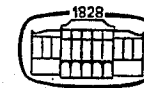


Strong Approximations in Probability and Statistics

M. Csörgő and P. Révész



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5.3. The law of iterated logarithm for the quantile process

In Section 4.5 we called attention to the fact that the analogue of the Glivenko–Cantelli theorem for the quantile process does not hold true without further restrictions. In this section we intend to point out that most of the strong laws proved for the empirical process are also true for the quantile process if we assume the conditions (4.5.10) and (4.5.12) of Theorem 4.5.6. We can do this in three ways: (i) apply Theorem 4.5.7 saying that the quantile process is near to a Kiefer process and deduce that theorems proved for a Kiefer process are also true for the quantile process; (ii) apply Theorem 5.2.2 saying that the quantile process is near to an empirical process and deduce that theorems proved for empirical processes are also true for quantile processes; (iii) apply Theorem 4.5.6 saying that a quantile process is near to a suitable uniform quantile process, hence theorems for the uniform quantile process extend to more general quantile processes. We followed the latter approach in the proofs of Theorems 5.2.2 and 5.2.5.

Applying any of the methods (i) and (ii) one gets (5.3.1)–(5.3.5) immediately, while method (iii) together with Theorem 4.5.6 implies (5.3.6). Thus we have:

Theorem 5.3.1. Under the conditions of Theorem 5.2.2 on F we have

$$(5.3.1) \quad \overline{\lim}_{n \rightarrow \infty} (\log \log n)^{-1/2} \sup_{0 < y < 1} f(\text{inv } F(y)) |q_n(y)| \stackrel{\text{a.s.}}{=} 2^{-1/2},$$

$$(5.3.2) \quad \overline{\lim}_{n \rightarrow \infty} (2 \log \log n)^{-1} \int_0^1 f^2(\text{inv } F(y)) q_n^2(y) dy \stackrel{\text{a.s.}}{=} \pi^{-2}, \text{ (cf. Theorem 5.1.3),}$$

$$(5.3.3) \quad \underline{\lim}_{n \rightarrow \infty} (\log \log n)^{1/2} \sup_{0 < y < 1} f(\text{inv } F(y)) |q_n(y)| \stackrel{\text{a.s.}}{=} 8^{-1/2} \pi, \text{ (cf. (5.1.8))}$$

$$(5.3.4) \quad \underline{\lim}_{n \rightarrow \infty} (\log \log n)^{1/2} \left(\int_0^1 f^2(\text{inv } F(y)) q_n^2(y) dy \right)^{1/2} \stackrel{\text{a.s.}}{=} 8^{-1/2}, \text{ (cf. (5.1.9)),}$$

$$(5.3.5) \quad \overline{\lim}_{n \rightarrow \infty} (2 \log \log n)^{-1/2} \sup_{\varepsilon < y < 1-\varepsilon} (y(1-y))^{-1/2} f(\text{inv } F(y)) |q_n(y)| \stackrel{\text{a.s.}}{=} 1.$$

Assume that conditions (4.5.10), (4.5.12) with $\gamma < 1$ hold true. Then there exists a $C > 0$ such that

$$(5.3.6) \quad \overline{\lim}_{n \rightarrow \infty} \sup_{\delta_n < y < 1-\delta_n} (y(1-y) \log \log n)^{-1/2} |f(\text{inv } F(y)) q_n(y)| \leq C \text{ a.s.}$$

where $\delta_n = 25n^{-1} \log \log n$.

The version of Theorem 5.1.2 in terms of $f(\text{inv } F(y)) q_n(y)$ is also straightforward. However the generalization of Theorem 5.1.5 resp. 5.1.6 to the quantile process is not immediate at all. The only such version we have at present is that of (5.3.6).

5.4. Asymptotic distribution results for some classical functionals of the empirical process

We have already seen that Donsker's theorem for the empirical process (Theorem 4.2.1) is a direct consequence of any one of the Brillinger and/or Kiefer type approximation theorems of Sections 4.3 and 4.4, and hence that, for example, Theorem 4.1.2 is implied by (1.5.3) and (1.5.4) of Theorem 1.5.1. Some further Corollary 4.4.1 type results follow here.

