

On the Breiman conjecture

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Breiman 1965

Coin tossing \longrightarrow random walk S_1, S_2, \dots

Put Y_1, Y_2, \dots the interarrival times between the zeros of S_1, S_2, \dots

X, X_1, X_2, \dots iid $\mathbf{P}\{X = 0\} = \frac{1}{2} = \mathbf{P}\{X = 1\}$.

$$T_n = \frac{\sum_{i=1}^n X_i Y_i}{\sum_{i=1}^n Y_i}$$

is the proportion of the time that the random walk spends in $[0, \infty)$.

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Arc-sine law

In this case:

$$\lim_{n \rightarrow \infty} \mathbf{P} \{ T_n \leq x \} = \frac{2}{\pi} \arcsin \sqrt{x}$$

In general

Y, Y_1, Y_2, \dots non-negative iid rv's with df G

X, X_1, X_2, \dots iid with df F , independent from Y, Y_1, Y_2, \dots ,

$E|X| < \infty$

$$T_n = \frac{\sum_{i=1}^n X_i Y_i}{\sum_{i=1}^n Y_i}$$

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Remark

If $\mathbf{E}Y < \infty$, then

$$\frac{\sum_{i=1}^n X_i Y_i}{\sum_{i=1}^n Y_i} = \frac{\sum_{i=1}^n X_i Y_i}{n} \xrightarrow{\text{a.s.}} \mathbf{E}X.$$

$\mathbf{E}|X| < \infty$ implies (T_n) is tight.

Remark

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Theorem (Breiman, 1965)

If T_n converges in distribution for every F , and the limit is non-degenerate for at least one F , then $Y \in D(\beta)$, for some $\beta \in [0, 1)$.

Conjecture (Breiman)

If T_n has a non-degenerate limit for some F , then $Y \in D(\beta)$ for some $\beta \in [0, 1)$.



Breiman, L.

On some limit theorems similar to the arc-sin law

Teor. Verojatnost. i Primenen. **10** 351–360, 1965.

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$D(\beta)$

Domain of attraction of an β -stable law:

$$Y \in D(\beta) \Leftrightarrow 1 - G(x) = \frac{\ell(x)}{x^\beta},$$

where ℓ is slowly varying ($\ell(\lambda x)/\ell(x) \rightarrow 1$ for any $\lambda > 0$ as $x \rightarrow \infty$).

$D(0)$

$Y \in D(0)$ if $1 - G(x)$ is slowly varying
in which case (Darling, 1952)

$$\frac{\max\{Y_i : i = 1, 2, \dots, n\}}{\sum_{i=1}^n Y_i} \xrightarrow{\mathbf{P}} 1$$

and so

$$\frac{\sum_{i=1}^n X_i Y_i}{\sum_{i=1}^n Y_i} \xrightarrow{\mathcal{D}} X$$

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Limits

Theorem (Breiman)

Assume that $Y \in D(\beta)$, $\beta \in (0, 1)$, and $\mathbf{E}|X|^{\beta+\varepsilon} < \infty$, for some $\varepsilon > 0$. Then $T_n \xrightarrow{\mathcal{D}} T$, where

$$\mathbf{P}\{T \leq x\} = \frac{1}{2} + \frac{1}{\pi\beta} \arctan \left[\frac{\int |u-x|^\beta \operatorname{sgn}(x-u) F(du)}{\int |u-x|^\beta F(du)} \tan \frac{\pi\beta}{2} \right].$$

$$\mathbf{P}\{T > x\} \approx \mathbf{P}\{X > x\}$$

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$$\mathbf{P}\{T > x\} \approx \mathbf{P}\{X > x\}$$

$$\mathbf{E}|X|^{2+\delta} < \infty$$

Theorem (Mason & Zinn, 2005)

Assume that $\mathbf{E}|X|^{2+\delta} < \infty$. Then $T_n \rightarrow R$, where R is non-degenerate, iff $Y \in D(\beta)$, $\beta \in [0, 1)$.

