \sum_{j=1}^{n} \mathbf{E}|A_{j}|^{r}, \quad \overline{L}_{r,n} = \sum_{j=1}^{n} \mathbf{E}|\overline{A}_{j}|^{r}, \quad \overline{\overline{L}}_{r,n} = \sum_{j=1}^{n} \mathbf{E}|\overline{\overline{A}}_{j}|^{r},$$

$$L_{s,n}^{*} = nd_{s,n}, \quad d_{s,n} = \max_{1 \leq j \leq n} \mathbf{E}|A_{j}|^{s},$$

$$\Delta_{n}(x) = \mathbf{P}\{Z_{n} < x\} - \Phi(x), \quad \Delta_{n} = \sup_{x} |\Delta_{n}(x)|, \quad ||\cdot||_{1} = \int_{-\infty}^{\infty} |\cdot| dx,$$

$$d_{i}^{(p)}(F_{n}, \Phi) = \sup_{h \in \mathcal{H}_{i}^{(p)}} |\mathbf{E}h(Z_{n}) - \mathbf{E}h(\mathcal{N})|/||h||_{BH_{i}^{(p)}},$$

$$d_{BL}(F_{n}, \Phi) = \sup_{h \in BL(R)} |\mathbf{E}h(Z_{n}) - \mathbf{E}h(\mathcal{N})|/||h||_{BL},$$

where  $\mathcal{H}_i^{(p)} = \{h : ||h||_{\infty} < \infty, \ H_i^{(p)} < \infty\}, \ i = 1, 2, \text{ and } 1_A \text{ is the indicator of event } A, t > 0 \text{ and } B_n > 0.$ 

The quantities  $d_i^{(p)}(F_n, \Phi)$  and  $d_{BL}(F_n, \Phi)$  may be expressed as follows:

$$d_i^{(p)}(F_n, \Phi) = \sup_{\substack{h \in \mathcal{H}_i^{(p)} \\ ||h||} BH_i^{(p)} \le 1} |\mathbf{E}h(Z_n) - \mathbf{E}h(\mathcal{N})|,$$

$$d_{BL}(F_n, \Phi) = \sup_{\substack{h \in BL(R) \\ ||h||} BL \le 1} |\mathbf{E}h(Z_n) - \mathbf{E}h(\mathcal{N})|.$$

The quantity  $d_{BL}(F_n, \Phi)$  is known as the bounded Lipschitz distance between the d.f.'s  $F_n$  and  $\Phi$ .

Later on, we shall omit the subscript n in the notation  $L_{r,n}$ ,  $\overline{L}_{r,n}$ ,  $\overline{\overline{L}}_{r,n}$ ,  $L_{s,n}^*$  and  $d_{s,n}$  and instead of  $d_i^{(p)}(F_n,\Phi)$  and  $d_{BL}(F_n,\Phi)$  we shall write  $d_i^{(p)}$  and  $d_{BL}$ .

The letter  $C(\cdot)$  with or without a subscript will denote a finite positive constant (not always the same one) that depends on the quantities indicated in the parentheses. C is an absolute positive constant;  $\Theta$  is a complex function not exceeding one in modulus;  $0 < K < \infty$  and  $\lambda > 0$  are constants.

We shall show later how the methods of Stein, Tikhomirov and Heinrich yield upper bounds for  $||\Delta_n(x)||_1$ ,  $d_i^{(p)}$ ,  $d_{BL}$  and  $\Delta_n$  if the sequence (12) satisfies one of the above weak dependence conditions.

The bounds for  $||\Delta_n(x)||_1$  (Theorem 1 and Corollary 1),  $d_i^{(p)}$  (Theorems 3–8 and Corollaries 2, 3) and  $\Delta_n$  (Theorem 10) for the mixing coefficients described above are new and are due to the author. More general results are contained in Propositions 2–4.

To describe the methods of Stein and Tikhomirov, we shall make use of the r.v.'s

$$Z_j^{(i)} = \sum_{|p-j| \le im} A_p$$
 and  $z_j^{(i)} = Z_n - Z_j^{(i)}$ 

with 2mi+1 < n and m = 1, 2, ...; in the case of m-dependence, m = 0, 1, .... The bounds derived by Stein's method make use of a subsidiary r.v. J which is uniformly distributed on the set  $\{1, 2, ..., n\}$  and independent of the r.v.'s  $X_1, X_2, ..., X_n$ .

Let

$$c_* = \begin{cases} m+1 & \text{under $m$-dependency,} \\ 1+2\psi_n & \text{under $\psi$-mixing,} \\ 1+4\Phi_n & \text{under u.s.m.,} \\ 1+12B_{n,r} & \text{under a.r.,} \\ 1+12A_{n,r} & \text{under s.m.} \end{cases}$$

where 
$$\psi_n = \sum_{\tau=1}^{n-1} \psi(\tau)$$
,  $\Phi_n = \sum_{\tau=1}^{n-1} \varphi^{1/2}(\tau)$ ,  $B_{n,r} = \sum_{\tau=1}^{n-1} (\beta(\tau))^{(r-2)/r}$ ,  $A_{n,r} = \sum_{\tau=1}^{n-1} (\alpha(\tau))^{(r-2)/r}$  and  $2 < r < \infty$ .

Applying the respective inequalities (1), (3) and (6), we find when the sequence (12) is either m-dependent,  $\psi$ -mixing or u.s.m. that

$$\mathbf{E}Z_n^2 \le c_* L_2. \tag{13}$$

For a.r. or s.m.,

$$\mathbf{E}Z_n^2 \le c_* \sum_{j=1}^n \mathbf{E}^{2/r} |A_j|^r, \quad 2 < r < \infty.$$
 (14)

However, for  $2 < r < \infty$ .

$$\sum_{j=1}^{n} \mathbf{E}^{2/r} |A_j|^r \le n^{(r-2)/r} L_r^{2/r}. \tag{15}$$

Therefore for the weakly dependent r.v.'s defined above in (12) with  $\psi_n$ ,  $\Phi_n$ ,  $B_{n,r}$  and  $A_{n,r}$  respectively finite, the estimates (13)–(15) imply that

$$n^{-(r-2)/2} \le c_*^{r/2} L_r \tag{16}$$

for  $2 < r < \infty$ . Consequently if  $C_1(\cdot)(m+1) \ge n$ , then

$$1 \le (C_1(\cdot))^{(r-2)/2} c_*^{r/2} (m+1)^{(r-2)/2} L_r. \tag{17}$$

By virtue of (17), it suffices to consider the case where  $C_1(\cdot)(m+1) < n$  when estimating  $||\Delta_n(x)||_1$ ,  $d_i^{(p)}$ ,  $d_{BL}$  and  $\Delta_n$ .

### 2.2. Estimation of $||\Delta_n(x)||_1$ .

**Theorem 1.** Suppose that the sequence (12) is m-dependent. Then for  $m \ge 0$  and t > 0,

$$||\Delta_n(x)||_1 \le C\{\overline{\overline{L}}_1 + (m+1)\overline{\overline{L}}_2 + (m+1)^2\overline{L}_3 + (m+1)^3\overline{L}_4\}.$$

If the truncation level in this estimate is taken to be t=1 and  $\overline{\overline{L}}_1$  is replaced by  $(m+1)\overline{\overline{L}}_1$ , then one obtains Erickson's result (1974) (see also Sunklodas (1982)).

Corollary 1. Let the hypotheses of Theorem 1 hold and let  $\mathbf{E}|X_j|^s < \infty$ ,  $2 < s \le 3$  and j = 1, 2, ..., n. Then for  $m \ge 0$ ,

$$||\Delta_n(x)||_1 \le C(m+1)^{s-1}L_s.$$

By means of (24) below, similar bounds may be obtained for  $||\Delta_n(x)||_1$  in the case of the mixing coefficients  $\psi$ ,  $\varphi$  and  $\beta$ . Bounds were found by Sunklodas (1982), (1986) and Takahata (1983) for other mixing coefficients. For the sake of simplifying the presentation, we confine ourselves here to estimating  $||\Delta_n(x)||_1$  just for m-dependent r.v.'s.

Heinrich's results for m-dependent r.v.'s (1985d) imply this particular one.

**Theorem 2** (Heinrich (1985d)). Suppose that the sequence (12) is m-dependent and that  $\mathbf{E}|X_j|^s < \infty$  with 2 < s < 3 and j = 1, 2, ..., n. Then for  $1 \le p < \infty$  and  $n \ge 1$ ,

$$\left(\int_{-\infty}^{\infty} (1+|x|)^{sp-1} |\Delta_n(x)|^p dx\right)^{1/p} \le C(m+1)^{s-1} L_s^*.$$

Shergin (1983) considered m-dependent r.v.'s (12) with finite k-th absolute moments,  $k \geq 2$ , such that  $\sum_{j=1}^n \mathbf{E} X_j^2 \leq M_0 B_n^2$  for all  $n \geq n_0$ , where  $M_0$  and  $n_0$  are positive constants. He used Stein's technique to derive a bound for  $\int_{-\infty}^{\infty} |x|^{\ell} |\Delta_n(x)| dx$  in terms of Lyapunov's quotients for  $0 \leq \ell \leq k-1$ .

#### Proof of Theorem 1.

Stein pointed out the following characterization of the standard normal law: If the r.v.  $W = \mathcal{N}$ , then  $\mathbf{E}f'(W) - \mathbf{E}[Wf(W)] = 0$  for a fairly broad class of functions f, and conversely (see Stein (1972), (1981)). Consequently, the quantity  $\mathbf{E}f'(W) - \mathbf{E}[Wf(W)]$  is a good measure of the proximity of the r.v. W to  $\mathcal{N}$ .

Therefore to estimate the difference  $\mathbf{E}h(W) - \mathbf{E}h(\mathcal{N})$ , it suffices to estimate the right-hand side of the relation

$$\mathbf{E}h(W) - \mathbf{E}h(\mathcal{N}) = \mathbf{E}f'(W) - \mathbf{E}[Wf(W)).$$

To this end, consider the linear differential equation

$$f'(y) - yf(y) = h(y) - \mathbf{E}h(\mathcal{N}). \tag{18}$$

Its solution is

$$f(y) = e^{y^2/2} \int_{-\infty}^{y} h_0(u) e^{-u^2/2} du = -e^{y^2/2} \int_{y}^{\infty} h_0(u) e^{-u^2/2} du, \qquad (19)$$

where  $h_0(y) = h(y) - \mathbf{E}h(\mathcal{N})$ .

For instance, to estimate  $\mathbf{P}\{W < x\} - \Phi(x)$ , it is necessary to take h in (18) to be the indicator function of the interval, i.e.,  $h(y) = h_x(y) = 1_{(-\infty,x)}(y)$ . Then (18) and (19) become respectively

$$f_x'(y) - y f_x(y) = 1_{(-\infty, x)}(y) - \Phi(x)$$
(20)

and

$$f_x(y) = \begin{cases} \Phi(-x)\Phi(y)/\varphi(y), & \text{if } y < x, \\ \Phi(x)\Phi(-y)/\varphi(y), & \text{if } y \ge x. \end{cases}$$
 (21)

To estimate  $||\Delta_n(x)||_1$ , we make use of the following result.

**Lemma 4** (Erickson (1974), Ho and Chen (1978)). Let  $f_x$  be given by (21). Then for all real y,

$$\int_{-\infty}^{\infty} |f_x(y)| dx = 1 \quad and \quad \int_{-\infty}^{\infty} |f'_x(y)| dx \le 1.$$

In order that all of the r.v.'s occurring in (22) and (24) below should be well defined, it is assumed further that 6m + 1 < n.