Corollary 5.4.1. Consider the empirical process $\beta_n(x) = \alpha_n(F(x))$. Let $y = F(x)$ be a continuous distribution function and let $g(y) \neq 0$ be a real valued function for which we also have

$$(5.4.1) \quad \sup_y |g(y)| < \infty.$$

There exist then a sequence of Brownian bridges $\{B_n(y); 0 \leq y \leq 1\}$ and a Kiefer process $\{K(y, t); 0 \leq y \leq 1, 0 \leq t < \infty\}$ such that

$$(5.4.2) \quad \left| \sup_{-\infty < x < \infty} \beta_n(x) g(F(x)) - \sup_{0 \leq y \leq 1} B_n(y) g(y) \right| \stackrel{\text{a.s.}}{=} O(n^{-1/2} \log n),$$

$$\left| \sup_{-\infty < x < \infty} \beta_n(x) g(F(x)) - \sup_{0 \leq y \leq 1} n^{-1/2} K(y, n) g(y) \right| \stackrel{\text{a.s.}}{=} O(n^{-1/2} \log^2 n),$$

$$(5.4.3) \quad \left| \sup_{-\infty < x < \infty} |\beta_n(x) g(F(x))| - \sup_{0 \leq y \leq 1} |B_n(y) g(y)| \right| \stackrel{\text{a.s.}}{=} O(n^{-1/2} \log n),$$

$$\left| \sup_{-\infty < x < \infty} |\beta_n(x) g(F(x))| - \sup_{0 \leq y \leq 1} |n^{-1/2} K(y, n) g(y)| \right| \stackrel{\text{a.s.}}{=} O(n^{-1/2} \log^2 n),$$

$$(5.4.4) \quad \left| \int_{-\infty}^{+\infty} \beta_n^2(x) g^2(F(x)) dF(x) - \int_0^1 B_n^2(y) g^2(y) dy \right|$$

$$\stackrel{\text{a.s.}}{=} O(n^{-1/2} \log n (\log \log n)^{1/2}),$$

$$\left| \int_{-\infty}^{+\infty} \beta_n^2(x) g^2(F(x)) dF(x) - \int_0^1 n^{-1} K^2(y, n) g^2(y) dy \right|$$

$$\stackrel{\text{a.s.}}{=} O(n^{-1/2} \log^2 n (\log \log n)^{1/2})$$

and

$$(5.4.5) \quad |n^{1/2}R_n - (\sup_y B_n(y)g(y) - \inf_y B_n(y)g(y))| \stackrel{\text{a.s.}}{=} O(n^{-1/2} \log n),$$

$$|n^{1/2}R_n - (\sup_y n^{-1/2}K(y, n)g(y) - \inf_y n^{-1/2}K(y, n)g(y))| \stackrel{\text{a.s.}}{=} O(n^{-1/2} \log^2 n),$$

where

$$R_n = D_n^+ + D_n^- = \sup_x (F_n(x) - F(x))g(F(x)) - \inf_x (F_n(x) - F(x))g(F(x)).$$

Since $B_n(y) \stackrel{\mathcal{D}}{=} B(y)$, and also $n^{-1/2}K(y, n) \stackrel{\mathcal{D}}{=} K(y, 1) \stackrel{\mathcal{D}}{=} B(y)$, a Brownian bridge for each n , we see that the limit distributions of the above-presented functionals of the empirical process agree with the distributions of the corresponding functionals of a Brownian bridge. For example, (5.4.2) and (5.4.3) with $g(y)=1$ give the asymptotic distribution of the classical Kolmogorov-Smirnov statistics (cf. Theorem 4.1.2) via (1.5.3) and (1.5.4). Again with $g(y)=1$ and applying (5.4.4), we get the limit distribution of the Cramér-von Mises statistic. Namely, Theorem 1.5.2 implies

$$(5.4.6) \quad \lim_{n \rightarrow \infty} P \left\{ \int_{-\infty}^{+\infty} \beta_n^2(x) dF(x) \leq u \right\} = P\{\omega^2 \leq u\}, \quad u \geq 0,$$

where the latter distribution function is given in Theorem 1.5.2. Another classical result, the distribution of the Kuiper (1960) statistic, can be obtained from (5.4.5) with $g(y)=1$ by Theorem 1.5.3. Namely we have

$$\lim_{n \rightarrow \infty} P\{n^{1/2}R_n \leq u\} = 1 - \sum_{j=1}^{\infty} 2(4(ju)^2 - 1)e^{-2j^2u^2}, \quad u \geq 0.$$