Studentization

Other type of self-normalization (Logan & Mallows & Rice & Shepp, 1973):

$$\frac{\sum_{i=1}^n X_i}{\sqrt{\sum_{i=1}^n X_i^2}},$$

X, X_1, X_2, \dots iid. Student's T -statistic:

$$\frac{\sum_{i=1}^n X_i}{\sqrt{n} \sqrt{\frac{\sum_{i=1}^n (X_i - \bar{X})^2}{n-1}}}$$

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Conjecture (Logan & Mallows & Rice & Shepp, 1973)

$$\frac{\sum_{i=1}^n X_i}{\sqrt{\sum_{i=1}^n X_i^2}} \xrightarrow{\mathcal{D}} W,$$

where $\mathbf{P}\{|W| = 1\} < 1$, iff $X \in D(\alpha)$, $\alpha \in (0, 2]$; if $\alpha > 1$, $\mathbf{E}X = 0$; if $\alpha = 1$, $X \in D(\text{Cauchy})$.

Giné & Götze & Mason (1997): W is standard normal iff $X \in D(2)$ and $\mathbf{E}X = 0$

Chistyakov & Götze (2004): in general

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$\mathbf{E}|X| < \infty$.

$$\phi_X(t) = \mathbf{E}e^{itX}$$

Theorem (K – Mason)

Assume that for some $\mathbf{E}X = 0$, $1 < \alpha \leq 2$, positive slowly varying function L at zero and $c > 0$,

$$\frac{-\log(\Re\phi_X(t))}{|t|^\alpha L(|t|)} \rightarrow c, \text{ as } t \rightarrow 0.$$

Whenever

$$\frac{\sum_{i=1}^n X_i Y_i}{\sum_{i=1}^n Y_i} \xrightarrow{\mathcal{D}} W \quad (W \text{ nondegenerate})$$

then $Y \in D(\beta)$ for some $\beta \in [0, 1)$.

What does this condition mean?

As $-\log \Re\phi_X(t) \sim 1 - \Re\phi_X(t)$, $t \rightarrow 0$,

$$\frac{-\log(\Re\phi_X(t))}{|t|^\alpha L(|t|)} \rightarrow c \Leftrightarrow \frac{1 - \Re\phi_X(t)}{|t|^\alpha L(|t|)} \rightarrow c.$$

For $\alpha < 2$ this holds iff (Pitman)

$$\mathbf{P}\{|X| > x\} \sim L(1/x)x^{-\alpha}c\Gamma(\alpha)\frac{2}{\pi}\sin\left(\frac{\pi\alpha}{2}\right)$$

If $\mathbf{E}X = 0$ and $X \in D(\alpha)$ then this condition is satisfied.

Also if $\mathbf{E}X = 0$, $\mathbf{E}X^2 < \infty$ then the condition of the theorem is satisfied ($\alpha = 2$, $c = \sigma^2/2$).

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Proposition

Assume that the assumptions of the theorem hold. Then for some $0 < \gamma \leq 1$

$$\mathbf{E} \frac{\sum_{i=1}^n Y_i^\alpha}{\left(\sum_{i=1}^n Y_i\right)^\alpha} \rightarrow \gamma. \quad (*)$$

Proposition

If (\star) holds with some $\gamma \in (0, 1]$ then $Y \in D(\beta)$, for some $\beta \in [0, 1)$, where $-\beta \in (-1, 0]$ is the unique solution of

$$\text{Beta}(\alpha - 1, 1 - \beta) = \frac{\Gamma(\alpha - 1)\Gamma(1 - \beta)}{\Gamma(\alpha - \beta)} = \frac{1}{\gamma(\alpha - 1)}.$$

In particular, $Y \in D(0)$ for $\gamma = 1$. Conversely, if $Y \in D(\beta)$, $0 \leq \beta < 1$, then (\star) holds with

$$\gamma = \frac{\Gamma(\alpha - \beta)}{\Gamma(\alpha)\Gamma(1 - \beta)} = \frac{1}{(\alpha - 1)\text{Beta}(\alpha - 1, 1 - \beta)}.$$

Extension of a result by Fuchs, Joffe and Teugels (2001), where $\alpha = 2$.

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$$\mathbf{E} \frac{\sum_{i=1}^n Y_i^\alpha}{\left(\sum_{i=1}^n Y_i\right)^\alpha} \rightarrow \gamma \quad (\star)$$

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In particular, $Y \in D(0)$ for $\gamma = 1$. Conversely, ...

For $\alpha = 2$ this gives $1 - \gamma = \beta$.