For brevity, put  $Q_j = A_j Z_j^{(1)} - \mathbf{E}(A_j Z_j^{(1)})$ . It follows from (20) that

$$\Delta_n(x) = \mathbf{E} \left\{ f_x'(Z_n) - \sum_{j=1}^n A_j f_x(Z_n) \right\}$$

$$= \mathbf{E} \left\{ f_x'(Z_n) + \sum_{j=1}^n A_j \left[ f_x(z_j^{(1)}) - f_x(Z_n) \right] - \sum_{j=1}^n A_j f_x(z_j^{(1)}) \right\}.$$

Applying the Newton-Leibniz formula to the differences  $f_x(z_j^{(1)}) - f_x(Z_n)$  and  $f_x(u) - f_x(Z_n)$  and noting (20), we find that

$$\Delta_{n}(x) = \mathbf{E}f'_{x}(Z_{n}) + \sum_{j=1}^{n} \mathbf{E} \left\{ A_{j} \int_{Z_{n}}^{z_{j}^{(1)}} u \int_{Z_{n}}^{u} f'_{x}(v) \, dv \, du \right\}$$

$$+ \frac{1}{2} \sum_{j=1}^{n} \mathbf{E} \left\{ A_{j} (Z_{j}^{(1)})^{2} f_{x}(Z_{n}) \right\} - \sum_{j=1}^{n} \mathbf{E} \left\{ A_{j} f_{x}(z_{j}^{(1)}) \right\}$$

$$+ \sum_{j=1}^{n} \mathbf{E} \left\{ A_{j} \int_{Z_{n}}^{z_{j}^{(1)}} [1_{(-\infty,x)}(u) - 1_{(-\infty,x)}(Z_{n})] du \right\}$$

$$- \sum_{j=1}^{n} \mathbf{E} \left\{ A_{j} Z_{j}^{(1)} f'_{x}(Z_{n}) \right\}.$$

Similar reasoning leads to

$$\sum_{j=1}^{n} \mathbf{E} \{ Q_{j} [f'_{x}(Z_{n}) - f'_{x}(z_{j}^{(2)})] \} = \sum_{j=1}^{n} \mathbf{E} \{ Q_{j} Z_{j}^{(2)} f_{x}(Z_{n}) \}$$

$$+ \sum_{j=1}^{n} \mathbf{E} \left\{ Q_{j} z_{j}^{(2)} \int_{z_{j}^{(2)}}^{Z_{n}} f'_{x}(u) du \right\}$$

$$+ \sum_{j=1}^{n} \mathbf{E} \{ Q_{j} [1_{(-\infty,x)}(Z_{n}) - 1_{(-\infty,x)}(z_{j}^{(2)})] \}.$$

Since  $\mathbf{E}Z_n^2 = 1$ , we have

$$1 - \sum_{j=1}^{n} A_j Z_j^{(1)} = -\sum_{j=1}^{n} Q_j + \sum_{j=1}^{n} \mathbf{E} \{ A_j z_j^{(1)} \}.$$

From the last three equalities, we conclude that (Sunklodas (1982))

$$\Delta_n(x) = \sum_{k=1}^{9} E_k(x), \tag{22}$$

where

$$E_{1}(x) = \sum_{j=1}^{n} \mathbf{E} \left\{ A_{j} \int_{Z_{n}}^{z_{j}^{(1)}} u \int_{Z_{n}}^{u} f'_{x}(v) \, dv \, du \right\},$$

$$E_{2}(x) = \sum_{j=1}^{n} \mathbf{E} \left\{ Q_{j} z_{j}^{(2)} \int_{Z_{n}}^{z_{j}^{(2)}} f'_{x}(u) \, du \right\},$$

$$E_{3}(x) = \frac{1}{2} \sum_{j=1}^{n} \mathbf{E} \{ A_{j} (Z_{j}^{(1)})^{2} f_{x}(Z_{n}) \},$$

$$E_{4}(x) = -\sum_{j=1}^{n} \mathbf{E} \{ Q_{j} Z_{j}^{(2)} f_{x}(Z_{n}) \},$$

$$E_{5}(x) = \sum_{j=1}^{n} \mathbf{E} \left\{ A_{j} \int_{Z_{n}}^{z_{j}^{(1)}} [1_{(-\infty,x)}(u) - 1_{(-\infty,x)}(Z_{n})] du \right\},$$

$$E_{6}(x) = -\sum_{j=1}^{n} \mathbf{E} \left\{ Q_{j} [1_{(-\infty,x)}(Z_{n}) - 1_{(-\infty,x)}(z_{j}^{(2)})] \right\},$$

$$E_{7}(x) = -\sum_{j=1}^{n} \mathbf{E} \{ A_{j} f_{x}(z_{j}^{(1)}) \},$$

$$E_{8}(x) = -\sum_{j=1}^{n} \mathbf{E} \{ Q_{j} f_{x}'(z_{j}^{(2)}) \},$$

$$E_{9}(x) = \sum_{j=1}^{n} \mathbf{E} \{ A_{j} z_{j}^{(1)} \} \mathbf{E} f'(Z_{n}).$$

It is precisely (22) that plays a basic role in estimating  $||\Delta_n(x)||_1$  for weakly dependent r.v.'s. A similar relation was used by Erickson (1974) for m-dependent r.v.'s.

Applying Lemma 4 and the fact that

$$\int_{-\infty}^{\infty} |1_{(-\infty,x)}(u) - 1_{(-\infty,x)}(v)| dx = |u - v|, \tag{23}$$

for any  $u, v \in R$ , we can deduce the next result from (22). Basic Inequality (Sunklodas (1982)):

$$||\Delta_n(x)||_1 \le I_1(m) + I_2(m)$$
 (24)

with

$$\begin{split} I_{1}(m) &= \frac{5}{6} n \mathbf{E} |A_{J}(Z_{J}^{(1)})^{3}| + n \mathbf{E} |A_{J}(Z_{J}^{(1)})^{2}| + 2n \mathbf{E} |Q_{J}^{(2)}Z_{J}^{(2)}| + \\ &+ \frac{n}{2} \mathbf{E} |A_{J}(Z_{J}^{(1)})^{2} (Z_{J}^{(2)} - Z_{J}^{(1)})| + n \mathbf{E} |Q_{J}Z_{J}^{(2)}(Z_{J}^{(3)} - Z_{J}^{(2)})| \end{split}$$

and

$$I_{2}(m) = \frac{n}{2} \mathbf{E} |A_{J}(Z_{J}^{(1)})^{2} z_{J}^{(2)}| + n \mathbf{E} |Q_{J} Z_{J}^{(2)} z_{J}^{(3)}| +$$

$$+ \sum_{j=1}^{n} [||\mathbf{E} \{A_{j} f_{x}(z_{j}^{(1)})\}||_{1} + ||\mathbf{E} \{Q_{j} f_{x}^{\prime}(z_{j}^{(2)})\}||_{1}] + n |\mathbf{E} \{A_{J} z_{J}^{(1)}\}|.$$

Since  $I_1(m)$  involves the moments of "close" r.v.'s, its estimation does not require the use of the weak dependence of  $X_1, X_2, \ldots$ 

It is easy to see that for any real  $r \geq 1$ 

$$\mathbf{E}|Z_I^{(i)}|^r \le (2mi+1)^r n^{-1} L_r, \quad i = 0, 1, \dots,$$
(25)

and that

$$\mathbb{E}|Z_I^{(i)} - Z_I^{(i-1)}|^r \le (2m)^r n^{-1} L_r, \quad i = 2, 3, \dots$$
 (26)

Therefore the application of Hölder's inequality and the estimates (25) and (26) yield

$$|I_1(m)| \le (2m+1)(18m+5)L_3 + \frac{1}{6}(2m+1)(128m^2 + 50m + 5)L_4.$$
 (27)

 $I_2(m)$  has to be estimated separately for each weak dependence condition. Let the sequence (12) be m-dependent. Then

$$I_2(m) = \frac{n}{2} \mathbf{E} |A_J(Z_J^{(1)})^2 z_J^{(2)}| + n \mathbf{E} |Q_J Z_J^{(2)} z_J^{(3)}|.$$

Put

$$t_j^i = \mathbf{E}|Z_j^{(2)}|^i, \quad \delta_j^i = \mathbf{E}|Z_j^{(3)}|^i.$$

Then

$$\begin{aligned} \mathbf{E}|A_{J}(Z_{J}^{(1)})^{2}z_{J}^{(2)}| &\leq \mathbf{E}|A_{J}(Z_{J}^{(1)})^{2}| + \mathbf{E}|A_{J}(Z_{J}^{(1)})^{2}t_{J}^{1}|, \\ \mathbf{E}|Q_{J}Z_{J}^{(2)}z_{J}^{(3)}| &\leq \mathbf{E}|Q_{J}Z_{J}^{(2)}| + \mathbf{E}|Q_{J}Z_{J}^{(2)}\delta_{J}^{1}|. \end{aligned}$$

If  $\xi_j = \sum_{p \in B_j} A_p$  and  $\tau_j^i = \mathbf{E} |\xi_j|^i$ , where  $B_j \subset \{1, 2, \dots, n\}$  for any  $j = 1, 2, \dots, n$ , it is easy to see that

$$\mathbf{E}|\tau_J^i|^{j/i} \le \mathbf{E}|\xi_J|^j, \quad 0 < i \le j. \tag{28}$$

Then (25), (26), (28) and Hölder's inequality imply that

$$\begin{aligned} \mathbf{E}|A_{J}(Z_{J}^{(1)})^{2}| &\leq (2m+1)^{2}n^{-1}L_{3}, \\ \mathbf{E}|A_{J}(Z_{J}^{(1)})^{2}t_{J}^{1}| &\leq (2m+1)^{2}(4m+1)n^{-1}L_{4}, \\ \mathbf{E}|Q_{J}Z_{J}^{(2)}| &\leq 2(2m+1)(4m+1)n^{-1}L_{3}, \\ \mathbf{E}|Q_{J}Z_{J}^{(2)}\delta_{J}^{1}| &\leq 2(2m+1)(4m+1)(6m+1)n^{-1}L_{4}. \end{aligned}$$

Consequently,

$$I_2(m) \le (2m+1)[(9m+2.5)L_3 + (52m^2 + 23m + 2.5)L_4).$$
 (29)

Therefore substituting (27) and (29) in (24) and dropping the condition 6m + 1 < n by means of (17), we obtain the following result.

**Proposition 1.** Suppose that the sequence (12) is m-dependent and  $\mathbf{E}X_j^4 < \infty$ , j = 1, 2, ..., n. Then for  $m \geq 0$ ,

$$||\Delta_n(x)||_1 \le (2m+1)[(27m+7.5)L_3 + (10/3)(22m^2 + 9.4m + 1)L_4).$$

Truncation. To complete the proof of Theorem 1, we truncate the r.v.'s  $A_j$ , j = 1, 2, ..., n, at the level t > 0 (§2, 2.1); in addition we put

$$\overline{A}_j^{(0)} = \overline{A}_j - \mathbf{E}\overline{A}_j, \quad \overline{\overline{A}}_j^{(0)} = \overline{\overline{A}}_j - \mathbf{E}\overline{\overline{A}}_j, \quad \overline{Z}_n^{(0)} = \sum_{j=1}^n \overline{A}_j^{(0)}.$$

Since any pair of r.v.'s  $\xi$  and  $\eta$  satisfies the inequality (see Erickson (1974), p. 527)

$$||\mathbf{P}\{\xi < x\} - \mathbf{P}\{\eta < x\}||_1 \le \mathbf{E}|\xi - \eta|,\tag{30}$$

it follows that

$$||\Phi(ax) - \Phi(x)||_1 \le \sqrt{2/\pi} |1 - a^{-1}| \tag{31}$$

for any positive a.

Therefore when  $\mathbf{E}(\overline{Z}_n^{(0)})^2 > 0$ , the sequence (12) satisfies

$$||\Delta_{n}(x)||_{1} \leq 2\overline{\overline{L}}_{1} + \sqrt{2/\pi} |1 - \mathbf{E}(\overline{Z}_{n}^{(0)})^{2}| + \mathbf{E}^{1/2}(\overline{Z}_{n}^{(0)})^{2}||\mathbf{P}\{\overline{Z}_{n}^{(0)} < x\mathbf{E}^{1/2}(\overline{Z}_{n}^{(0)})^{2}\} - \Phi(x)||_{1}.$$
(32)

We note that the truncation inequality (32) holds for any dependency of the sequence (12).

Since

$$|1 - \mathbf{E}(\overline{Z}_n^{(0)})^2| \le 9\overline{\overline{L}}_2 + \sum_{1 \le i \ne j \le n} |\mathbf{E}(\overline{\overline{A}}_i^{(0)} \overline{\overline{A}}_j^{(0)}) + 2\mathbf{E}(\overline{A}_i^{(0)} \overline{\overline{A}}_j^{(0)})|, \tag{33}$$

any m-dependent sequence (12) satisfies

$$|1 - \mathbf{E}(\overline{Z}_n^{(0)})^2| < 9(2m+1)\overline{\overline{L}}_2 = \varepsilon_1. \tag{34}$$

To prove Theorem 1, it is assumed that  $\varepsilon_1 \leq 1/2$  (otherwise the estimates are trivial). From (34), we have that  $\frac{1}{2} \leq \mathbf{E}(\overline{Z}_n^{(0)})^2 \leq \frac{3}{2}$  and consequently,

$$\sum_{j=1}^{n} \mathbf{E} |\overline{A}_{j}^{(0)} / \mathbf{E}^{1/2} (\overline{Z}_{n}^{(0)})^{2}|^{r} \le 2^{3r/2} \overline{L}_{r}$$
(35)

for positive r.

By virtue of the truncation inequality (32) and the bounds (34) and (35), to complete the proof of Theorem 1, it is sufficient to estimate  $||\Delta_n(x)||_1$  for untruncated, r.v.'s. In other words, we make use of Proposition 1.