Taking now for example

$$(5.4.7) \quad g_\varepsilon(y) = \begin{cases} 0 & \text{if } 0 \leq y \leq \varepsilon \\ y^{-1} & \text{if } \varepsilon < y \leq 1, \end{cases}$$

$$g_\delta(y) = \begin{cases} (1-y)^{-1} & \text{if } 0 \leq y \leq \delta \\ 0 & \text{if } \delta < y \leq 1, \end{cases}$$

respectively

$$(5.4.8) \quad g_{\varepsilon, \delta}(y) = \begin{cases} (y(1-y))^{-1/2} & \text{if } 0 < \varepsilon \leq y \leq \delta < 1 \\ 0 & \text{otherwise,} \end{cases}$$

in (5.4.2) and (5.4.3), we conclude that the limit distributions of the Rényi (1953) statistics, respectively those of the Anderson-Darling (1952) statistics can be evaluated via the distributions of the corresponding functionals of

a Brownian bridge. We mention only two typical results obtainable this way. The first one of these is

$$(5.4.9) \quad \lim_{n \rightarrow \infty} P \left\{ \sup_{\varepsilon \leq F(x)} \frac{\beta_n(x)}{F(x)} \leq u \right\} = P \left\{ \sup_{\varepsilon \leq y} \frac{B(y)}{y} \leq u \right\}$$

$$= 2\Phi \left(u \left(\frac{\varepsilon}{1-\varepsilon} \right)^{1/2} \right) - 1 \quad u \geq 0, \varepsilon > 0.$$

The latter equality is true, since

$$\left\{ \frac{B(y)}{y}; 0 < y \leq 1 \right\} \stackrel{\mathcal{D}}{=} \left\{ W \left(\frac{1-y}{y} \right); 0 < y \leq 1 \right\},$$

and hence

$$P \left\{ \sup_{\varepsilon \leq y} \frac{B(y)}{y} \leq u \right\} = P \left\{ \sup_{\varepsilon \leq y} W \left(\frac{1-y}{y} \right) \leq u \right\}$$

$$= P \left\{ \sup_{0 < t \leq \frac{1-\varepsilon}{\varepsilon}} W(t) \leq u \right\},$$

which by (1.5.1) now gives the desired result. The second one we have in mind is

$$(5.4.10) \quad \lim_{n \rightarrow \infty} P \left\{ \sup_{\varepsilon \leq F(x)} \frac{|\beta_n(x)|}{F(x)} \leq u \right\} = P \left\{ \sup_{\varepsilon \leq y} \frac{|B(y)|}{y} \leq u \right\}$$

$$= \frac{4}{\pi} \sum_{k=0}^{\infty} \frac{(-1)^k}{(2k+1)} \exp \{ -(2k+1)^2 \pi^2 (1-\varepsilon) / 8\varepsilon u^2 \}, \quad u > 0, \varepsilon > 0.$$

The above two results ((5.4.9) and (5.4.10)) were first given by Rényi (1953) via classical limiting arguments. For a proof of these and some further similar ones along these lines via the invariance principle, we refer to Csörgő (1966, 1967).

So far we have seen how strong invariance principles (cf. Corollary 5.4.1) can be used to prove asymptotic distribution result like e.g., (5.4.6), (5.4.9), (5.4.10). In proving these results, we have not utilized the rates of approximation of Corollary 5.4.1 at all, and did not say anything about the problem of how fast these distribution functions themselves converged to their limits. The next result gives an answer to this problem.

Corollary 5.4.2 (Komlós, Major, Tusnády 1975a). *Let B_n, α_n be as in Theorem 4.4.1, and let ψ be a functional defined on $D(0, 1)$, satisfying the Lipschitz condition*

$$(5.4.11) \quad |\psi(u) - \psi(v)| \leq L \sup_{0 \leq y \leq 1} |u(y) - v(y)|$$

with some positive constant L . Assume further that the distribution of the random variable $\psi(B(y)) \stackrel{d}{=} \psi(B_n(y))$ ($n=1, 2, \dots$) has a bounded density with respect to Lebesgue measure. Then

$$(5.4.12) \quad \sup_{-\infty < x < \infty} |P\{\psi(\alpha_n(y)) \leq x\} - P\{\psi(B_n(y)) \leq x\}| = O\left(\frac{\log n}{n^{1/2}}\right).$$

The proof of this corollary is similar to that of Corollary 5.4.3 whose proof is given below.