Sketch of the proof

$$\begin{aligned}\mathbf{E} \frac{\sum_{i=1}^n Y_i^\alpha}{\left(\sum_{i=1}^n Y_i\right)^\alpha} &= n \mathbf{E} \frac{Y_1^\alpha}{\left(\sum_{i=1}^n Y_i\right)^\alpha} \\ &= \frac{n}{\Gamma(\alpha)} \mathbf{E} \int_0^\infty Y_1^\alpha e^{-t \sum_{i=1}^n Y_i} t^{\alpha-1} dt \\ &= \frac{n}{\Gamma(\alpha)} \int_0^\infty t^{\alpha-1} \mathbf{E} \left(e^{-t Y_1} Y_1^\alpha \right) \left(\mathbf{E} e^{-t Y_1} \right)^{n-1} dt \\ &=: \frac{n}{\Gamma(\alpha)} \int_0^\infty t^{\alpha-1} \phi_\alpha(t) \phi_0(t)^{n-1} dt.\end{aligned}$$

Note that for $\alpha = 2$ we have $\phi_\alpha = \phi_0''$.

$$s \int_0^\infty t^{\alpha-1} \phi_\alpha(t) e^{s \log \phi_0(t)} dt \rightarrow \gamma \Gamma(\alpha), \quad s \rightarrow \infty.$$

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$$\phi_\alpha(t) = \mathbf{E}e^{-tY} Y^\alpha, \quad \phi_0(t) = \mathbf{E}e^{-tY}$$

By Karamata's Tauberian theorem

$$\lim_{t \rightarrow 0} \frac{\int_0^t y^{\alpha-1} \phi_\alpha(y) dy}{1 - \phi_0(t)} = \gamma \Gamma(\alpha).$$

After some further calculation

$$t^{\alpha-1} \frac{\int_0^\infty \bar{G}(u) u^{\alpha-1} e^{-ut} du}{\int_0^\infty \bar{G}(u) e^{-ut} du} \rightarrow \gamma \Gamma(\alpha), \text{ as } t \searrow 0.$$

$$u^{1-\alpha} e^{-ut} = \frac{1}{\Gamma(\alpha-1)} \int_0^\infty y^{\alpha-2} e^{-(y+t)u} dy,$$

which holds for $u > 0$ and $\alpha \in (1, 2]$. *Weyl-transform*, or *Weyl-fractional integral* of the function e^{-ut}

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We obtain

$$\int_1^\infty (u-1)^{\alpha-2} u^{-\alpha} \frac{g_\infty(x/u)}{g_\infty(x)} du = \frac{k^M * g_\infty(x)}{g_\infty(x)} \rightarrow [\gamma(\alpha-1)]^{-1}$$

with

$$g_\infty(x) = \int_0^\infty \bar{G}(ux) u^{\alpha-1} e^{-u} du.$$

$$k^M * h(x) = \int_0^\infty h(x/u) k(u) / u du$$

Mellin-convolution of h and k .

Drasin-Shea theorem implies that $g_\infty(x)$ is regularly varying at infinity with index $0 \geq \rho > -1$.

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Notation

$\text{id}(a, b, \nu)$ infinitely divisible distribution on \mathbb{R}^d with characteristic exponent

$$iu'b - \frac{1}{2}u'au + \int \left(e^{iu'x} - 1 - iu'xI(|x| \leq 1) \right) \nu(dx),$$

where $b \in \mathbb{R}^d$, $a \in \mathbb{R}^{d \times d}$ is a positive semidefinite matrix and ν is the Lévy measure.

Theorem (K & Mason, 2012)

If along a subsequence $\{n'\}$

$$\frac{1}{a_{n'}} \sum_{i=1}^{n'} Y_i \xrightarrow{\mathcal{D}} W_2, \text{ as } n' \rightarrow \infty,$$

where $W_2 \sim \text{id}(0, b, \Lambda)$, then

$$\left(\frac{\sum_{i=1}^{n'} X_i Y_i}{a_{n'}}, \frac{\sum_{i=1}^{n'} Y_i}{a_{n'}} \right) \xrightarrow{\mathcal{D}} (W_1, W_2), \quad n' \rightarrow \infty,$$

where $(W_1, W_2) \sim \text{id}(\mathbf{0}, \mathbf{b}, \Pi)$

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Theorem (K & Mason, 2012)