Corollary 1 follows from Theorem 1 with  $t = (m+1)^{-1}$ .

# 2.3. Estimation of $d_i^{(p)}$ and $d_{BL}$ .

**Theorem 3.** Suppose that the sequence (12) is m-dependent with  $\mathbf{E}|X_j|^{2+p+\alpha} < \infty$ , j = 1, 2, ..., n. If 6m + 1 < n, then

$$|\mathbf{E}h(Z_n) - \mathbf{E}h(\mathcal{N})| \le C(p,\alpha)H_i^{(p)}\{(m+1)^{1+\alpha}L_{2+\alpha}(1+\mathbf{E}|Z_n|^p) + (m+1)^{1+p+\alpha}L_{2+p+\alpha}\}.$$

Thus by Theorem 3, the estimation of  $|\mathbf{E}h(Z_n) - \mathbf{E}h(\mathcal{N})|$ , where h satisfies either the condition  $H_1^{(p)} < \infty$  or  $H_2^{(p)} < \infty$ , reduces to estimating the absolute moment  $\mathbf{E}|Z_n|^p$ ,  $p \geq 0$ .

Bounds for the absolute moments of a sum of weakly dependent r.v.'s were obtained by Utev (1984). For  $\psi$ -mixing or absolutely regular r.v.'s, estimating  $d_i^{(p)}$ , i=1,2, can also be reduced to estimating the absolute moments of  $Z_n$  (although of a higher order than p). However, for the sake of simplicity, we shall only consider here the estimation of  $d_i^{(0)}$ , i=1,2.

**Theorem 4.** Let the sequence (12) be  $\psi$ -mixing with  $\psi(\tau) \leq K\tau^{-\mu}$ , where  $\mu \geq (r-1)r$  and let  $\mathbf{E}|X_j|^r < \infty$ ,  $4 \leq r < \infty$  and j = 1, 2, ..., n. Then

$$d_i^{(0)} \le C(K, \mu, r, \alpha) \{ n^{(1+\alpha)(r-1)/\mu} L_{2+\alpha} + L_r \}.$$

**Theorem 5.** Let the sequence (12) be absolutely regular and let  $\mathbf{E}|X_j|^r < \infty$ ,  $4 \le r < \infty$  and j = 1, 2, ..., n.

1. If  $\beta(\tau) \leq Ke^{-\lambda \tau}$ , then

$$d_i^{(0)} \le C(K, \lambda, r, \alpha) \{ L_{2+\alpha} \ln^{1+\alpha} (n+1) + L_r \}.$$

2. If  $\beta(\tau) \leq K\tau^{-\mu}$  with  $\mu \geq 2(r-1)r$ , then

$$d_i^{(0)} \le C(K, \mu, r, \alpha) \{ n^{2(1+\alpha)(r-1)/\mu} L_{2+\alpha} + L_r \}.$$

Truncation of the r.v.'s leads, for example, to the following results.

**Theorem 6.** Let the sequence (12) be m-dependent. Then for  $m \geq 0$  and t > 0

$$d_1^{(0)} \le C(\alpha) \{ \overline{\overline{L}}_1^{\alpha} + (m+1)^{\alpha} \overline{\overline{L}}_2^{\alpha} + (m+1)^{1+\alpha} \overline{L}_{2+\alpha} \}.$$

Corollary 2. Let the hypotheses of Theorem 6 hold.

1. If  $\mathbf{E}|X_j|^{2+\alpha} < \infty$ , j = 1, 2, ..., n, then for  $m \ge 0$ 

$$d_1^{(0)} \le C(\alpha)(m+1)^{\alpha(1+\alpha)} L_{2+\alpha}^{\alpha}.$$

2. If  $\mathbf{E}|X_j|^s < \infty$ ,  $2 < s \le 3$  and  $j = 1, 2, \dots, n$ , then for  $m \ge 0$ 

$$d_{BL} \le C(m+1)^{s-1} L_s.$$

**Theorem 7.** Suppose that (12) is  $\psi$ -mixing with  $\psi(\tau) \leq Ke^{-\lambda \tau}$ . Then

$$d_1^{(0)} \le C(K, \lambda, r, \alpha) \{ \overline{\overline{L}}_1^{\alpha} + \overline{\overline{L}}_2^{\alpha} + \overline{L}_{2+\alpha} \ln^{1+\alpha} (n+1) + \overline{L}_r \}$$

for t > 0 and  $4 \le r < \infty$ .

Corollary 3. Suppose that the hypotheses of Theorem 7 hold.

1. If  $\mathbf{E}|X_j|^{2+\alpha} < \infty$ , j = 1, 2, ..., n, then

$$d_1^{(0)} \le C(K, \lambda, \alpha) L_{2+\alpha}^{\alpha} \ln^{\alpha(1+\alpha)}(n+1).$$

2. If  $\mathbf{E}|X_j|^s < \infty$ ,  $2 < s \le 3$  and j = 1, 2, ..., n, then

$$d_{BL} \le C(K,\lambda)L_s \ln^{s-1}(n+1).$$

**Theorem 8.** Suppose that (12) is  $\psi$ -mixing with  $\psi(\tau) \leq K\tau^{-\mu}$ , where  $\mu \geq 12$ .

1. If  $\mathbf{E}|X_j|^{2+\alpha} < \infty$ , j = 1, 2, ..., n, then

$$d_1^{(0)} \le C(K, \mu, \alpha) n^{3(1+\alpha)/\mu} L_{2+\alpha}^{\alpha}$$

2. If  $\mathbf{E}|X_j|^s < \infty$ ,  $2 < s \le 3$  and j = 1, 2, ..., n, then

$$d_{BL} \le C(K,\mu) n^{3(s-1)/\mu} L_s.$$

The reader can learn more details about estimating with the bounded Lipschitz metric  $d_{BL}$  by consulting Sunklodas (1989).

Proofs of Theorems 3–8.

The following result is true.

**Lemma 5.** Let f and f' be given by (18) and (19).

1. If  $||h||_{\infty} < \infty$ , then (see Erickson (1974), Barbour and Eagleson (1985))

$$|f(y)| \le c_1 ||h_0||_{\infty} \quad and \quad |f'(y)| \le 2||h_0||_{\infty}$$
 (36)

for  $\forall y \in R$ , where  $c_1 = \sup_{x>0} \Xi(x)$  and  $\Xi(x) = \Phi(-x)/\varphi(x)$ .

2. If  $H_1^{(p)} < \infty$  or  $H_2^{(p)} < \infty$ , then (Barbour (1986))

$$\mathcal{L}(f'; p, \alpha) \le C_i(p, \alpha) H_i^{(p)}, \tag{37}$$

where i = 1 if  $H_1^{(p)} < \infty$  and i = 2 if  $H_2^{(p)} < \infty$ .

Furthermore, let 6m + 1 < n. Then (18) implies that

$$\mathbf{E}h_0(Z_n) = \sum_{j=1}^n \{ \mathbf{E}(A_j Z_j^{(1)}) \mathbf{E}f'(Z_n) - \mathbf{E}[A_j f(Z_n)] + \mathbf{E}(A_j Z_j^{(1)}) \mathbf{E}f'(Z_n) \}.$$

In the summation, we first add and subtract  $\mathbf{E}(A_j Z_j^{(1)}) \mathbf{E} f'(z_j^{(1)})$ ,  $\mathbf{E}[A_j f(z_j^{(1)})]$  and  $\mathbf{E}[A_j Z_j^{(1)} f'(z_j^{(1)})]$  and then  $\mathbf{E}[Q_j f'(z_j^{(2)})]$ . Then using the identity

$$f(Z_n) - f(z_j^{(1)}) - f'(z_j^{(1)})Z_j^{(1)} = \int_0^{Z_j^{(1)}} [f'(z_j^{(1)} + u) - f'(z_j^{(1)})]du,$$

we find that

$$\mathbf{E}h(Z_n) - \mathbf{E}h(\mathcal{N}) = I_1 + I_2, \tag{38}$$

where

$$I_{1} = \sum_{j=1}^{n} \mathbf{E}(A_{j}Z_{j}^{(1)})\mathbf{E}[f'(Z_{n}) - f'(z_{j}^{(1)})]$$

$$-\sum_{j=1}^{n} \mathbf{E}\{Q_{j}[f'(z_{j}^{(1)}) - f'(z_{j}^{(2)})]\}$$

$$-\sum_{j=1}^{n} \mathbf{E}\left\{A_{j}\int_{0}^{Z_{j}^{(1)}}[f'(z_{j}^{(1)} + u) - f'(z_{j}^{(1)})]du\right\}$$

and

$$I_{2} = -\sum_{j=1}^{n} \mathbf{E}\{Q_{j}f'(z_{j}^{(2)})\} + \sum_{j=1}^{n} \mathbf{E}(A_{j}z_{j}^{(1)})\mathbf{E}f'(Z_{n})$$
$$-\sum_{j=1}^{n} \mathbf{E}[A_{j}f(z_{j}^{(1)})].$$

Put

$$\nu_j^i = \mathbf{E}|Z_j^{(1)}|^i, \ \gamma_j^i = \mathbf{E}|Z_j^{(2)} - Z_j^{(1)}|^i, \ t_j^i = \mathbf{E}|Z_j^{(2)}|^i, \ \delta_j^i = \mathbf{E}|Z_j^{(3)}|^i, 
\omega_j = \mathbf{E}|(Z_j^{(1)})^{\alpha}(Z_j^{(2)} - Z_j^{(1)})^p|, \ w_j = \mathbf{E}|(Z_j^{(2)} - Z_j^{(1)})^{\alpha}(Z_j^{(3)} - Z_j^{(2)})^p|.$$

Estimating  $|I_1|$  with the help of (37), we obtain

$$|I_1| \le C(p,\alpha)H_i^{(p)}(nI_1' + I_1''),$$
 (39)

where

$$\begin{split} I_{1}' &= \mathbf{E}|A_{J}Z_{J}^{(1)}\nu_{J}^{\alpha}| + \mathbf{E}|A_{J}Z_{J}^{(1)}\nu_{J}^{p+\alpha}| + \mathbf{E}|A_{J}Z_{J}^{(1)}\omega_{J}| \\ &+ \mathbf{E}|A_{J}Z_{J}^{(1)}\gamma_{J}^{\alpha}| + \mathbf{E}|A_{J}Z_{J}^{(1)}\gamma_{J}^{p+\alpha}| + \mathbf{E}|A_{J}Z_{J}^{(1)}w_{J}| \\ &+ \mathbf{E}|A_{J}(Z_{J}^{(1)})^{1+\alpha}| + \mathbf{E}|A_{J}(Z_{J}^{(1)})^{1+p+\alpha}| \\ &+ \mathbf{E}|A_{J}(Z_{J}^{(1)})^{1+\alpha}(Z_{J}^{(2)} - Z_{J}^{(1)})^{p}| + \mathbf{E}|A_{J}Z_{J}^{(1)}(Z_{J}^{(2)} - Z_{J}^{(1)})^{\alpha}| \\ &+ \mathbf{E}|A_{J}Z_{J}^{(1)}(Z_{J}^{(2)} - Z_{J}^{(1)})^{p+\alpha}| \\ &+ \mathbf{E}|A_{J}Z_{J}^{(1)}(Z_{J}^{(2)} - Z_{J}^{(1)})^{\alpha}(Z_{J}^{(3)} - Z_{J}^{(2)})^{p}| \end{split}$$

and

$$I_{1}'' = \sum_{j=1}^{n} \mathbf{E}|A_{j}Z_{j}^{(1)}|[\mathbf{E}|(Z_{j}^{(1)})^{\alpha}(z_{j}^{(2)})^{p}| + \mathbf{E}|(Z_{j}^{(2)} - Z_{j}^{(1)})^{\alpha}(z_{j}^{(3)})^{p}|]$$

$$+ \sum_{j=1}^{n} [\mathbf{E}|A_{j}(Z_{j}^{(1)})^{1+\alpha}(z_{j}^{(2)})^{p}| + \mathbf{E}|A_{j}Z_{j}^{(1)}(Z_{j}^{(2)} - Z_{j}^{(1)})^{\alpha}(z_{j}^{(3)})^{p}|].$$

Applying (25), (26), (28) and Hölder's inequalities, we find that  $\mathbf{E}|A_JZ_J^{(1)}\nu_J^{\alpha}|$ ,  $\mathbf{E}|A_JZ_J^{(1)}\gamma_J^{\alpha}|$ ,  $\mathbf{E}|A_JZ_J^{(1)}(Z_J^{(2)}-Z_J^{(1)})^{\alpha}|$  are bounded above by  $(2m+1)^{1+\alpha}n^{-1}L_{2+\alpha}$  and the remaining terms of  $I_1'$  are bounded above by  $(2m+1)^{1+p+\alpha}n^{-1}L_{2+p+\alpha}$ . Therefore

$$I_1' \le C(p,\alpha)n^{-1}[(m+1)^{1+\alpha}L_{2+\alpha} + (m+1)^{1+p+\alpha}L_{2+p+\alpha}). \tag{40}$$

We point out that it would have been possible to manage without the weak dependency of the r.v.'s (12) because  $I'_1$  involves moments of close r.v.'s.