As to the nearness of the processes α_n and B_n , Theorem 4.4.1 gives the best possible rate. It is an open question whether the rate of (5.4.12) is the best possible or not for any given Lipschitzian functional ψ . For example, in the Kolmogorov-Smirnov case (that is when ψ is the sup-functional) the rate of convergence is known to be $O(n^{-1/2})$ (cf. Gnedenko, Korolyuk and Skorohod 1960; Bickel 1974).

The rate in (5.4.12) does not hold necessarily for functionals not satisfying the Lipschitzian condition (5.4.11), but Theorem 4.4.1 might still give us a handle on occasions. The case we have in mind is that of the Cramér-von Mises statistic $\omega_n^2 = \int_0^1 \alpha_n^2(y) dy$. Let $N_n(x) = P\{\omega_n^2 \leq x\}$, $N(x) = P\{\int_0^1 B^2(y) dy \leq x\}$ and put $\Delta_n = \sup_{0 < x < \infty} |N_n(x) - N(x)|$. We have, of course, that $\lim_{n \rightarrow \infty} N_n(x) = N(x)$ for every real x (cf. (5.4.6)) and, in addition to this, it can be easily deduced from the first statement of (4.4.25), or from that of (5.4.4) that $\Delta_n = O(n^{-1/2} \log n (\log \log n)^{1/2})$. But, if we are a bit more circumspect, we can actually prove also

Corollary 5.4.3 (S. Csörgő 1976).

$$(5.4.13) \quad \Delta_n = O(n^{-1/2} \log n).$$

The latter statement is of interest, because it turns out to be a refinement of the best available result of this kind so far for the distribution of ω_n^2 . Namely Orlov (1974) proved that for any $\varepsilon > 0$ there exists a positive constant $b(\varepsilon)$ such that for each n $\Delta_n \leq b(\varepsilon)n^\varepsilon/n^{1/2}$. For a complete set of earlier work on this problem we refer to S. Csörgő (1976). We should also remark here that, though the rate of convergence of Δ_n in (5.4.13) is the best available one so far, it is probably far away from the best possible one. Indeed, a complete asymptotic expansion for the Laplace transform of ω_n^2 is given by S. Csörgő (1976) and, on the basis of his work, he

conjectures that Δ_n has the order of $1/n$ (concerning this latter problem we refer also to S. Csörgő and Stachó, 1979). For an improved result in this direction we refer to Götze (1979).

Proof of Corollary 5.4.3. Let $\omega^2(n) = \int_0^1 B_n^2(y) dy$, where $\{B_n(y); 0 \leq y \leq 1\}$ is the sequence of Brownian bridges for which the statement of Theorem 4.4.1 holds true, and let $\omega^2 = \int_0^1 B^2(y) dy$, where $B(y)$ is an arbitrary Brownian bridge. We will use the elementary fact that, if X, Y, Z are arbitrary r.v. such that $P\{|X - Y| > Z\} < \varepsilon$ for some $\varepsilon > 0$, then, for every real x ,

$$(5.4.14) \quad P\{Y \leq x - Z\} - \varepsilon \leq P\{X \leq x\} \leq P\{Y \leq x + Z\} + \varepsilon.$$

In (4.4.1) let $x = \frac{1}{2\lambda} \log n$, $D = C + \frac{1}{2\lambda}$. Then, with the notation $D_n = \sup_{0 \leq y \leq 1} |\alpha_n(y) - B_n(y)|$ and $\varepsilon_n = D \log n / \sqrt{n}$, we have

$$\begin{aligned} & P\{|\omega_n^2 - \omega^2(n)| > \varepsilon_n^2 + 2\varepsilon_n \omega(n)\} \\ &= P\left\{\left|\int_0^1 (\alpha_n(y) - B_n(y))(\alpha_n(y) + B_n(y)) dy\right| > \varepsilon_n^2 + 2\varepsilon_n \omega(n)\right\} \\ &\leq P\left\{\int_0^1 (\alpha_n(y) - B_n(y))^2 dy + 2 \int_0^1 |\alpha_n(y) - B_n(y)| |B_n(y)| dy > \varepsilon_n^2 + 2\varepsilon_n \omega(n)\right\}. \end{aligned}$$

Let the event of the latter probability statement be denoted by $E(\varepsilon_n)$. Then, from the above inequality,

$$\begin{aligned} P\{|\omega_n^2 - \omega^2(n)| > \varepsilon_n^2 + 2\varepsilon_n \omega(n)\} &\leq P\{E(\varepsilon_n), D_n \leq \varepsilon_n\} + P\{D_n > \varepsilon_n\} \\ &\leq P\left\{\varepsilon_n^2 + 2\varepsilon_n \int_0^1 |B_n(y)| dy > \varepsilon_n^2 + 2\varepsilon_n \omega(n)\right\} + n^{-1/2} L \\ &= P\left\{\int_0^1 |B_n(y)| dy > \omega(n)\right\} + n^{-1/2} L \\ &\leq P\{\omega(n) > \omega(n)\} + n^{-1/2} L = n^{-1/2} L, \end{aligned}$$

where the last inequality follows from that of Schwartz, and L is the positive absolute constant of (4.4.1).