i.e. its characteristic function

$$\begin{aligned} \Psi(\theta_1, \theta_2) = \mathbf{E} e^{i(\theta_1 W_1 + \theta_2 W_2)} = \exp \left\{ i(\theta_1 b_1 + \theta_2 b_2) \right. \\ \left. + \int_0^\infty \int_{-\infty}^\infty \left(e^{i(\theta_1 x + \theta_2 y)} - 1 - (i\theta_1 x + i\theta_2 y) \mathbf{1}_{\{x^2 + y^2 \leq 1\}} \right) \right. \\ \left. F(dx/y) \wedge(dy) \right\}. \end{aligned}$$

$$H(x) = \mathbf{P} \left\{ \frac{W_1}{W_2} \leq x \right\} = \frac{1}{2} - \frac{1}{2\pi} \int_{-\infty}^\infty \frac{\Im \Psi(u, -ux)}{u} du.$$

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i.e. its characteristic function

$$\begin{aligned} \Psi(\theta_1, \theta_2) = \mathbf{E} e^{i(\theta_1 W_1 + \theta_2 W_2)} = \exp \left\{ i(\theta_1 b_1 + \theta_2 b_2) \right. \\ \left. + \int_0^\infty \int_{-\infty}^\infty \left(e^{i(\theta_1 x + \theta_2 y)} - 1 - (i\theta_1 x + i\theta_2 y) \mathbf{1}_{\{x^2 + y^2 \leq 1\}} \right) \right. \\ \left. F(dx/y) \wedge(dy) \right\}. \end{aligned}$$

$$H(x) = \mathbf{P} \left\{ \frac{W_1}{W_2} \leq x \right\} = \frac{1}{2} - \frac{1}{2\pi} \int_{-\infty}^\infty \frac{\Im \Psi(u, -ux)}{u} du.$$

Feller class

ξ, ξ_1, \dots iid with df F , $S_n = \sum_{i=1}^n \xi_i$. F is in the *centered Feller class*, if there exists B_n , such that every subsequence n' has a further subsequence n'' , such that

$$\frac{S_{n''}}{B_{n''}} \xrightarrow{\mathcal{D}} W,$$

where W is non-degenerate.

Theorem (Feller (1966), Maller (1979))

Y is in the centered Feller class, iff

$$\limsup_{x \rightarrow \infty} \frac{x^2 \mathbf{P}\{|Y| > x\} + x |\mathbf{E} Y I(|Y| \leq x)|}{\mathbf{E}[Y^2 I(|Y| \leq x)]} < \infty.$$

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Surprising result

Theorem (K & Mason, 2012)

The subsequential limit distributions of

$$T_n = \frac{\sum_{i=1}^n X_i Y_i}{\sum_{i=1}^n Y_i}$$

are continuous for all X with finite expectation if and only if $Y \in \mathcal{F}_c$.

Outline

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Earlier results

Partial solution

Results

Sketch of the proof

Subsequential limits

Results

Further remarks

Towards Lévy processes

$$(W_1, W_2) \stackrel{\mathcal{D}}{=} (a_1 + U, a_2 + V),$$

where $(a_1, a_2) = \left(\left(b - \int_0^1 x \Lambda(dx) \right) \mathbf{E}X, b - \int_0^1 x \Lambda(dx) \right)$

$$\mathbf{E}e^{i(\theta_1 U + \theta_2 V)} = \exp \left\{ \int_0^\infty \int_{-\infty}^\infty \left(e^{i(\theta_1 x + \theta_2 y)} - 1 \right) F(dx/y) \Lambda(dy) \right\}$$

Under the conditions of the theorem

$$\left(\frac{\sum_{1 \leq i \leq n't} X_i Y_i}{a_{n'}}, \frac{\sum_{1 \leq i \leq n't} Y_i}{a_{n'}} \right)_{t>0} \stackrel{\mathcal{D}}{\rightarrow} (a_1 t + U_t, a_2 t + V_t)_{t>0},$$

where (U_t, V_t) , $t \geq 0$, is the corresponding Lévy process.

$$\frac{U_t}{V_t} \xrightarrow{\mathcal{D}} ?, \quad t \rightarrow 0 \text{ or } t \rightarrow \infty$$



Kevei, P, Mason, D.M.

Randomly Weighted Self-normalized Lévy Processes
Stochastic Processes and their Applications, **123** (2) 2013,
490–522.