The quantities  $I_1''$  and  $|I_2|$  will be estimated separately for each weak dependency.

Let the sequence (12) be m-dependent. Since

$$\mathbf{E}|z_i^{(i)}|^p \le (1 \vee 2^{p-1})(\mathbf{E}|Z_n|^p + \mathbf{E}|Z_i^{(i)}|^p), \quad i = 1, 2,$$

the m-dependence implies that

$$I_{1}^{"} \leq (1 \vee 2^{p-1}) n ([\mathbf{E}|A_{J}Z_{J}^{(1)}\nu_{J}^{\alpha}| + \mathbf{E}|A_{J}Z_{J}^{(1)}\gamma_{J}^{\alpha}| + \mathbf{E}|A_{J}(Z_{J}^{(1)})^{1+\alpha}|$$

$$+ \mathbf{E}|A_{J}Z_{J}^{(1)}(Z_{J}^{(2)} - Z_{J}^{(1)})^{\alpha}|]\mathbf{E}|Z_{n}|^{p} + \mathbf{E}|A_{J}Z_{J}^{(1)}\nu_{J}^{\alpha}t_{J}^{p}|$$

$$+ \mathbf{E}|A_{J}Z_{J}^{(1)}\gamma_{J}^{\alpha}\delta_{J}^{p}| + \mathbf{E}|A_{J}(Z_{J}^{(1)})^{1+\alpha}t_{J}^{p}|$$

$$+ \mathbf{E}|A_{J}Z_{J}^{(1)}(Z_{J}^{(2)} - Z_{J}^{(1)})^{\alpha}\delta_{J}^{p}|).$$

$$(41)$$

The expression multiplying  $\mathbf{E}|Z_n|^p$  has already been estimated. The estimates (25), (26), (28) and Hölder's inequality can be used to show that each of the remaining terms in (41) is bounded above by at least  $(6m + 1)^{1+p+\alpha}n^{-1}L_{2+p+\alpha}$ . Consequently,

$$I_1'' \le C(p,\alpha)[(m+1)^{1+\alpha}L_{2+\alpha}\mathbf{E}|Z_n|^p + (m+1)^{1+p+\alpha}L_{2+p+\alpha}). \tag{42}$$

 $I_2=0$  by virtue of the m-dependence. Therefore to complete the proof of Theorem 3, it suffices to substitute (40) and (42) in (39).

When p = 0,

$$I_1'' = n[\mathbf{E}|A_J Z_J^{(1)} 
u_J^{lpha}| + \mathbf{E}|A_J Z_J^{(1)} \gamma_J^{lpha}| + \mathbf{E}|A_J (Z_J^{(1)})^{1+lpha}| + \mathbf{E}|A_J Z_J^{(1)} (Z_J^{(2)} - Z_J^{(1)})^{lpha}|]$$

occurs in the expression for  $nI'_1$ . Thus from (39) and (40), we find that

$$|I_1| \le C(\alpha)H_i^{(0)}m^{1+\alpha}L_{2+\alpha}.$$
 (43)

If  $H_1^{(0)}<\infty$  or  $H_2^{(0)}<\infty$ , it follows from (38) and (43) that any weakly dependent sequence (12) satisfies

$$|\mathbf{E}h(Z_n) - \mathbf{E}h(\mathcal{N})| \le C(\alpha)H_i^{(0)}m^{1+\alpha}L_{2+\alpha} + |I_2|. \tag{44}$$

Suppose that (12) is  $\psi$ -mixing. Then estimating with (9) and using (36), we obtain

$$|I_2| \le 6||h_0||_{\infty} n[2\mathbf{E}^{1/2}(A_J Z_J^{(1)})^2 + (2^{3/2} + c_1)\mathbf{E}^{1/2} A_J^2 + 2^{3/2}\mathbf{E}^{1/2}(A_J^2 \nu_J^2)]\psi^{1/2}(m+1).$$
(45)

By virtue of (25), (28) and Hölder's inequality,

$$|I_2| \le C||h||_{\infty} n^{1/2} [L_2^{1/2} + mL_4^{1/2}] \psi^{1/2}(m+1).$$
 (46)

The estimates (44) and (46) yield the following result.

**Proposition 2.** Let sequence (12) be  $\psi$ -mixing and let  $\mathbf{E}X_j^4 < \infty$ ,  $j = 1, 2, \ldots, n$ . Then

$$d_i^{(0)} \le C(\alpha) \{ m^{1+\alpha} L_{2+\alpha} + n^{1/2} [L_2^{1/2} + m L_4^{1/2}] \psi^{1/2}(m+1) \}$$

providing 6m + 1 < n.

Under a.r., the quantity  $|I_2|$  can be estimated with the help of (10) and in similar fashion we deduce the following

**Proposition 3.** Let sequence (12) be absolutely regular and let  $\mathbf{E}X_j^4 < \infty$ ,  $j = 1, 2, \ldots, n$ . Then

$$d_i^{(0)} \le C(\alpha) \{ m^{1+\alpha} L_{2+\alpha} + [n^{3/4} L_4^{1/4} + mn^{1/2} L_4^{1/2}] \beta^{1/4} (m+1) \}.$$

Theorems 4 and 5 are consequences of the respective Propositions 2 and 3. Truncation. As in the estimation of  $||\Delta_n(x)||_1$ , the r.v.'s  $A_j$ , j = 1, 2, ..., n, are truncated at level t > 0 (§2, 2.1). Recall the subsidiary variables

$$\overline{A}_j^{(0)} = \overline{A}_j - \mathbf{E}\overline{A}_j \quad \text{and} \quad \overline{Z}_n^{(0)} = \sum_{j=1}^n \overline{A}_j^{(0)}.$$

Let  $h \in \mathcal{H}_1^{(0)}$ . Then any real r.v.'s  $\xi$  and  $\eta$  satisfy

$$|\mathbf{E}h(\xi) - \mathbf{E}h(\eta)| \le 3H_1^{(0)}\mathbf{E}|\xi - \eta|^{\alpha}.$$
(47)

It is easy to see that  $\mathbf{E}|Z_n - \overline{Z}_n^{(0)}|^{\alpha} \leq 2^{\alpha} \overline{\overline{L}}_1^{\alpha}$  and, when  $\mathbf{E}(\overline{Z}_n^{(0)})^2 > 0$ , that

$$\mathbf{E}|\overline{Z}_{n}^{(0)}(1-\mathbf{E}^{-1/2}(\overline{Z}_{n}^{(0)})^{2})|^{\alpha} \leq |1-\mathbf{E}(\overline{Z}_{n}^{(0)})^{2}|^{\alpha}.$$

Therefore estimating according to (47), we conclude that

$$|\mathbf{E}h(Z_n) - \mathbf{E}h(\mathcal{N})| \le 3H_1^{(0)} \left[2^{\alpha} \overline{\overline{L}}_1^{\alpha} + |1 - \mathbf{E}(\overline{Z}_n^{(0)})^2|^{\alpha}\right] + |\mathbf{E}h(\overline{Z}_n^{(0)}/\mathbf{E}^{1/2}(\overline{Z}_n^{(0)})^2) - \mathbf{E}h(\mathcal{N})|$$

$$(48)$$

when  $h \in \mathcal{H}_1^{(0)}$  and  $\mathbf{E}(\overline{Z}_n^{(0)})^2 > 0$ .

The truncation inequality (48) clearly holds for any dependency of the sequence (12). Therefore, the proofs of Theorems 6–8 are completed on merely applying it to Theorems 3–5.

#### 2.4. Estimation of $\Delta_n$ .

**Theorem 9.** Let the sequence (12) be m-dependent and let  $\mathbf{E}|X_j|^s < \infty$ ,  $2 < s \le 3$  and j = 1, 2, ..., n. Then for  $m \ge 0$  and  $n \ge 1$ 

$$\Delta_n \le C(m+1)^{s-1} L_s^*.$$

**Theorem 10.** Let (12) be a strongly mixing sequence and let  $\mathbf{E}|X_j|^s < \infty$ ,  $2 < s \leq 3$  and j = 1, 2, ..., n.

1. If  $\alpha(\tau) \leq Ke^{-\lambda \tau}$ , then

$$\Delta_n \le C(K, \lambda, s) L_s^* \ln^{s-1} (n+1).$$

2. If  $\alpha(\tau) \leq K\tau^{-\mu}$  with  $\mu \geq \frac{2(s-1)}{s(s-2)^2}[\beta(5s-6)-(s-2)(3-s)]$  and  $\beta > 1$ , then

$$\Delta_n \le C(K, \beta, s) n^c d_s,$$

where 
$$c = (4\beta + s^2 - 4)/(2(2\beta + s - 4))$$
.

For the exponential s.m. coefficient, the power of the logarithm may be lowered but under more stringent conditions on the moments of the terms. We now give a nonuniform bound in the case of strict stationarity.

**Theorem 11** (Tikhomirov (1980)). Suppose that (12) is a strictly stationary and strongly mixing sequence with  $\alpha(\tau) \leq Ke^{-\lambda \tau}$  and that  $\mathbf{E}|X_1|^{4+\nu} < \infty$  for some positive  $\nu$ . Then

$$\Delta_n \le C(K, \lambda, \nu) n^{-1/2} \ln n$$

and

$$|\Delta_n(x)| \le C(K, \lambda, \nu) \frac{\ln^3 n}{\sqrt{n}(1+|x|)^4}.$$

Heinrich (1985d) obtained a nonuniform bound for  $|\Delta_n(x)|$  for *m*-dependent r.v.'s. In particular, he proved the following result.

**Theorem 12** (1985d). Suppose that the sequence (12) is m-dependent and that  $\mathbf{E}|X_j|^s < \infty$ ,  $2 < s \le 3$  and j = 1, 2, ..., n. Then

$$|\Delta_n(x)| \le C \frac{(m+1)^{s-1} L_s^*}{(1+|x|)^s}$$

for all real x and  $n \ge 1$ .

When s = 3, the rate of convergence in Theorem 12 cannot be improved in relation to n, x (see Petrov (1987)) and m (see Berk (1973)).

We shall limit ourselves to proving Theorems 9 and 10.

Shergin (1979), (1990) was able to replace  $L_s^*$  by Lyapunov's quotient  $L_s$  in Theorems 9 and 12.

Tikhomirov (1986) obtained an estimate for the rate at which the d.f. of the  $\max_{1 \le k \le n} (X_1 + \ldots + X_k)$  for a strictly stationary and strongly mixing sequence converges to the d.f.

$$\sqrt{2/\pi} \int_0^{x^+} e^{-u^2/2} du$$

in the uniform metric, where  $x^+ = \max(0, x)$ .

Proofs of Theorems 9 and 10.

Tikhomirov's method (presented in detail in Tikhomirov (1980) for a strictly stationary sequence) consists in deriving a linear differential equation for the c.f.  $f_n(t) = \mathbf{E}e^{itZ_n}$  of a centered and normalized sum  $Z_n$ . This differential equation is "close" to the homogeneous differential equation f'(t) = -tf(t) whose solution is the c.f. of the standard normal law, that is,  $f(t) = e^{-t^2/2}$ .

It will be recalled that Theorems 9 and 10 do not assume the stationarity of the r.v.'s (12). Similar bounds for  $\Delta_n$  were found by Sunklodas (1984) for a sequence of either m-dependent or strongly mixing r.v.'s but under the additional condition  $B_n^2 \geq c_0 n$ ,  $0 < c_0 < \infty$ .