Now we apply (5.4.14) with $X = \omega_n^2$, $Y = \omega(n)$, $Z = \varepsilon_n^2 + 2\varepsilon_n \omega(n)$ and $\varepsilon = n^{-1/2} L$, and get

$$P\{A_n(x)\} - n^{-1/2} L \leq P\{\omega_n^2 \leq x\} \leq P\{G_n(x)\} + n^{-1/2} L,$$

where $A_n(x) = \{\omega^2 \leq x - \varepsilon_n^2 - 2\varepsilon_n \omega\}$, $G_n(x) = \{\omega^2 \leq x + \varepsilon_n^2 + 2\varepsilon_n \omega\}$ and $x > 0$.

Solving the corresponding quadratic inequalities for the events $A_n(x)$ and $G_n(x)$ we find that $P\{A_n(x)\}=0$ if $x \leq \varepsilon_n^2$, while

$$P\{A_n(x)\} = P\{\omega^2 \leq x + \varepsilon_n^2 - (4\varepsilon_n^2 x)^{1/2}\} \cong P\{\omega^2 \leq x - \delta_n\}, \quad \text{if } x \geq \varepsilon_n^2,$$

and

$$G_n(x) \subset \{\omega^2 \leq x + \delta_n\} \quad \text{with } \delta_n = 3\varepsilon_n^2 + (8\varepsilon_n^4 + 4\varepsilon_n^2 x)^{1/2}, \quad x > 0.$$

Hence, with $\delta_n = \delta_n(x)$ ($x > 0$) as in the preceding line,

$$N(x - \delta_n(x)) - n^{-1/2}L \leq N_n(x) \leq N(x + \delta_n(x)) + n^{-1/2}L,$$

and whence

$$\begin{aligned} A_n &\leq \sup_x P\{x - \delta_n(x) < \omega^2 \leq x + \delta_n(x)\} + n^{-1/2}L \\ &= \sup_x \int_x^{x + 2\delta_n(x)} v(y) dy + n^{-1/2}L \leq 2 \sup_x (\delta_n(x) \sup_{x \leq y \leq x + 2\delta_n(x)} v(y)) + n^{-1/2}L \\ &\leq 2K \sup_x (\delta_n(x) \sup_{x \leq y \leq x + 2\delta_n(x)} (y+1)^{-1/2}) + n^{-1/2}L \\ &\leq 2K \sup_x \frac{\delta_n(x)}{(x+1)^{1/2}} + n^{-1/2}L = 2K \sup_x \frac{3\varepsilon_n^2 + (8\varepsilon_n^4 + 4\varepsilon_n^2 x)^{1/2}}{(x+1)^{1/2}} + n^{-1/2}L \\ &\leq 2K \sup_x \frac{3\varepsilon_n^2(x+1)^{1/2} + (8\varepsilon_n^4(x+1) + 4\varepsilon_n^2(x+1))^{1/2}}{(x+1)^{1/2}} + n^{-1/2}L \\ &= O(\varepsilon_n) = O(n^{-1/2} \log n), \end{aligned}$$

where $v(x) = \frac{d}{dx} N(x)$, and the third inequality for A_n above is by $v(y)(y+1)^{1/2} \leq v(y)(y^{1/2}+1) \leq K$ uniformly in $y \in (0, \infty)$, since it is known that not only the density function $v(x)$ of ω^2 but $v(x)x^{1/2}$ too is bounded on the positive half-line (cf. Lemma 8 in § 5, S. Csörgő 1976).

While Corollary 5.4.1 provides us with weak convergence results for the classical functionals of the empirical process with weight functions $g(y)$ satisfying the condition (5.4.1), the weight function $g(y) = (y(1-y))^{-1/2}$, $0 < y < 1$, which is probably the most natural one, does not fit into its framework. An application of Theorem 4.4.1, however, turns out to be a good initial step also towards this direction of weak convergence problems, which we are going to consider now.

Let $V_n(x) = n^{1/2}(F_n(x) - F(x))/(F(x)(1-F(x)))^{1/2}$, where F is a continuous distribution function and, for $0 \leq \varepsilon < \delta \leq 1$, define $V^n(\varepsilon, \delta) = \sup_{\varepsilon < F(x) < \delta} V_n(x)$ and $W^n(\varepsilon, \delta) = \sup_{\varepsilon < F(x) < \delta} |V_n(x)|$. We note that Anderson and Darling (1952)

derived the Laplace transform of the asymptotic distribution of $W^n(\varepsilon, \delta)$ with $0 < \varepsilon < \delta < 1$. It is also natural to ask how one could choose normalizing factors for given sequences ε_n, δ_n so that $V^n(\varepsilon_n, \delta_n), W^n(\varepsilon_n, \delta_n), V^n(0, 1)$ and $W^n(0, 1)$ should have a non-degenerate asymptotic distribution. Indeed this question was asked and answered recently by Jaeschke (1975) and Eicker (1976, 1979) and further studied and developed by Jaeschke (1976, 1979).

For the construction of confidence intervals for F ,

$$\hat{V}_n(x) = \begin{cases} 0, & \text{if } F_n(x) = 0 \text{ or } 1 \\ n^{1/2}(F_n(x) - F(x))/(F_n(x)(1-F_n(x)))^{1/2}, & \text{otherwise,} \end{cases}$$

is more convenient than $V_n(x)$. It is shown by Jaeschke (1976) that for \hat{V}_n the same assertions hold as for V_n , including also the case $\varepsilon_n = 0, \delta_n = 1$. The latter is also a generalization of the earlier results of Eicker (1976). Here we formulate these results with $\varepsilon_n = 0, \delta_n = 1$.

Towards stating these results, let $E_c(t) = \exp(-c \exp(-t))$ ($c \geq 0$) and $a(\cdot, \cdot)$ be as in Theorem 1.9.1.

Theorem 5.4.1 (Jaeschke 1976). *We have*

$$(5.4.15) \quad \lim_{n \rightarrow \infty} P\{V^n(0, 1) \leq a(t, \log n)\} = E_1(t)$$

and

$$(5.4.16) \quad \lim_{n \rightarrow \infty} P\{W^n(0, 1) \leq a(t, \log n)\} = E_2(t), \quad -\infty < t < +\infty.$$

Theorem 5.4.2 (Eicker 1976, Jaeschke 1976). *Let*

$$\hat{V}^n(\varepsilon, \delta) = \sup_{\varepsilon < F(x) < \delta} \hat{V}_n(x) \quad \text{and} \quad \hat{W}^n(\varepsilon, \delta) = \sup_{\varepsilon < F(x) < \delta} |\hat{V}_n(x)|.$$

Then

$$(5.4.17) \quad \lim_{n \rightarrow \infty} P\{\hat{V}^n(0, 1) \leq a(t, \log n)\} = E_1(t)$$

and

$$(5.4.18) \quad \lim_{n \rightarrow \infty} P\{\hat{W}^n(0, 1) \leq a(t, \log n)\} = E_2(t), \quad -\infty < t < +\infty.$$

Just as in the case of the classical empirical process $\beta_n(x) = \alpha_n(F(x))$, we may from now on take $F \in U(0, 1)$, since F is assumed to be continuous. As to the proof of the two theorems formulated above, we give only the main steps, in order to demonstrate how Theorem 4.4.1 can be applied in this situation. Here we follow Jaeschke (1976).

The proof of Theorem 5.4.1 is based on the following lemma:

Lemma 5.4.1 (Jaeschke 1976). For $-\infty < t < +\infty$ we have

$$(5.4.19) \quad \lim_{n \rightarrow \infty} P\{W^n(0, \varepsilon_n) \vee W^n(1 - \varepsilon_n, 1) \leq a(t, \log n)\} = 1,$$

where $\varepsilon_n = n^{-1} \log^3 n$.

For a proof of this lemma we refer to that of Lemma 4 in Jaeschke (1976).