In addition to the notation of §2, 2.1, write

$$\xi_j^{(l)} = e^{it(z_j^{(l-1)} - z_j^{(l)})} - 1, \quad \eta_j^{(r)} = e^{-itZ_j^{(r)}} - 1,$$

$$a_j^{(r-1)} = \mathbf{E}\left(iA_j \prod_{l=1}^{r-1} \xi_j^{(l)}\right), \quad z_j^{(0)} = Z_n.$$

Since  $f'_n(t) = \sum_{j=1}^n \mathbf{E}(iA_je^{itZ_n})$ , by successively adding and subtracting the quantities  $\mathbf{E}(iA_je^{itz_j^{(1)}})$ ,  $\mathbf{E}(iA_j\xi_j^{(1)}e^{itz_j^{(2)}})$ , ...,  $\mathbf{E}(iA_j\prod_{l=1}^{k-1}\xi_j^{(l)}e^{itz_j^{(k)}})$  in the summation, one finds that (Tikhomirov (1980))

$$f'_{n}(t) = \sum_{j=1}^{n} \mathbf{E}(iA_{j}e^{itz_{j}^{(1)}}) + \sum_{j=1}^{n} \sum_{r=2}^{k} \mathbf{E}\left(iA_{j} \prod_{l=1}^{r-1} \xi_{j}^{(l)} e^{itz_{j}^{(r)}}\right) + \sum_{j=1}^{n} \mathbf{E}\left(iA_{j} \prod_{l=1}^{k} \xi_{j}^{(l)} e^{itz_{j}^{(k)}}\right).$$

The relation

$$\mathbf{E}e^{itz_{j}^{(r)}} = \mathbf{E}(\eta_{i}^{(r)} + 1)f_{n}(t) + \mathbf{E}[(\eta_{i}^{(r)} - \mathbf{E}\eta_{i}^{(r)})e^{itZ_{n}}]$$

can be utilized to prove that the derivative of  $f_n(t)$  with respect to t is (Sunklodas (1984))

$$f_n'(t) = (E_1 + E_2)f_n(t) + E_3 + E_4 + E_5 + E_6, (49)$$

in which

$$E_{1} = \sum_{j=1}^{n} a_{j}^{(1)}, \quad E_{2} = \sum_{j=1}^{n} \left[ a_{j}^{(1)} \mathbf{E} \eta_{j}^{(2)} + \sum_{r=3}^{k} a_{j}^{(r-1)} \mathbf{E} (\eta_{j}^{(r)} + 1) \right],$$

$$E_{3} = \sum_{r=2}^{k} \sum_{j=1}^{n} a_{j}^{(r-1)} \mathbf{E} \left[ (\eta_{j}^{(r)} - \mathbf{E} \eta_{j}^{(r)}) e^{itZ_{n}} \right],$$

$$E_{4} = \sum_{j=1}^{n} \mathbf{E} \left( iA_{j} \prod_{l=1}^{k} \xi_{j}^{(l)} e^{itz_{j}^{(k)}} \right), \quad E_{6} = \sum_{j=1}^{n} \mathbf{E} (iA_{j} e^{itz_{j}^{(1)}}),$$

$$E_{5} = \sum_{r=2}^{k} \sum_{j=1}^{n} \mathbf{E} \left[ iA_{j} \prod_{l=1}^{r-1} \xi_{j}^{(l)} (e^{itz_{j}^{(r)}} - \mathbf{E} e^{itz_{j}^{(r)}}) \right].$$

It is assumed here that 2km + 1 < n.

The linear differential equation (49) is the starting point for proving Theorems 9–11. The passage from (49) to the difference of the c.f.'s is accomplished by means of the following lemma.

Lemma 6 (Sunklodas (1984)). Let the linear differential equation

$$f'(t) = (-t + \theta a(t))f(t) + \theta b(t), \quad f(0) = 1, \tag{50}$$

be given for  $|t| \leq T_1$ , where

$$a(t) = a^{(0)} + a^{(1)}|t| + a^{(2)}t^2 + a^{(3)}|t|^{s-1}, \quad 2 < s \le 3,$$
  
 $b(t) = b^{(0)} + b^{(2)}t^2.$ 

Here the coefficients  $a^{(i)} \geq 0$  (i = 0,1,2,3) and  $b^{(j)} \geq 0$  (j = 0,2) are independent of t; and  $\theta$  is a complex function such that  $|\theta| \leq 1$ .

If  $a^{(1)} \leq 1/6$ , then

$$|f(t) - e^{-t^{2}/2}| \le C[a^{(0)}|t| + a^{(1)}t^{2} + a^{(2)}|t|^{3} + a^{(3)}|t|^{s}]e^{-t^{2}/4} + C[b^{(0)}\min(|t|^{-1}, |t|) + b^{(2)}|t|]$$
(51)

for  $|t| \leq \min(T_1, T_2)$ , where

$$T_2 = \min \left\{ \frac{1}{a^{(0)}}, \frac{1}{6a^{(2)}}, \left(\frac{1}{6a^{(3)}}\right)^{1/(s-2)} \right\}.$$

Lemma 6 and Essen's inequality (see, for example, Petrov (1987), p. 154) lead easily to the next assertion:

If the c.f.  $f(t) = \mathbf{E}e^{it\xi}$  of a real r.v.  $\xi$  satisfies the linear differential equation (50) for  $|t| \leq T_1$ , then

$$\sup_{x} |\mathbf{P}\{\xi < x\} - \Phi(x)| \le C(a^{(0)} + a^{(1)} + a^{(2)} + a^{(3)} + b^{(0)} + b^{(2)}T_1 + T_1^{-1}). \tag{52}$$

Therefore to estimate  $\Delta_n$ , all attention will be directed to deriving linear differential equation (50) for  $f_n(t)$  so that its coefficients and  $T_1$  assure a bound for  $\Delta_n$  as close as possible to being optimum.

Let us proceed to prove Theorem 9 and 10.

Assume that  $\mathbf{E}|X_j|^s < \infty$ ,  $2 < s \le 3$  and j = 1, 2, ..., n. For any dependency,

$$E_1 = -t \sum_{j=1}^{n} \mathbf{E}(A_j Z_j^{(1)}) + \theta(2^{3-s}/(s-1))(2m+1)^{s-1} L_s |t|^{s-1}.$$
 (53)

Now let the sequence (12) be m-dependent. Then

$$E_1 = -t + \theta(2^{3-s}/(s-1))(2m+1)^{s-1}L_s|t|^{s-1}.$$
 (54)

It is easily seen that

$$\mathbf{E}^{1/2}|\xi_i^{(l)}|^2 \le \sqrt{2}(m+1)d_s^{1/s}|t| \tag{55}$$

and that

$$\mathbf{E}^{1/2}|\eta_i^{(r)}|^2 \le \sqrt{2r}(m+1)d_s^{1/s}|t|. \tag{56}$$

By virtue of the m-dependence and (55),

$$|a_{j}^{(r-1)}| \leq \mathbf{E}^{1/2} A_{j}^{2} \prod_{l=1}^{r-1} \mathbf{E}^{1/2} |\xi_{j}^{(l)}|^{2}$$

$$\leq d_{s}^{1/s} (\sqrt{2}(m+1) d_{s}^{1/s} |t|)^{r-1} = a^{(r-1)}. \tag{57}$$

Since

$$|E_2| \le \sum_{j=1}^n \left[ |a_j^{(1)}| \mathbf{E} |\eta_j^{(2)}| + \sum_{r=3}^k |a_j^{(r-1)}| \right),$$
 (58)

(56) and (57) can be used to show that

$$|E_2| \le C(m+1)^2 n d_s^{3/s} t^2 \tag{59}$$

for  $|t| \le (\sqrt{2}e^2(m+1)d_s^{1/s})^{-1} = T_1$ .

By the *m*-dependence and the estimates (56) and (57), we find for  $|t| \leq T_1$  that

$$|E_{3}| \leq \sum_{r=2}^{k} a^{(r-1)} \left( \sum_{j=1}^{n} \sum_{|p-j| \leq 3rm} \mathbf{E}^{1/2} |\eta_{j}^{(r)}|^{2} \mathbf{E}^{1/2} |\eta_{p}^{(r)}|^{2} \right)^{1/2}$$

$$\leq C(m+1)^{5/2} n^{1/2} d_{s}^{3/s} t^{2}. \tag{60}$$

By virtue of (57),

$$|E_4| < na^{(k)} < C(m+1)^2 n^{1/2} d_s^{3/s} t^2$$
(61)

for  $k \geq 2 + \frac{1}{4} \ln n$  and  $|t| \leq T_1$ . Since  $E_5 = E_6 = 0$  because of the m-dependence, on substituting (54), (59)–(61) in (49), we find that

$$f_n'(t) = (-t + \theta a_n(t))f_n(t) + \theta b_n(t) \tag{62}$$

providing 2km + 1 < n,  $k \ge 2 + \frac{1}{4} \ln n$  and  $|t| < T_1$ , where

$$a_n(t) = a_n^{(2)} t^2 + a_n^{(3)} |t|^{s-1}, b_n(t) = b_n^{(2)} t^2, a_n^{(2)} = C(m+1)^2 n d_s^{3/s}, a_n^{(3)} = C(m+1)^{s-1} L_s^*, b_n^{(2)} = C(m+1)^{5/2} n^{1/2} d_s^{3/s}.$$

Relations (52) and (62) yield Theorem 9 if  $m+1 < n/(C \ln(n+1))$ , C > 1. Theorem 9 follows for all  $m \ge 0$  from the third inequality in Lemma 7 of §2, 2.5 and Esseen's inequality.

Now suppose that (12) is a s.m. sequence. Write  $z_j^{(i)} = \hat{z}_j^{(i)} + \hat{z}_j^{(i)}$ , where  $\hat{z}_j^{(i)}$  is the sum of those  $A_p$  in  $z_j^{(i)}$  for which p < j - im and  $\hat{z}_j^{(i)}$  is the sum of  $A_p$  for which p > j + im.

Then according to (6),

$$\sum_{j=1}^{n} |\mathbf{E}(A_{j}z_{j}^{(1)})| \leq \sum_{j=1}^{n} [|\mathbf{E}(A_{j}\hat{z}_{j}^{(1)})| + |\mathbf{E}(A_{j}\hat{z}_{j}^{(1)})|]$$
$$\leq 6n^{2}d_{s}^{2/s}(\alpha(m+1))^{(s-2)/s}.$$

Therefore (53) yields

$$E_1 = -t + \theta(2^{3-s}/(s-1))(2m+1)^{s-1}L_s^*|t|^{s-1} + \theta 6n^2d_s^{2/s}(\alpha(m+1))^{(s-2)/s}|t|.$$
(63)

Put 
$$\hat{\xi}_{j}^{(l)} = e^{itx} - 1$$
,  $\hat{\xi}_{j}^{(l)} = e^{ity} - 1$ , where  $x = \sum_{p=j-lm}^{j-(l-1)m-1} A_p$ ; and

 $y = \sum_{p=j+(l-1)m+1}^{j+lm} A_p$ . Then by Minkowski's inequality, for  $1 \le \mu \le s$ 

$$\max\{\mathbf{E}^{1/\mu}|\hat{\xi}_{i}^{(l)}|^{\mu},\ \mathbf{E}^{1/\mu}|\hat{\hat{\xi}}_{i}^{(l)}|^{\mu}\} \le (m+1)d_{s}^{1/s}|t| \tag{64}$$

and

$$\mathbf{E}^{1/\mu} |\eta_i^{(r)}|^{\mu} \le (2rm+1)d_s^{1/s}|t| = \eta^{(r)}. \tag{65}$$

Similarly,

$$a_j^{(1)} \le (2m+1)d_s^{2/s}|t|. (66)$$

We next estimate  $a_j^{(r-1)}$  for  $r=3,4,\ldots,k$ . Since  $|\xi_j^{(l)}| \leq |\hat{\xi}_j^{(l)}| + |\hat{\hat{\xi}}_j^{(l)}|$ , we have

$$|a_j^{(r-1)}| \le \sum_{n=0}^{r-1} \sum^* \mathbf{E} \left| A_j \prod_{\nu=1}^p \hat{\xi}_j^{(l_{\nu})} \prod_{\mu=n+1}^{r-1} \hat{\hat{\xi}}_j^{(l_{\mu})} \right|, \tag{67}$$

in which  $\sum^*$  denotes summation over all collections of indices  $1 \le l_1 < l_2 < \ldots < l_p \le r-1$  and  $1 \le l_{p+1} < l_{p+2} < \ldots < l_{r-1} \le r-1$  such that  $l_{\nu} \ne l_{\mu}$  for  $\nu \ne \mu$  (see Tikhomirov (1980)).

By Hölder's inequality,

$$\mathbf{E} \left| A_{j} \prod_{\nu=1}^{p} \hat{\xi}_{j}^{(l_{\nu})} \prod_{\mu=p+1}^{r-1} \hat{\xi}_{j}^{(l_{\mu})} \right| \\
\leq \mathbf{E}^{1/s} \left| \prod_{\nu}' \hat{\xi}_{j}^{(l_{\nu})} \prod_{\mu}' \hat{\xi}_{j}^{(l_{\mu})} \right|^{s} \mathbf{E}^{(s-1)/s} \left| A_{j} \prod_{\nu}'' \hat{\xi}_{j}^{(l_{\nu})} \prod_{\mu}'' \hat{\xi}_{j}^{(l_{\mu})} \right|^{s/(s-1)},$$

where  $\prod''$  and  $\prod'$  are respectively products over all even and all odd l from 1 up to r-1.

Let r-1 be even (if r-1 is odd, one proceeds in the same way). Put  $(\cdot) = (m+1)d_s^{1/s}|t|$ . Then by (7), (8) and (64),

$$\mathbf{E} \left| \prod_{\nu}' \hat{\xi}_{j}^{(l_{\nu})} \prod_{\mu}' \hat{\xi}_{j}^{(l_{\mu})} \right|^{s} \leq (\cdot)^{s(r-1)/2} + 8(r-3)2^{s(r-1)/2} \alpha(m+1)$$

and

$$\mathbf{E} \left| A_j \prod_{\nu} {}'' \hat{\xi}_j^{(l_{\nu})} \prod_{\mu} {}'' \hat{\xi}_j^{(l_{\mu})} \right|^{s/(s-1)} \le d_s^{1/(s-1)} \left[ (\cdot)^{s(r-1)/(2(s-1))} + 8(r-1)2^{s(r-1)/(2(s-1))} (\alpha(m+1))^{(s-2)/(s-1)} \right].$$

Therefore

$$\mathbf{E} \left| A_j \prod_{\nu=1}^p \hat{\xi}_j^{(l_{\nu})} \prod_{\mu=p+1}^{r-1} \hat{\xi}_j^{(l_{\mu})} \right| \le 8d_s^{1/s} (\alpha(m+1))^{(s-1)/s} (r-1) 2^{r-1} + d_s^{1/s} (\cdot)^{r-1} + 7d_s^{1/s} (2(\cdot))^{(r-1)/2} (\alpha(m+1))^{(s-2)/s} (r-1).$$

Since there are at most  $2^{r-1}$  terms in (67), for  $|t| \leq (32(m+1)d_s^{1/s})^{-1} = T_3$  and r = 3, 4, ..., k, we obtain

$$|a_j^{(r-1)}| \le C \left\{ m^2 d_s^{3/s} t^2 \left( \frac{1}{2} \right)^{4r} + d_s^{1/s} (\alpha(m+1))^{(s-2)/s} \left[ r \left( \frac{1}{2} \right)^r + r 4^r (\alpha(m+1))^{1/s} \right] \right\} = a^{(r-1)}. (68)$$

It is now possible to estimate  $|E_4|$ . For  $k \ge \ln n/(8 \ln 2) \ge 3$  and  $|t| \le T_3$ , we find that

$$|E_4| \le na^{(k)} \le C[m^2 n^{1/2} d_s^{3/s} t^2 + d_s^{1/s} n(\alpha(m+1))^{(s-2)/s}). \tag{69}$$

We estimate the terms  $E_2$ ,  $E_3$  and  $E_5$  subject to the additional condition

$$k^{3/2}4^k(\alpha(m+1))^{1/s} \le 1. \tag{70}$$

Substituting (66), (65) and (68) in (58), we find for  $|t| \leq T_3$  that

$$|E_2| \le C[m^2 n d_s^{3/s} t^2 + n d_s^{1/s} (\alpha(m+1))^{(s-2)/s}] \tag{71}$$

under condition (70).

Consider  $cov(\hat{\xi}, \eta) = \mathbf{E}(\xi - \mathbf{E}\xi)(\overline{\eta - \mathbf{E}\eta})$ . By virtue of (6),

$$|E_{3}| \leq \sum_{r=2}^{k} a^{(r-1)} \left( \sum_{j=1}^{n} \sum_{p=1}^{n} |\operatorname{cov}(\eta_{j}^{(r)}, \eta_{p}^{(r)})| \right)^{1/2}$$

$$\leq \sum_{r=2}^{k} a^{(r-1)} \left( \sum_{j=1}^{n} \sum_{|p-j| \leq 2rm} \mathbf{E}^{1/2} |\eta_{j}^{(r)}|^{2} \mathbf{E}^{1/2} |\eta_{p}^{(r)}|^{2} + \sum_{j=1}^{n} \sum_{|p-j| > 2rm} 24 \mathbf{E}^{1/s} |\eta_{j}^{(r)}|^{s} \mathbf{E}^{1/s} |\eta_{p}^{(r)}|^{s} (\alpha(|p-j| - 2rm))^{(s-2)/s} \right)^{1/2}.$$

Noting (66), (65) and (68), we obtain

$$|E_3| \le C[m^{1/2} + A_{n,s}^{1/2}][m^2n^{1/2}d_s^{3/s}t^2 + n^{1/2}d_s^{1/s}(\alpha(m+1))^{(s-2)/s}]$$
 (72)

for  $|t| \leq T_3$  under assumption (70).

The definition of  $z_j^{(r)}$  shows that  $\hat{z}_j^{(r)}$  and  $\hat{z}_j^{(r)}$  cannot vanish simultaneously. We shall consider that both do not vanish (otherwise the computations merely become simpler). Observe that

$$\begin{split} &\left|\mathbf{E}\left[iA_{j}\prod_{l=1}^{r-1}\xi_{j}^{(l)}(e^{tiz_{j}^{(r)}}-\mathbf{E}e^{itz_{j}^{(r)}})\right]\right| \\ &\leq \left|\mathbf{E}\left(iA_{j}\prod_{l=1}^{r-1}\xi_{j}^{(l)}e^{itz_{j}^{(r)}}\right)-\mathbf{E}\left(iA_{j}\prod_{l=1}^{r-1}\xi_{j}^{(l)}e^{it\hat{z}_{j}^{(r)}}\right)\mathbf{E}e^{it\hat{z}_{j}^{(r)}}\right| \\ &+\left|\mathbf{E}\left[iA_{j}\prod_{l=1}^{r-1}\xi_{j}^{(l)}(e^{it\hat{z}_{j}^{(r)}}-\mathbf{E}e^{it\hat{z}_{j}^{(r)}})\right]\right| \\ &+\left|\mathbf{E}e^{it\hat{z}_{j}^{(r)}}\mathbf{E}e^{it\hat{z}_{j}^{(r)}}-\mathbf{E}e^{itz_{j}^{(r)}}\right|\cdot\left|\mathbf{E}\left(iA_{j}\prod_{l=1}^{r-1}\xi_{j}^{(l)}\right)\right|. \end{split}$$

The right-hand side of this last inequality can be estimated by means of inequalities (5) and (4) and one can show that it does not exceed  $48 \cdot 2^{r-1} d_s^{1/s} (\alpha(m+1))^{(s-1)/s}$ . Summing the resultant inequalities over all j and r, we conclude that

$$|E_5| \le 48nd_s^{1/s}(\alpha(m+1))^{(s-2)/s}$$
 (73)

under condition (70).

 $|E_6|$  may be estimated in similar fashion with the result

$$|E_6| \le 32nd_s^{1/s}(\alpha(m+1))^{(s-1)/s}.$$
 (74)

Put  $A_s = \sum_{\tau=1}^{\infty} (\alpha(\tau))^{(s-2)/s}$ . The substitution of (63), (69), (71)–(74) in (49) yields this: If  $A_s < \infty$ ,  $k^{3/2} 4^k (\alpha(m+1))^{1/s} \le 1$ ,  $3 \le \ln n/(8 \ln 2) \le k$  and  $2k(m+1) \le n$ , then

$$f_n'(t) = (-t + \theta a_n(t)) f_n(t) + \theta b_n(t) \tag{75}$$

for  $|t| \leq T_3$ , where

$$a_n(t) = a_n^{(0)} + a_n^{(1)}|t| + a_n^{(2)}t^2 + a_n^{(3)}|t|^{s-1},$$
  

$$b_n(t) = b_n^{(0)} + b_n^{(2)}t^2,$$

$$\begin{array}{ll} a_n^{(0)} = Cnd_s^{1/s}(\alpha(m+1))^{(s-2)/s}, & a_n^{(1)} = Cn^2d_s^{2/s}(\alpha(m+1))^{(s-2)/s}, \\ a_n^{(2)} = Cm^2nd_s^{3/s}, & a_n^{(3)} = Cm^{s-1}L_s^*, \\ b_n^{(0)} = C(A_s)nd_s^{1/s}(\alpha(m+1))^{(s-2)/s}, & b_n^{(2)} = C(A_s)m^{5/2}n^{1/2}d_s^{3/s}. \end{array}$$

The differential equations (50) and (75) are of the same form and so (52) yields the following

**Proposition 4.** Suppose that (12) is a s.m. sequence,  $\mathbf{E}|X_j|^s < \infty$ ,  $2 < s \le 3$  and  $j = 1, 2, \ldots, n$ ,  $A_s < \infty$  and  $k^{3/2}4^k(\alpha(m+1))^{1/s} \le 1$ , with the positive integers k and m satisfying  $3 \le \ln n/(8\ln 2) \le k$  and  $2k(m+1) \le n$ . Then

 $\Delta_n \le C(A_s)[m^{s-1}L_s^* + n^2 d_s^{2/s}(\alpha(m+1))^{(s-2)/s}).$ 

It merely remains to select k and m depending on the rate of decay of the s.m. coefficient.

If  $n \ge C(K, \lambda, s)$ , part 1 of Theorem 10 follows from Proposition 4 with  $m = \left[\frac{3s}{\lambda(s-2)}\ln(n+1)\right]$  and  $k = \left[\frac{1}{(s-2)}\ln(n+1)\right]$ .

If  $n > C(K, \beta, s)$ , part 2 of Theorem 10 follows from Proposition 4 with  $m = [n^{\varepsilon}], \ \varepsilon = s(s-2)/(2(s-1)(2\beta+s-2))$  and  $k = \left[\frac{5\beta}{8(2\beta+1)}\ln(n+1)\right].$ 

For small n, Theorem 10 is a consequence of the estimate (17).

#### 2.5. Heinrich's Method for m-Dependent Random Variables.

Heinrich's method (1982) is a fairly general way of deducing various limit theorems for the sums  $Z_n$  of m-dependent r.v.'s. It is based on the factorization of the c.f.  $f_n(t) = \mathbf{E}e^{itZ_n}$  (or moment-generating function  $\mathbf{E}e^{zZ_n}$ ,  $z \in (C^1)$  in a neighborhood of t = 0 (or z = 0).

We shall only consider here the factorization of  $f_n(t)$ . By making use of the factorization of  $f_n(t)$  once it has been found, one can obtain, for instance, these sorts of results for sums of m-dependent r.v.'s: convergence to unbounded distributions, uniform and non-uniform bounds for the rate of convergence in the CLT, asymptotic expansions, moderate and large deviations and so on.

The gist of Heinrich's method will be demonstrated by the proof of a single lemma.

**Lemma 7.** Suppose that (12) is an m-dependent sequence and that  $\mathbf{E}|X_j|^s < \infty$ ,  $2 < s \leq 3$  and j = 1, 2, ..., n. Then the following inequalities hold in the interval

$$|t| \le \left(\frac{1-2c}{2C_1(s)} \cdot \frac{1}{(m+1)^{s-1}L_s^*}\right)^{1/(s-2)} = T_4$$

for  $m \geq 0$ :

1. 
$$\left| \ln f_n(t) + \frac{t^2}{2} \right| \le C_1(s)(m+1)^{s-1} L_s^* |t|^s$$
,

$$2. |f_n(t)| \le e^{-ct^2},$$

3. 
$$|f_n(t) - e^{-t^2/2}| < C_1(s)(m+1)^{s-1}L_s^*|t|^s e^{-ct^2}$$

where

$$C_1(s) = \left(149 + \frac{1}{s}\right) \frac{2^{4-s}}{s-1}, \quad 0 < c < 1/2.$$

When s=3, Lemma 7 follows (with different constants) from Corollary 3.2 in Heinrich (1982). Since  $|e^z| \le e^{|z|}$  and  $|e^z-1| \le |z|e^{|z|}$  for any complex z, the second and third inequalities in Lemma 7 are consequence of the first one. To prove the first inequality, we need several preliminary assertions.

For any sequence of complex r.v.'s  $\xi_1, \xi_2, \ldots$  such that  $\mathbf{E}|\xi_j|^k < \infty, j = 1, 2, \ldots, k$ , the symbol  $\hat{\mathbf{E}}\xi_1\xi_2\ldots\xi_k$  means that  $\hat{\mathbf{E}}\xi_1 = \mathbf{E}\xi_1$  and for  $k \geq 2$ , that (see Heinrich (1982))

$$\hat{\mathbf{E}}\xi_1\xi_2\dots\xi_k = \mathbf{E}\xi_1\xi_2\dots\xi_k - \sum_{j=1}^{k-1}\hat{\mathbf{E}}\xi_1\dots\xi_j\mathbf{E}\xi_{j+1}\dots\xi_k.$$
 (76)

This symbol was first introduced in another way by Statulevičius (1970b) and is known as the k-th centered moment. Among the many interesting properties of centered moments are the following.