Proof of Theorem 5.4.1. Since by Theorem 4.4.1

$$(2 \log \log n)^{1/2} \sup_{\varepsilon_n < y < 1 - \varepsilon_n} |V_n(y) - B_n(y)/(y(1-y))^{1/2}| \\ \stackrel{\text{a.s.}}{=} O\left(\left(\frac{\log \log n}{\log n}\right)^{1/2}\right) = o(1),$$

Corollary 1.9.1 and Lemma 5.4.1 imply Theorem 5.4.1.

For a similar proof of the second theorem we need two further lemmas.

Lemma 5.4.2 (Jaeschke 1976). For $\varepsilon_n \geq n^{-1} \log n$ and $a_n = (2 \log \log n)^{1/2}$ we have

$$(5.4.20) \quad a_n \hat{V}^n(\varepsilon_n, 1 - \varepsilon_n) \stackrel{\text{a.s.}}{=} a_n V^n(\varepsilon_n, 1 - \varepsilon_n) + o(1).$$

Proof. Due to Theorem 5.1.6 we have

$$\sup_{n^{-1} \log n < y < 1} |1 - y^{-1} F_n(y)| \stackrel{\text{a.s.}}{=} O((\log \log n / \log n)^{1/2})$$

and

$$\sup_{0 < y < 1 - n^{-1} \log n} |1 - (1 - F_n(y))/(1 - y)| \stackrel{\text{a.s.}}{=} O((\log \log n / \log n)^{1/2}).$$

Whence

$$a_n \hat{V}^n(\varepsilon_n, 1 - \varepsilon_n) \stackrel{\text{a.s.}}{=} a_n V^n(\varepsilon_n, 1 - \varepsilon_n) (1 + O((\log \log n / \log n)^{1/2})).$$

Again Theorem 5.1.6 implies $a_n V^n(\varepsilon_n, 1 - \varepsilon_n) \stackrel{\text{a.s.}}{=} O(\log \log n)$, and this is the assertion of (5.4.20).

Lemma 5.4.3 (Eicker 1976, Jaeschke 1976). With $\varepsilon_n = n^{-1} \log n$ and $-\infty < t < +\infty$, we have

$$(5.4.21) \quad \lim_{n \rightarrow \infty} P\{\hat{W}^n(0, \varepsilon_n) \vee \hat{W}^n(1 - \varepsilon_n, 1) \leq a(t, \log n)\} = 1.$$

For a proof of this lemma we refer to that of Lemma 6 in Jaeschke (1976).

Proof of Theorem 5.4.2. A combination of Lemmas 5.4.2, 5.4.3 and Theorem 5.4.1 yields (5.4.17) and (5.4.18).

5.5. Asymptotic distribution results for some classical functionals of the quantile process

On the basis of Theorems 4.5.6 and 4.5.7 it is quite immediate to construct a Corollary 5.4.1 type statement for the quantile process $f(\text{inv } F(y))q_n(y)$, provided we assume conditions (4.5.10) and (4.5.12). Such an analogue of Corollary 5.4.1 immediately implies, among others, the following typical statements:

$$(5.5.1) \quad \lim_{n \rightarrow \infty} P\left\{\sup_{0 < y < 1} f(\text{inv } F(y))q_n(y) \leq u\right\} = P\left\{\sup_{0 \leq y \leq 1} B(y) \leq u\right\} \\ = 1 - e^{-2u^2}, \quad u \geq 0, \quad (\text{cf. (1.5.3)}),$$

$$(5.5.2) \quad \lim_{n \rightarrow \infty} P\left\{\sup_{0 < y < 1} |f(\text{inv } F(y))q_n(y)| \leq u\right\} = P\left\{\sup_{0 \leq y \leq 1} |B(y)| \leq u\right\} \\ = 1 - \sum_{k \neq 0} (-1)^{k+1} e^{-2k^2 u^2}, \quad u \geq 0, \quad (\text{cf. (1.5.4)}),$$

$$(5.5.3) \quad \lim_{n \rightarrow \infty} P\{\mathcal{R}_n \leq u\} = P\left\{\sup_{0 < y < 1} B(y) - \inf_{0 < y < 1} B(y) \leq u\right\} \\ = 1 - \sum_{j=1}^{\infty} 2(4(ju)^2 - 1)e^{-2j^2 u^2}, \quad u \geq 0, \quad (\text{cf. Theorem 1.5.3}),$$

where $\mathcal{R}_n = \sup_{0 < y < 1} f(\text{inv } F(y))q_n(y) - \inf_{0 < y < 1} f(\text{inv } F(y))q_n(y)$,