**Lemma 8** (Heinrich (1982)). Let  $\xi_1, \xi_2, \dots, \xi_k$  be 1-dependent complex r.v.'s.

1. If  $\mathbf{E}|\xi_j|^k < \infty$ , j = 1, 2, ..., k, then

$$\hat{\mathbf{E}}(\xi_1 + a_1)(\xi_2 + a_2) \dots (\xi_k + a_k) = \hat{\mathbf{E}}\xi_1 \xi_2 \dots \xi_k, \tag{77}$$

where  $a_1, a_2, \ldots, a_k$  are any complex numbers.

2. If  $\mathbf{E}|\xi_j|^2 < \infty$ , j = 1, 2, ..., k, then

$$|\hat{\mathbf{E}}\xi_1\xi_2\dots\xi_k| \le 2^{k-1} \prod_{j=1}^k \mathbf{E}^{1/2} |\xi_j|^2.$$
 (78)

Corresponding to a sequence (12) of m-dependent r.v.'s, we form new 1-dependent r.v.'s

$$Y_j = \sum_{p=(j-1)(m+1)+1}^{j(m+1)} A_p, \quad j=1,2,\ldots,N = [n/(m+1)),$$

$$Y_{N+1} = \begin{cases} \sum_{p=N(m+1)+1}^{n} A_p & \text{if } N(m+1) < n, \\ 0 & \text{if } N(m+1) = n, \end{cases}$$

where [x] is the integer part of x.

Put  $U_j = \sum_{i=1}^j Y_i$ , j = 1, 2, ..., N+1,  $w = \max_{1 \le j \le N+1} \mathbf{E}^{1/2} |e^{itY_j} - 1|^2$  and  $u_j(t) = 2\mathbf{E}^{1/2} |e^{itY_{j-1}} - 1|^2 \mathbf{E}^{1/2} |e^{itY_j} - 1|^2$ .

Then the following is true.

**Lemma 9** (Heinrich (1982)). If  $w \leq 1/6$ , then

1.  $f_n(t) = \prod_{j=1}^{N+1} g_j(t), \tag{79}$ 

where  $g_1(t) = \mathbf{E}e^{itY_1}$ , and for j = 2, 3, ..., N + 1,

$$g_{j}(t) = \frac{\mathbf{E}e^{itU_{j}}}{\mathbf{E}e^{itU_{j-1}}} = \mathbf{E}e^{itY_{j}} + \sum_{a=1}^{j-1} \frac{\hat{\mathbf{E}}(e^{itY_{a}} - 1)(e^{itY_{a+1}} - 1)\dots(e^{itY_{j}} - 1)}{\prod_{p=a}^{j-1} g_{p}(t)};$$
(80)

2. for j = 1, 2, ..., N + 1,

$$|g_j(t) - 1| \le |\mathbf{E}e^{itY_j} - 1| + 3u_j(t)$$
 (81)  
  $\le 2w.$  (82)

Lemma 9 is proved by induction using relations (76)–(78). By virtue of (80),

$$\begin{split} &\sum_{j=1}^{N+1} [g_j(t)-1] = \sum_{j=1}^{N+1} (\mathbf{E}e^{itY_j}-1) \\ &+ \sum_{j=2}^{N+1} \sum_{a=1}^{j-1} \left(\frac{1}{g_a(t)\dots g_{j-1}(t)}-1\right) \hat{\mathbf{E}}(e^{itY_a}-1)\dots (e^{itY_j}-1) \\ &+ \sum_{j=3}^{N+1} \sum_{a=1}^{j-2} \hat{\mathbf{E}}(e^{itY_a}-1)\dots (e^{itY_j}-1) \\ &+ \sum_{j=2}^{N+1} \hat{\mathbf{E}}(e^{itY_{j-1}}-1)(e^{itY_j}-1). \end{split}$$

Since  $Y_1, Y_2, \dots, Y_{N+1}$  are 1-dependent r.v.'s with zero expectations,

$$\frac{t^2}{2} = \sum_{j=1}^{N+1} \left[ \mathbf{E}e^{itY_j} - 1 - \frac{(it)^2}{2} \mathbf{E}Y_j^2 \right] 
+ \sum_{j=2}^{N+1} [\hat{\mathbf{E}}(e^{itY_{j-1}} - 1)(e^{itY_j} - 1) - (it)^2 \mathbf{E}(Y_{j-1}Y_j)] 
- \sum_{j=1}^{N+1} (\mathbf{E}e^{itY_j} - 1) - \sum_{j=2}^{N+1} \hat{\mathbf{E}}(e^{itY_{j-1}} - 1)(e^{itY_j} - 1).$$

Adding the last two relations, we find that if (12) is a sequence of m-dependent r.v.'s and  $\mathbf{E}|X_j|^s < \infty$ ,  $2 < s \le 3$  and j = 1, 2, ..., n, then

$$\ln f_n(t) + \frac{t^2}{2} = \Sigma_1 + \ldots + \Sigma_5,$$
 (83)

where

$$\Sigma_{1} = \sum_{j=1}^{N+1} [\ln g_{j}(t) - (g_{j}(t) - 1)],$$

$$\Sigma_{2} = \sum_{j=2}^{N+1} \sum_{a=1}^{j-1} \left( \frac{1}{g_{a}(t) \dots g_{j-1}(t)} - 1 \right) \hat{\mathbf{E}}(e^{itY_{a}} - 1) \dots (e^{itY_{j}} - 1),$$

$$\Sigma_{3} = \sum_{j=3}^{N+1} \sum_{a=1}^{j-2} \hat{\mathbf{E}}(e^{itY_{a}} - 1) \dots (e^{itY_{j}} - 1),$$

$$\Sigma_{4} = \sum_{j=1}^{N+1} \left[ \mathbf{E}e^{itY_{j}} - 1 - \frac{(it)^{2}}{2} \mathbf{E}Y_{j}^{2} \right],$$

$$\Sigma_{5} = \sum_{j=2}^{N+1} \left[ \hat{\mathbf{E}}(e^{itY_{j-1}} - 1)(e^{itY_{j}} - 1) - (it)^{2} \mathbf{E}(Y_{j-1}Y_{j}) \right].$$

Everywhere below when estimating the right-hand side of (83), we are assuming that  $w \le 1/6$ .

By the simple inequality  $|\ln z - (z-1)| \le |z-1|^2$ , which is true for  $|z-1| \le 1/2$ , and the estimates (81)–(82), it follows that

$$|E_1| \le 2w \sum_{j=1}^{N+1} |\mathbf{E}e^{itY_j} - 1| + 6w \sum_{j=2}^{N+1} u_j(t).$$
 (84)

According to (82),

$$\left| \frac{1}{g_a(t) \dots g_{j-1}(t)} - 1 \right| \le 3(j-a)2^{j-a-1}w,$$

and the estimate (78) leads to

$$|\Sigma_2| \le 27w \sum_{j=2}^{N+1} u_j(t).$$
 (85)

The estimate (78) also assures that

$$|\Sigma_3| \le 3w \sum_{j=3}^{N+1} u_j(t).$$
 (86)

For  $2 < s \le 3$ ,

$$\begin{aligned} |e^{ix} - 1| &\leq 2^{3-s} |x|^{s-2}, \\ |e^{ix} - 1 - ix| &\leq (2^{3-s}/(s-1))|x|^{s-1}, \\ \left|e^{ix} - 1 - ix - \frac{(ix)^2}{2}\right| &\leq (2^{3-s}/(s-1)s)|x|^s, \end{aligned}$$

and so

$$|\Sigma_4| \le (2^{3-s}/(s-1)s)|t|^s \sum_{j=1}^{N+1} \mathbf{E}|Y_j|^s,$$
 (87)

$$|\Sigma_5| \le 3(2^{3-s}/(s-1))|t|^s \sum_{j=1}^{N+1} \mathbf{E}|Y_j|^s,$$
 (88)

and

$$w\sum_{j=1}^{N+1} |\mathbf{E}e^{itY_j} - 1| \le (2^{3-s}/(s-1))(N+1) \max_{1 \le j \le N+1} \mathbf{E}|Y_j|^s |t|^s.$$
 (89)

Adding (84)-(88) and using (89) in conjunction with the fact that

$$\sum_{j=2}^{N+1} u_j(t) \le 4 \sum_{j=1}^{N+1} |\mathbf{E}e^{itY_j} - 1|,$$

we find for  $w \leq 1/6$  that

$$\left| \ln f_n(t) + \frac{t^2}{2} \right| \le C_1(s)(m+1)^{s-1} L_s^* |t|^s. \tag{90}$$

It remains to observe that  $w \leq 1/6$  when  $|t| \leq T_4$ . Consequently, the first inequality of Lemma 7 has been proved and thereby all of Lemma 7.

From Lemma 9 it is seen that the functions  $g_j(t)$  (whose product is the c.f. of  $Z_n$ ) although not c.f.'s, behave primarily like the c.f.'s of the  $Y_i$ 's. The next lemma, for example, underscores this fact.

**Lemma 10** (Heinrich (1982)). Let the sequence (12) be m-dependent and let  $\max_{1 \le j \le n} \mathbf{E}|X_j|^p < \infty$  for some  $p = 1, 2, \ldots$  Then when  $w \le 1/6$ ,

1. 
$$\max_{1 \le j \le N+1} \left| \frac{d^p}{dt^p} g_j(t) \right| \le C(p) \max_{1 \le j \le N+1} \mathbf{E} |Y_j|^p,$$
2. 
$$\max_{1 \le j \le N+1} \left| \frac{d^p}{dt^p} \ln g_j(t) \right| \le C(p) \max_{1 \le j \le N+1} \mathbf{E} |Y_j|^p,$$

where the constants C(p) can be determined explicitly.

Lemma 9 therefore plays a fundamental role in the study of the limiting law for the distribution of the normalized sum  $Z_n$  of m-dependent r.v.'s (see Heinrich (1982), (1984), (1985a,b,c,d)).

§3. Estimation of the Rate of Convergence in the Central Limit Theorem for Weakly Dependent Random Fields

Let  $Z^d=\{a=(a_1,\ldots,a_d): a_i\in\{0,\pm 1,\ldots\},\ i=1,2,\ldots,d\},\ ||a||=\max_{1\leq i\leq d}|a_i|\ \text{and}\ \mathcal{V}=\{V\subset Z^d: |V|<\infty\},\ \text{where}\ |V|=\#\{a:a\in V\}\ \text{is}$  the number of elements in V. The distance between  $V_1,V_2\in\mathcal{V}$  is defined as follows:  $d(V_1,V_2)=\min\{||a-b||:a\in V_1,b\in V_2\}.$   $\mathcal{F}_V$  denotes the  $\sigma$ -algebra of events generated by the r.v.'s  $\{X_a,a\in V\}.$ 

In what follows, we shall consider a real random field  $\{X_a, a \in Z^d\}, d \ge 1$ , satisfying one of the following weak dependence conditions:

- 1. m-dependence:  $\mathcal{F}_{V_1}$  and  $\mathcal{F}_{V_2}$  are independent for  $\forall V_1, V_2 \in \mathcal{V}$  with  $d(V_1, V_2) > m$ ;
- 2. strong mixing (s.m.): if there exist functions  $M: Z_+^2 \to [1, \infty)$  and  $\alpha: N \to [0, \infty)$  such that M is nondecreasing in each argument,  $\alpha(r) \downarrow 0$  as  $r \to \infty$  and for  $\forall V_1, V_2 \in \mathcal{V}$ ,

$$\sup_{\substack{A \in \mathcal{F}_{V_1} \\ B \in \mathcal{F}_{V_2}}} |\mathbf{P}\{AB\} - \mathbf{P}\{A\}\mathbf{P}\{B\}| \le M(|V_1|, |V_2|)\alpha(d(V_1, V_2)).$$

The definition of these and other mixing coefficients for random fields as well as references on the subject may be found, for instance, in Dobrushin (1968), Bulinskii (1987), (1989), Takahata (1983), (1984), Sunklodas (1986) and Nakhapetyan (1987). Let

$$\{X_a, \ a \in \mathbb{Z}^d\}, \quad d \ge 1, \tag{91}$$

be a real random field with  $\mathbf{E}X_a=0$  and  $\mathbf{E}X_a^2<\infty$  for  $a\in V$ . For  $V\in\mathcal{V},$   $V\neq\emptyset$ , put

$$\begin{split} S_{V} &= \sum_{a \in V} X_{a}, \quad B_{V}^{2} = \mathbf{E} S_{V}^{2}, \quad Z_{V} = S_{V}/B_{V}, \\ F_{V}(x) &= \mathbf{P}\{Z_{V} < x\}, \quad A_{a} = X_{a}/B_{V}, \quad L_{r} = \sum_{a \in V} \mathbf{E}|A_{a}|^{r}, \\ L_{s}^{*} &= |V|d_{s}, \quad d_{s} = \max_{a \in V} \mathbf{E}|A_{a}|^{s}, \quad \Delta_{V}(x) = F_{V}(x) - \Phi(x), \\ \Delta_{V} &= \sup_{x} |\Delta_{V}(x)|, \quad ||\Delta_{V}(x)||_{1} = \int_{-\infty}^{\infty} |\Delta_{V}(x)|dx, \\ d_{i}^{(p)}(F_{V}, \Phi) &= \sup_{h \in \mathcal{H}^{(p)}} |\mathbf{E}h(Z_{V}) - \mathbf{E}h(\mathcal{N})|/||h||_{BH_{i}^{(p)}}, \end{split}$$

where  $\mathcal{H}_{i}^{(p)}$  is the class of functions  $h: R \to R$  with norm  $||h||_{BH_{i}^{(p)}}$  defined in §2, 2.1, and i=1,2.

The rate of convergence in the CLT for weakly dependent random fields has been estimated by generalizing the methods developed for sequences of weakly depenent r.v.'s. The specific difficulties that have to be overcome in estimating  $\Delta_V$  for multi-indexed terms, the distinctive features of mixing fields, the limits of applicability of Bernstein's method to random fields and other related question are discussed in detail in Bulinskii's book (1989). Leonenko (1975) found a bound for  $\Delta_V$  in the case of m-dependent random fields for integer parallelepipeds  $V \subset \mathbb{Z}^d$ . When d=1, it reduces to Petrov's result (1960) cited above. By generalizing Maejima's results (1978) for r.v.'s, Rao (1981) found a nonuniform bound for  $|\Delta_V(x)|$  for integer parallelepipeds for m-dependent random fields. He conjectured particularly that it was impossible to obtain an estimate for  $\Delta_V$  of order  $O(|V|^{-\gamma})$ ,  $0 < \gamma \le 1/2$ , even for an m-dependent random field. This conjecture was disproved by Takahata (1983) and Guyon and Richardson (1984). A more precise uniform estimate (compared to those of Leonenko (1975) and Rao (1981)) for weakly dependent random additive functions (encompassing the class of m-dependent ones) was found by Bulinskii (1977). He subsequently strengthened this estimate (1987). The proofs by Rao, Leonenko and Bulinskii (1977) utilize Bernstein's method.

More exact estimates of the rate of convergence in the CLT for weakly dependent random fields have been found by means of the techniques of Stein and Tikhomirov.

Guyon and Richardson (1984) study the rate of convergence in the CLT for centered weakly dependent random fields  $\{X_a, a \in Z^d\}$  (either m-dependent or s.m. with  $M \equiv 1$ ) that satisfy  $\sup_{a \in Z^d} \mathbf{E} |X_a|^{2+\delta} < \infty, \ \delta > 0$ . The summation  $S_{V_n} = \sum_{a \in V_n} X_a$  is over a strictly increasing sequence of sets  $V_n \in \mathcal{V}$  such that  $\liminf_{n \to \infty} B_{V_n}^2/|V_n| > 0$ , where  $B_{V_n}^2 = \mathbf{E} S_{V_n}^2$ .

In particular, they show that

1. for m-dependent fields

$$\Delta_{V_n} = \begin{cases} O(B_{V_n}^{-\delta}) & \text{if } 0 < \delta < 1; \\ O[B_{V_n}^{-1}(\log B_{V_n})^{(d-1)/2}) & \text{if } \delta \ge 1; \end{cases}$$
(92)

2. for s.m. random fields with  $M\equiv 1$  and  $\alpha$  an exponentially decreasing function (here  $a\wedge b=\min(a,b)$ )

$$\Delta_{V_n} = O[B_V^{-(\delta \wedge 1)}(\log B_{V_n})^{d(1+\delta) \wedge 2}). \tag{94}$$

If  $\sup_{a\in\mathbb{Z}^d}\mathbf{E}|X_a|^{4+\delta}<\infty$ ,  $\delta>0$ , the last estimate can be improved to

$$\Delta_{V_n} = O[B_{V_n}^{-1} (\log B_{V_n})^d). \tag{95}$$

Guyon and Richardson (1984) also investigated the case where  $M \not\equiv 1$  and  $\alpha$  decreases like a power function. The proofs are carried out by Tikhomirov's method.

Takahata (1983) found upper bounds for  $\Delta_{V_n}$  and  $||\Delta_{V_n}(x)||_1$  when the random field is m-dependent or s.m. with  $M \not\equiv 1$  and  $\alpha$  exponentially decreasing; the summation  $S_{V_n} = \sum_{a \in V_n} X_a$  is over a sequence of sets  $V_n \in \mathcal{V}$ 

such that  $|V_n| \to \infty$   $(n \to \infty)$  and  $\liminf_{n \to \infty} B_{V_n}^2/|V_n| > 0$ . Under these conditions, the following was shown:

1. for m-dependent random fields

$$\Delta_{V_n} = O(|V_n|^{-1/2}) \tag{96}$$

if  $\sup_{a\in\mathbb{Z}^d}\mathbf{E}X_a^8<\infty$ , and

$$||\Delta_{V_n}(x)||_1 = O(|V_n|^{-1/2}) \tag{97}$$

if  $\sup_{a \in \mathbb{Z}^d} \mathbf{E} X_a^4 < \infty$ ;

2. for s.m. random fields with  $M(n,m) \leq B(n+m)^k$  for some k>1 and  $\alpha(\tau) \leq Ke^{-\lambda \tau}$ 

$$\Delta_{V_n} = O[|V_n|^{-1/2} (\log |V_n|)^d] \tag{98}$$

if  $\sup_{a \in \mathbb{Z}^d} \mathbf{E} |X_a|^{8+\delta} < \infty$ ,  $\delta > 0$ , and

$$||\Delta_{V_n}(x)||_1 = O[|V_n|^{-1/2}(\log|V_n|)^d]$$
(99)

if  $\sup_{a \in \mathbb{Z}^d} \mathbf{E} |X_a|^{4+\delta} < \infty, \ \delta > 0.$ 

The estimate (96) established by Takahata (1983) for an *m*-dependent random field refines Riauba's paper (1980). Takahata (1983) used Stein's technique to prove his result.

The methods of Stein (Takahata (1983), Sunklodas (1986)) and Tikhomirov (Guyon and Richardson (1984), Sunklodas (1986), Bulinski (1986a), (1987)) may be extended to weakly dependent nonstationary random fields whose terms have finite absolute moments of order  $s, 2 < s \leq 3$ . Without any assumptions about the linear growth of the variance  $B_V^2$  of the sum  $S_V, \tilde{V} \in \mathcal{V}$ , one may estimate  $||\Delta_V(x)||_1$  (by Stein's method) and  $\Delta_V$  (by Tikhomirov's method) in such a way that for d=1 these estimates yield the best known estimates (or ones close to them) for a weakly dependent sequence of r.v.'s (Tikhomirov (1980), Erickson (1974), Sunklodas (1982)).

We state a number of results of this kind.

**Theorem 13** (Sunklodas (1986)). Suppose that the random field (91) is m-dependent and that  $\mathbf{E}|X_a|^s < \infty$ ,  $2 < s \le 3$  and  $a \in V$ . Then

$$\Delta_V \le C(d)\{(m+1)^{d(s-1)}L_s^* + (m+1)^d d_s^{1/s}(\ln(|V|+1))^{(d-1)/2}\}$$

for  $m+1 \le |V|^{1/d}/(C\ln(|V|+1))$ , C > 1.

**Theorem 14** (Sunklodas (1986)). Suppose that (91) is a s.m. random field with  $M(n,m) \leq B(n+m)^p$ ,  $\alpha(\tau) \leq Ke^{-\lambda \tau}$  and that  $\mathbf{E}|X_a|^s < \infty$ ,  $2 < s \leq 3$ , with  $a \in V$ . If  $0 \leq p < \infty$ , then

$$\Delta_V \le C(B, K, \lambda, d, p, s) \{ L_s^* (\ln(|V| + 1))^{d(s-1)} + |V|^{1/2} d_s^{2/s} (\ln(|V| + 1))^{1 + (dp(s-2)/(2s))} \}.$$

**Theorem 15** (Sunklodas (1986)). Suppose that (91) is an m-dependent random field and  $\mathbb{E}|X_a|^s < \infty$ ,  $2 < s \le 3$ ,  $a \in V$ . Then for  $m \ge 0$ 

$$||\Delta_V(x)||_1 \le C(d)(m+1)^{d(s-1)}L_s.$$

The author has also found more precise estimates for  $||\Delta_V(x)||_1$  when the terms have finite second moments.

It should be noted that for weakly dependent random fields (just as for a sequence of r.v.'s),  $\Delta_V$  can be estimated in terms of  $L_s^*$  and  $d_s$  and  $||\Delta_V(x)||_1$  can be estimated in terms of Lyapunov's quotient  $L_s$ .

Close results to Theorems 13 and 14 were obtained independently by Bulinskii (1986a), (1987). His mixing conditions are more general since they take into account a "geometric" aspect of selecting the sets used to define the mixing coefficient (see also Bulinskii (1989)).

Herrndorf (1983b) constructed an example of a s.m. strictly stationary sequence of r.v.'s whose s.m. coefficient decrease arbitrarily fast and whose partial sums have a variance increasing regularly. However, the CLT is not obeyed if just the second moments of the terms exist. It is therefore reasonable to estimate the rate of convergence in the CLT for weakly dependent random fields by imposing moment restrictions such as  $\sup_{a\in Z^d}\mathbf{E}\mathcal{G}(|X_a|)<\infty$  on the terms, where  $\mathcal{G}$  satisfies the condition  $\lim_{x\to\infty}x^{-2}\mathcal{G}(x)=\infty$ . Bulinskii and Doukhan (1990) obtained an estimate for the rate of convergence in the CLT for  $\tilde{\alpha}$ -mixing random fields (see (1990)) assuming the finiteness of the moments of the terms of "small" order (for example, of the type  $\mathbf{E}X^2\ln^{4}_{h}(x)<\infty$ ). These estimates were derived by means of truncation as applied to Bulinskii's results (1986a). The author (1990) found estimates for  $d_i^{(p)}$ , i=1,2, for various types of mixing random fields. The problem is reduced to estimating an absolute moment of  $Z_V$  whose order depends on p and the type of mixing.

Here we state just one estimate which generalizes Theorem 3 to random fields.

**Theorem 16** (Sunklodas (1990)). Suppose that the function  $h: R \to R$  satisfies the condition  $H_1^{(p)} < \infty$  or  $H_2^{(p)} < \infty$  (see §2, 2.1) and that (91) is an m-dependent random field with  $\mathbf{E}|X_a|^{2+p+\alpha} < \infty$  and  $a \in V$ . If  $(6m+1)^d < |V|$ , then

$$|\mathbf{E}h(Z_n) - \mathbf{E}h(\mathcal{N})| \le C(d, p, \alpha) H_i^{(p)} \{ (m+1)^{d(1+\alpha)} L_{2+\alpha} (1 + \mathbf{E}|Z_V|^p) + (m+1)^{d(1+p+\alpha)} L_{2+p+\alpha} \}.$$

The boundedness of h is not required in Theorem 16.

Zuev (1989) found an estimate for  $\Delta_V$  for m(d)-dependent random fields (for an exact definition, see that paper).

Shergin (1988), (1990) found estimates for  $\Delta_V$ ,  $\Delta_V(x)$  and  $\int_{-\infty}^{\infty} |x|^l |\Delta_V(x)| dx$  for finitely dependent r.v.'s (see Chen (1978)).

Mukhamedov (1987) obtained estimates for  $\int_{-\infty}^{\infty} |x|^l |\Delta_V(x)| dx$  using Stein's technique for s.m. and u.s.m. random fields.

For weakly dependent stationary r.v.'s and random fields with a very slowly increasing mixing coefficient, Nakhapetyan (1987), (1989) applied his small blocking method to find estimates of the rate of convergence in the CLT that are sufficient for the law of the iterated logarithm.

Tikhomirov (1983) found a nonuniform estimate for  $|\Delta_V(x)|$  for strongly mixing strictly stationary random fields defined on the integer lattice  $Z_+^d$  and taking values in a finite-dimensional Euclidean space  $R^k$ . Tikhomirov (1983) extended his method to the estimation of the rate of convergence in the CLT for strictly stationary Hilbert-valued r.v.'s.

Asymptotic expansions for weakly dependent r.v.'s and random fields may be found in the papers by Heinrich (1985a), (1986), (1990a), Rhee (1985) and Götze and Hipp (1983), (1989). Stein's method was investigated by Barbour (1990) in the context of functional approximation of Wiener and other Gaussian processes.

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161

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