$$(5.5.4) \quad \lim_{n \rightarrow \infty} P\left\{\sup_{\varepsilon \leq y} \frac{f(\text{inv } F(y))q_n(y)}{y} \leq u\right\} = P\left\{\sup_{\varepsilon \leq y} \frac{B(y)}{y} \leq u\right\} \\ = 2\Phi\left(u\left(\frac{\varepsilon}{1-\varepsilon}\right)^{1/2}\right) - 1, \quad u \geq 0, \quad \varepsilon > 0, \quad (\text{cf. (5.4.9)}),$$

and

$$(5.5.5) \quad \lim_{n \rightarrow \infty} P\left\{\sup_{\varepsilon \leq y} \frac{|f(\text{inv } F(y))q_n(y)|}{y} \leq u\right\} = P\left\{\sup_{\varepsilon \leq y} \frac{|B(y)|}{y} \leq u\right\} \\ = \frac{4}{\pi} \sum_{k=0}^{\infty} \frac{(-1)^k}{(2k+1)} \exp\{-(2k+1)^2 \pi^2 (1-\varepsilon)/8\varepsilon u^2\}, \quad u > 0, \quad \varepsilon > 0, \\ (\text{cf. (5.4.10)}).$$

We again call attention to the fact that the above asymptotic results hold true, assuming only the reasonably weak conditions (4.5.10) and (4.5.12).

An analogue of Corollary 5.4.2 is also immediate.

Corollary 5.5.1. Let B_n, q_n be as in Theorem 4.5.7, and let ψ be defined on $D(0, 1)$, satisfying the Lipschitz condition of (5.4.11). Assume further

Theorem 7.3.2. Let $\alpha_n(y)$, $u_n(y)$, $q_n(y)$, $K(y, n)$ be as in Theorems 4.4.3, 4.5.3 and 4.5.7 respectively, and let $\{v_n\}$ be a sequence of positive integer valued random variables, and assume that $v_n \xrightarrow{P} \infty$ as $n \rightarrow \infty$. Then

$$(7.3.2) \quad \sup_{0 \leq y \leq 1} |\alpha_{v_n}(y) - v_n^{-1/2} K(y, v_n)| \xrightarrow{P} 0,$$

$$(7.3.3) \quad \sup_{0 \leq y \leq 1} |u_{v_n}(y) - v_n^{-1/2} K(y, v_n)| \xrightarrow{P} 0,$$

$$(7.3.4) \quad \sup_{0 < y < 1} |f(F^{-1}(y))q_{v_n}(y) - v_n^{-1/2} K(y, v_n)| \xrightarrow{P} 0.$$

We note that (7.3.2) also holds true with $y = F(x)$ for $\alpha_n(F(x)) = \beta_n(x)$. Further, if instead of $v_n \xrightarrow{P} \infty$ we have that $v_n \xrightarrow{\text{a.s.}} \infty$, then (7.3.2)–(7.3.4) also hold true with probability one.

Proof of Theorem 7.3.2. Combining the respective statements (b) and (c) of Theorem 7.1.1 with (4.4.23), (4.5.8) and (4.5.25), the above statements follow.

Theorem 7.3.3. If v_n of Theorem 7.3.2 also satisfies the condition (7.2.7), then

$$(7.3.5) \quad \alpha_{v_n}(\cdot) \xrightarrow{\mathcal{D}} B(\cdot),$$

$$(7.3.6) \quad f(F^{-1}(\cdot))q_{v_n}(\cdot) \xrightarrow{\mathcal{D}} B(\cdot).$$

Proof. First we note that (7.2.7) implies that $v_n \xrightarrow{P} \infty$ as $n \rightarrow \infty$. It follows then from (7.3.2) and (7.3.4) that, in order to prove (7.3.5) and (7.3.6), it suffices to show that

$$(7.3.7) \quad v_n^{-1/2} K(\cdot, v_n) \xrightarrow{\mathcal{D}} B(\cdot).$$

Now the proof of (7.3.7) can be done along the lines of that of (7.2.8).

Remark 7.3.1. The first paper on the random sample size empirical process α_{v_n} was written by Pyke (1968). He proved (7.3.5) under the assumption that $v_n/n \xrightarrow{P} 1$. The above method of proof also extends to empirical processes defined in terms of multivariate random variables (cf. M. Csörgő, S. Csörgő, Fischler and Révész 1975). As to random sum limit theorems, one of the first papers was that of Anscombe (1952) (cf. also Doeblin 1938, 1940).

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